Continuous-time stochastic games

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A B S T R A C T
We study continuous-time stochastic games, with a focus on the existence of their equilibria that are insensitive to a small imprecision in the specification of players’ evaluations of streams of payoffs.
We show that the stationary, namely, time-independent, discounting game has a stationary equilibrium and that the discounting game and the more general game with time-separable payoffs have an epsilon equilibrium that is an epsilon equilibrium in all games with a sufficiently small perturbation of the players’ valuations.
A limit point of discounting valuations need not be a discounting valuation as some of the “mass” may be pushed to infinity; it is represented by an average of a discounting valuation and a mass at infinity.
We show that for every such limit point there is a strategy profile that is an epsilon equilibrium of all the discounting games with discounting valuations that are sufficiently close to the limit point.

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1. Introduction

1.1. Motivating continuous-time stochastic games

A fundamental shortfall of economic forecasting is its inability to make deterministic forecasts. A notable example is the repeated occurrence of financial crises. Even though each financial crisis was forecasted by many economic experts, many other experts claimed just before each occurrence that “this time will be different,” as documented in the book This Time Is Different: Eight Centuries of Financial Folly by Reinhart and Rogoff. On the other hand, many economic experts have forecasted financial crises that never materialized. Finally, even those who correctly predicted a financial crisis could not predict the time of its occurrence and many of them warned that it was imminent when in fact it occurred many years later.

An analogous observation comes from sports, e.g., soccer. In observing the play of a soccer game, or knowing the qualities of both competing teams, we can recognize a clear advantage of one team over the other. Therefore, a correct forecast is that the stronger team is more likely than the other team to score the next goal. However, it is impossible to be sure which team will be the first to score the next goal, and it is impossible to forecast the time of the scoring.

In both cases, the state of the economy and the score of the game can dramatically change in a split second, and players’ actions impact the likelihood of each state change.

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A game-theoretic model that accounts for the change of a state between different stages of the interaction, and where such change is impacted by the players’ actions, is a stochastic game. However, no single deterministic-time dynamic game can capture the important common feature of these two examples: namely, the probability of a state change in any short time interval can be positive yet arbitrarily small. This feature can be analyzed by studying the asymptotic behavior in a sequence of discrete-time stochastic games, where the individual stage represents short time intervals that converge to zero and the transition probabilities to a new state also converge to zero. The limit of such a sequence of discrete-time stochastic games is a continuous-time stochastic game; see Neyman (2012, 2013).

The analysis of continuous-time stochastic games enables us to model and analyze the important properties that (1) the state of the interaction can change, (2) the stochastic law of states is impacted by players’ actions, and (3) in any short time interval the probability of a discontinuous state change can be positive (depending on the players’ actions), but arbitrarily small for sufficiently short time intervals.

Accounting for stochastic state changes – both gradual (continuous) and instantaneous (discontinuous) but with infinitesimal probability in any short time interval – is common in the theory of continuous-time finance; see, e.g., Merton (1992). However, in continuous-time finance theory the stochastic law of states is not impacted by agents’ actions. Accounting for (mainly deterministic and more recently also stochastic) continuous state changes that are impacted by agents’ actions is common in the theory of differential games; see, e.g., Isaacs (1999).

In the present paper we study the three above-mentioned properties, with a special focus on discontinuous state changes. To do this we study the continuous-time stochastic game with finitely many states, since if there are finitely many states, then any state change is discontinuous.

1.2. Difficulties with continuous-time strategies

A classical game-theoretic analysis of continuous-time games entails a few unavoidable pathologies, as for some naturally defined strategies there is no distribution on the space of plays that is compatible with these strategies, and for other naturally defined strategies there are multiple distributions on the space of plays that are compatible with these strategies. In addition, what defines a strategy is questionable. In spite of these pathologies, we will describe unambiguously equilibrium payoffs and equilibrium strategies in continuous-time stochastic games.

Earlier game-theoretic studies overcome the above-mentioned pathologies by restricting the study either to strategies with inertia, or to Markov (memoryless) strategies that select an action as a function of (only) the current time and state. See, e.g., Bergin and MacLeod (1993) and Simon and Stinchcombe (1989) for the study of continuous-time supergames, Perry and Reny (1993) for the study of continuous-time bargaining, Isaacs (1999) for the study of differential games, and Zachrisson (1964), Lai and Tanaka (1982), and Guo and Hernandez-Lerma (2003, 2005) for the study of continuous-time stochastic games, termed in the above-mentioned literature Markov games, or Markov chain games. A more detailed discussion of the relation between these earlier contributions and the present paper appears in Section 7.

The common characteristic of these earlier studies is that they each consider only a subset of strategies, so that a profile of strategies selected from the restricted class defines a distribution over plays of the games, and thus optimality and equilibrium are well defined, but only within the restricted class of strategies. Therefore, in an equilibrium, there is no beneficial unilateral deviation within only the restricted class of strategies. Therefore, there is neither optimality nor nonexistence of beneficial unilateral deviation claims for general strategies.

The restriction to Markov (memoryless) strategies (see, e.g., Zachrisson, 1964; Rykov, 1966; Miller, 1968a, 1968b; Kacurman, 1969, 1971, 1975; Guo and Hernandez-Lerma, 2009) is (essentially) innocuous in the study of continuous-time, discounted or finite-horizon Markov decision processes and two-person zero-sum stochastic games.

The two properties of the discounted or finite-horizon payoffs that make this restriction innocuous are: (1) impatience, namely, the contribution to the payoff of the play in the distant future is negligible, and (2) time-separability of the payoff, namely, the payoff of a play on $[0, \infty)$ is, for every $s > 0$, a sum of a function of the play on the time interval $[0, s)$ and a function of the play on the time interval $[s, \infty)$.

The Markov (memoryless) assumption is restrictive in the Markov decision processes with a payoff that is not time-separable and in the zero-sum stochastic games with non-impatient payoffs.

The Markov (memoryless) assumption is restrictive in the non-zero-sum game-theoretic framework. It turns out that in discounted stochastic games (with finitely many states and actions) a Markov strategy profile that is an equilibrium in the universe of Markov strategies is also an equilibrium in the universe of history-dependent strategies. However, the set of equilibrium strategies and equilibrium payoffs in the universe of Markov strategies is a proper subset of those in the universe of history-dependent strategies.

A fundamental shortcoming of the restriction to memoryless strategies (and to oblivious strategies, which depend only on the state process) arises in the study of approximate equilibria that are robust to small changes of the discounting valuation. For example, there is a continuous-time (and discrete-time, even two-person zero-sum, e.g., the Big Match, Blackwell and Ferguson, 1968) stochastic game for which there is no Markov strategy profile that is an approximate equilibrium in all the discounted games with a sufficiently small discounting rate.

1 I.e., a list, one for each player.
The present paper proves that every continuous-time stochastic game with finitely many states and actions has such strategy profiles.

1.3. Discrete-time stochastic games

We recall the basic properties of the discrete-time stochastic game model (see, e.g., Shapley, 1953; Neyman and Sorin, 2003; Mertens et al., 2015); this will enable us to point out its similarities and differences with the continuous-time model to be described later.

In a discrete-time stochastic game, play proceeds in stages and the stage state, the stage action profile, the previous stage, and the next stage are well defined. The stage payoff is a function \( g(z,a) \) of the stage state \( z \) and the stage action profile \( a \), and the transitions to the next stage \( z' \) are defined by conditional probabilities \( p(z' | z, a) \) of the next state \( z' \) given the present state \( z \) and the stage action profile \( a \). Players’ stage-action choices are made simultaneously and are observed by all players following the stage play. Pure, mixed, and behavioral strategies are well and unambiguously defined in the discrete-time model.

If at stage \( m = 0, 1, \ldots \), the state is \( z_m \) and the action profile played is \( a_m \), the stage payoff is \( g_m := g(z_m, a_m) \). A play of the discrete-time stochastic game generates a stream of payoffs \( g_0, g_1, \ldots \).

A measure \( w \) on the nonnegative integers defines a payoff function \( u_w \) on streams of payoffs by \( u_w(g_0, g_1, \ldots) = \sum_{m=0}^{\infty} w(m) g_m \). The stationary discounting corresponds to measures \( w \) of the form \( w(m) = \lambda^m \) where \( \lambda \) is a positive constant and \( 0 < \lambda < 1 \). A (not necessarily stationary) discounting corresponds to a measure \( w \) with \( 0 \leq w(m+1) \leq w(m) \).

The common normalization \( u_w(1, 1, \ldots) = 1 \) corresponds to \( \sum_{m=0}^{\infty} w(m) = 1 \) and is useful in comparing outcomes of the same game with different discountings.

The measures that define the payoff function of different players need not be identical.

1.4. Continuous-time stochastic games: payoffs

In a continuous-time game, the payoff is usually defined as an accumulation of infinitesimal payoffs. If at time \( t \in [s, s+ds) \) the state is \( z \) and the action profile played is \( a \) the accumulation of the payoff in the time interval \( [s, s+ds) \) is \( g(z,a) ds \). Therefore, if the function \( t \mapsto (z_t, q_t) \), where \( z_t \) is the state at time \( t \) and \( q_t \) is the action profile at time \( t \), is measurable, then the accumulation of payoffs in the time interval \( [0,s) \) is \( \int_0^s g_t dt \) where \( g_t := g(z_t, q_t) \).

The unnormalized, respectively, the normalized, \( \rho \)-discounted payoff, \( \rho > 0 \), is \( \int_0^s e^{-\rho t} g_t dt \), respectively, \( \rho \int_0^s e^{-\rho t} g_t dt \), and the unnormalized, respectively, the normalized, \( s \)-horizon payoff, \( s > 0 \), is \( \int_0^s g_t dt \), respectively, \( s \int_0^s g_t dt \).

Similarly, the payoffs are also defined on the space of measurable functions \( t \mapsto (z_t, x_t) \), where \( z_t \) is the state at time \( t \) and \( x_t \) is the mixed action-profile (i.e., mixtures of pure action profiles) at time \( t \), by setting \( g(z, x) \) to be the linear extension of \( g(z, a) \) and \( g_t = g(z_t, x_t) \).

In general, the payoff in a stochastic game can be an arbitrary function of the function \( t \mapsto (z_t, x_t) \).

Consider for example a continuous-time stochastic game that models a single quarter-final or semi-final in the UEFA Champions League. The rule that specifies the winner, which qualifies for the next stage, follows.

A single quarter-final or semi-final is played under the knockout system, on a home-and-away basis (two legs). The team that scores the greater aggregate of goals in the two matches qualifies for the next stage. Otherwise, the team that scores more away goals qualifies for the next stage.

If this procedure does not produce a result, i.e., if both teams score the same number of goals at home and away, a 30-minute period of overtime is played at the end of the second leg. If both teams score the same number of goals during overtime, away goals count double (i.e., a visiting team that scored in the overtime qualifies).

If no goals are scored during overtime, kicks from the penalty mark determine which team qualifies for the next stage. For simplicity, assume that the outcome of the penalty kicks is random, e.g., each team is equally likely to win this phase.

Each match lasts 90 minutes. Therefore, if at the state of the stochastic game, \( z_t \), is the cumulative score at the \( t \)-th minute, the rules provide us with a function of \( (z_0, z_{30}, z_{60}) \) that specifies the probability (1, 0, or 1/2) that team 1 wins.

In the present paper we focus on the structure of the stochastic game that satisfies several conditions.

First, the payoff to a player, say player \( i \), is a function \( u_1 \) of the stream \( g : t \mapsto g_t \) of his “stage payoffs.”

Second, we assume that the payoff function \( u_1 \) is linear and monotonic and is defined over the ordered linear space of all bounded measurable streams of payoffs \( f : t \mapsto f_t \) (with the order \( f \geq g \) iff \( f_t \geq g_t \) \( \forall t \)), and the natural linear structure \( \alpha f + \beta g : t \mapsto \alpha f_t + \beta g_t \).

We allow the payoff functions to differ among players.

The payoff function \( u_1 \) is normalized if \( u(1) = 1 \), where \( 1 \) is the stream \( f \) with \( f_1 = 1 \) for every \( t \geq 0 \).

The payoff function \( u_1 \) is impatient if \( u(1_{\geq s}) \rightarrow s \rightarrow 0 \), where \( 1_{\geq s} \) is the stream \( f \) with \( f_1 = 1 \) for every \( t \geq s \) and \( f_t = 0 \) for every \( 0 \leq t < s \).

A payoff function \( u_1 \) that is defined (on the space of bounded measurable streams of payoffs \( f \)) by \( u(f) := \int_0^\infty f_t dw(t) \), where \( w \) is a measure on \( [0, \infty) \), is called a time-separable payoff and is denoted by \( u_w \).

A time-separable payoff function is linear and monotonic. If, in addition, \( w \) is a probability measure, then \( u_w \) is normalized.
In the other direction, any linear and monotonic payoff function \( u \) that is impatient defines a unique measure \( w \) on \([0, \infty)\) such that for any bounded continuous stream of payoffs \( f \) we have \( u(f) = w(f) := \int_0^\infty f(t) \, dw(t) \).

An important special class of time-separable payoff functions \( u \), called *impatient valuations*, is defined below.

For a given measurable subset \( B \) of \( \mathbb{R}_+ = [0, \infty) \), we denote by \( 1_B \) the indicator function of \( B \) (i.e., the stream of payoffs \( f \) such that \( f_t = 1 \) if \( t \in B \) and \( f_t = 0 \) if \( t \not\in B \)).

An impatient valuation is a time-separable payoff function \( u \) that is normalized and for every measurable subset \( B \) of \( \mathbb{R}_+ \) we have \( u(1_B) \geq \rho(\mathbb{1}_{c+B}) \), where \( c \geq 0 \) and \( c + B := \{ c + b \mid b \in B \} \). The last condition reflects the preference of advancing the time of positive payoffs.

The time-separable payoff function that is defined by a measure \( w \) on \( \mathbb{R}_+ \) is an impatient valuation if and only if \( w \) is a probability measure and \( w((b, b + c)) \) is, for every fixed \( c > 0 \), nonincreasing in \( b \geq 0 \).

An approximate description of a normalized, linear, monotonic, and impatient payoff function \( u \) is obtained by specifying, within a small positive error term, its value on finitely many continuous streams of payoffs with bounded support.

Therefore, we say that a sequence \( u_k \) of normalized time-separable payoff functions \( u_k \) converges if the sequence \( u_k(f) \) converges for every continuous function \( f \) with bounded support.

Examples of converging time-separable payoff functions include: the normalized \( \rho \)-discounted payoff as \( \rho \) goes to 0, the normalized \( s \)-horizon payoff as \( s \) goes to infinity, and the sequence of payoff functions \( u_k \) that are defined by the Dirac measures on a sequence of points that converge to some fixed point \( x_0 \).

Let \( u_k = u_{w_k} \) be a converging sequence of normalized time-separable payoff functions. Then, the limit \( u(f) := \lim_{k \to \infty} u_k(f) \) exists for every continuous function \( f \) that is defined on \([0, \infty)\). (Note that a continuous function \( f \) on \([0, \infty)\) is identified with a continuous function \( f \) on \([0, \infty)\) for which the limit of \( f_t \) as \( t \) goes to infinity, exists.)

The limit \( u \), which is defined on the space of all continuous functions on \([0, \infty)\), is linear, monotonic, and normalized. However, it need not coincide with a time-separable payoff function \( u \), as some of the corresponding \( w_k \)-mass may be pushed to infinity.

The restriction of the limit \( u \) to the space of continuous functions \( f \) on \([0, \infty)\) is represented by a (unique) probability measure \( w \) on \([0, \infty)\):

\[
\lim_{k \to \infty} u_k(f) = \lim_{k \to \infty} \int_0^\infty f(t) \, dw_k(t) = \int_0^\infty f(t) \, dw(t).
\]

If \( u_k \) are impatient valuations then the limiting probability measure \( w \), which may have an atom at \( \infty \) and at 0, is absolutely continuous\(^2\) on \([0, \infty)\) and \( dw \) is nonincreasing on \((0, \infty)\); equivalently, \( w \) is a probability measure on \([0, \infty)\) with \( w((b, b + c)) \) being, for every fixed \( c > 0 \), nonincreasing in \( b \geq 0 \).

Therefore, in studying approximate optimization that is insensitive to a small imprecision in the specification of an impatient valuation, it is useful to enlarge the space of measures on \([0, \infty)\) to the space \( W \) of all measures on \([0, \infty)\).

For a given measure \( w \in W \), the \( w \)-payoff is the integral \( \int_{[0, \infty]} g(t) \, dw(t) \), where \( g_\infty = \lim_{s \to \infty} \frac{1}{s} \int_0^s g(t) \, dt \) if the limit exists.

The definition of \( g_\infty \) is motivated by the characterization of valuations that follows.

A payoff function \( u \) satisfies the time value of money principle if

\[
u(f) \geq u(g) \text{ whenever } f_0 \geq g_0 \text{ and } \int_0^s f(t) \, dt \geq \int_0^s g(t) \, dt \text{ for all } s > 0.
\]

The time value of money principle reflects the preference for expediting the time of positive payoffs: the faster the accumulation of payoffs, the better. The principle states that a unit of payoffs in a given time period is preferable to its being spread over future time periods.

Note that a payoff function that satisfies the time value of money principle is monotonic.

A valuation is a normalized linear payoff function \( u \) that satisfies the time value of money principle. A patient valuation is a valuation \( u \) such that \( u(1_{[0, \infty]}) = 1 \) for all \( s \geq 0 \).

One can show that a payoff function \( u \) is a patient valuation if and only if it is a linear function over the bounded stream of payoffs such that

\[
\lim_{s \to \infty} \frac{1}{s} \int_0^s f(t) \, dt \leq u(f) \leq \lim_{s \to \infty} \frac{1}{s} \int_0^s f(t) \, dt,
\]

and that any valuation is a mixture of an impatient valuation and a patient one.

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\(^2\) A finite measure \( \mu \) on Borel subsets of the real line is absolutely continuous if for every positive number \( \varepsilon > 0 \) there is a positive number \( \delta > 0 \) such that \( \mu(A) < \varepsilon \) for all Borel sets \( A \) of Lebesgue measure less than \( \delta \). If \( \mu \) is absolutely continuous then there exists a Lebesgue integrable function \( g \), denoted by \( g_\mu \), on the real line such that \( \mu(A) = \int_A g \, dt \) for all Borel subsets \( A \) of the real line.
1.5. Continuous-time stochastic games: transitions

The transitions of states are described by nonnegative real-valued transition rates $\mu(z', z, a)$, defined for all triples $z', z, a$, of two distinct states $z' \neq z$ and an action profile $a$.

The nonnegative real number $\mu(z', z, a)$ describes the rate of transition from state $z$ to state $z'$ when action profile $a$ is played at state $z$.

If the state at time $t$ is $z_t = z$ and the action profile at all times $t \leq s < t + \delta$ is $a$, then the probability that the state $z_{t+\delta} = z' \neq z$ is approximately $\delta \mu(z', z, a)$ for small $\delta$; explicitly, for two distinct states $z, z'$ and a time $t$, conditional on the history of the play up to time $t$, $z_t = z$, and $a_i = a$ for all $t \leq s < t + \delta$, we have $P(z_{t+\delta} = z')/\delta \to \mu(z', z, a)$ as $\delta \to 0+$.

The actions in a continuous-time game can depend on time. Therefore, one has to define the transitions resulting from time-dependent actions.

Given a measurable time-dependent action profile $a_t$, $t \geq 0$, and two distinct states $z' \neq z$, the transitions obey

$$P(z_{t+\delta} = z' | z_t = z) = \int_t^{t+\delta} \mu(z', z, a_s) \, ds + o(\delta) \text{ as } \delta \to 0 + .$$

1.6. Continuous-time stochastic games: strategies

The classical definition of a pure strategy in a discrete-time game is a local definition. It defines, for each stage $m$, the selected action at stage $m$ as a function of the information derived from (signals from) players' and nature's moves.

In a stochastic game with observable states and actions, a strategy specifies the action at the discrete stage $m$ as a function of the sequence of past states and actions of other players and of the current state.

An equivalent definition is the global one that specifies the sequence of actions of the player as a function of the entire sequence of states and actions of other players, but with the obvious no-looking-ahead property. The no-looking-ahead property, termed also no anticipation, asserts that the specified action at stage $t$ depends only on players' past observable actions and past observable chance moves.

The equivalence of the two definitions follows from the fact that the set of decision times does not order; see Section 2.2.1.

The set of times $t \geq 0$ is not well ordered. Therefore, there are naturally defined local strategies that do not integrate into a global one; see Section 2.2.1. Therefore, one reverts to the global form of a strategy.

A full pure strategy of player $i$ is a function $\sigma^i$ from plays – functions $h : t \mapsto (z_t, a_t)$, $t \geq 0$ – to functions $t \mapsto \sigma^i_t(h)$ (with proper measurability assumptions), with the no-anticipation property.

Explicitly, if $h = (z_t, a_t)_{t \geq 0}$ and $h' = (z_t', b_t)_{t \geq 0}$ are two plays with $(z_t, a_t) = (z_t', b_t)$ for $t \leq s$, then $\sigma^i_t(h) = \sigma^i_t(h')$ on $t \leq s$.

Full pure strategies are important in the study of subgame perfect equilibria. Our focus in the present paper is on equilibria. Therefore, we confine attention to reduced strategies, namely strategies that define the action to be taken only on histories that are compatible with the strategy.

A reduced pure strategy, henceforth a pure strategy, of player $i$ is a function $\sigma^i$ from plays – functions $h : t \mapsto (z_t, a_t)$, $t \geq 0$ – to functions $t \mapsto \sigma^i_t(h)$ (with proper measurability assumptions), with the no-anticipation property and independent of its own past actions.

Explicitly, if $h = (z_t, a_t)_{t \geq 0}$ and $h' = (z_t', b_t)_{t \geq 0}$ are two plays with $(z_t, a^{-1}_t) = (z'_t, b^{-1}_t)$ for $t \leq s$, then $\sigma^i_t(h) = \sigma^i_t(h')$ on $t \leq s$.

The assumption that the (reduced pure) strategy choice of player $i$ is independent of player $i$'s past actions is conceptually innocuous, as the strategy itself and other players' past actions define player $i$'s past actions.

However, the independence of $\sigma^i_t(h)$ of at least player $i$'s very recent past own actions is required for technical reasons, as otherwise, even in the one-person game, there are strategies that do not define a play; see Section 2.2.1.

The reversion to global strategies does not phase out all pathologies, even in the special case of continuous-time supergames.

Indeed, there are (“naturally” defined global) pure strategy profiles $\sigma$, e.g., in the continuous-time matching pennies game, for which there is no play $h$ such that $\sigma(h) = h$, and there are ("naturally") defined global pure strategy profiles $\sigma$ in continuous-time supergames for which there is more than one play $h$ with $\sigma(h) = h$ (see Proposition 1 in Section 2.2.1).

We illustrate the above comment by explicit examples of such strategy profiles. This illustration may serve also, in particular, as examples of general strategies. Assume that each player has two actions, labeled 0 and 1.

Define the global strategy $\sigma^1$ of player 1 by $\sigma^1((a^1_t)_{t \geq 0}) = (a^1_t)_{t \geq 0}$ where $a^1_0 = 1$, and for $t > 0$, $a^1_t = 1$ if $\exists t > t_n \downarrow 0$ such that $a^2_{t_n} = 0$, and $a^1_t = 0$ otherwise. Note that there is a sequence $t > t_n \downarrow 0$ such that $a^2_{t_n} = 0$ if and only if there is a sequence $t_n \downarrow 0$ such that $a^2_{t_n} = 0$.

Define the global strategy $\sigma^2$ of player 2 by $\sigma^2((a^2_t)_{t \geq 0}) = (a^2_t)_{t \geq 0}$ where $a^2_0 = 1$, and for $t > 0$, $a^2_t = 1$ if $\exists t > t_n \downarrow 0$ such that $a^1_{t_n} = 1$, and $a^2_t = 0$ otherwise.

Assume that $\sigma((a^1_t)_{t \geq 0}) = (a^1_t)_{t \geq 0}$, where $\sigma = (\sigma^1, \sigma^2)$. If there is a sequence $t_n \downarrow 0$ such that $a^2_{t_n} = 0$, then, by the definition of $\sigma^1$, $a^1_t = 1$ for every $t \geq 0$, and therefore, by the definition of $\sigma^2$, $a^2_t = 1$ for every $t \geq 0$, contradicting the
existence of a sequence \( t_n \downarrow 0 \) such that \( a_{t_n}^2 = 0 \). If there is no sequence \( t_n \downarrow 0 \) such that \( a_{t_n}^2 = 0 \), then, by the definition of \( \sigma^1, a_t^1 = 0 \) for every \( t > 0 \), and therefore, by the definition of \( \sigma^2, a_t^2 = 0 \) for every \( t > 0 \), contradicting the nonexistence of a sequence \( t_n \downarrow 0 \) such that \( a_{t_n}^2 = 0 \). Therefore, there is no play \((a_t)_{t \geq 0}\) such that \( \sigma((a_t)_{t \geq 0}) = (a_t)_{t \geq 0}\).

Define the global strategy \( \tau^2 \) of player 2 by \( \tau^2((a_t^1)_{t \geq 0}) = (a_t^2)_{t \geq 0} \) where \( a_0^2 = 1 \), and for \( t > 0, a_t^2 = 0 \) if \( \exists t_n > t \) such that \( a_{t_n}^1 = 1 \), and \( a_t^2 = 1 \) otherwise. Note that the two plays \((a_t)_{t \geq 0}\) where \( a_0 = (1,1) \) and either \( a_t = (1,1) \) for all \( t > 0 \) or \( a_t = (0,1) \) for all \( t > 0 \) are compatible with the strategy pair \((\sigma^1, \tau^2)\).

Therefore, there is no proper strategic normal form defined over all strategies.

However, equilibrium is unambiguously defined as a profile of strategies with no unilateral beneficial deviation, as formally defined below.

1.7. Continuous-time stochastic games: equilibria

A strategy profile \( \sigma = (\sigma^i)_{i \in N} \) in a continuous-time stochastic game with player set \( N \) is called admissible if for every player \( i \) and every strategy \( \tau^i \) (recall that \( \tau^i \) is independent of its past actions), the strategy profile \((\sigma^{-i}, \tau^i)\) (where \( \sigma^{-i} = (\sigma^j)_{j \neq i} \)) defines a unique distribution \( P_{\sigma^{-i},\tau^i} \) on plays \( h = (z_t, x_t)_{t \geq 0} \) with right continuous state function \([0, \infty) \ni t \mapsto z_t\).

The uniqueness applies only to the space of plays with right continuous state functions. Therefore we henceforth assume that the state function of a play is right continuous.

An example of an admissible strategy profile in a continuous-time stochastic game is a profile of pure stationary strategies.

An example of a strategy profile that is not admissible is \((\sigma^1,0)\), where \( \sigma^1 \) is the strategy of player 1 that is defined in the previous section and 0 is the strategy of player 2 that always plays 0. The play with \( a_t = (1,0) \) for every \( t \) is the only one that is compatible with \((\sigma^1,0)\). However, if player 2 deviates to the strategy \( \sigma^2 \) (of the previous section) then there is no play that is compatible with \((\sigma^1,\sigma^2)\). Similarly, for any strategy \( \eta^1 \) of player 1, \((\eta^1,\sigma^2)\) is not an admissible strategy profile.

An \( \epsilon \)-equilibrium is an admissible strategy profile for which no single player can benefit by more than \( \epsilon \) from a unilateral deviation. An equilibrium is a 0-equilibrium.

For example, if \( \tilde{w} = (w^i)_{i \in N} \) is a profile of measures on \([0, \infty)\), then a strategy profile \( \sigma \) is an \( \epsilon \)-equilibrium of \( \Gamma_{\tilde{w}} \) if \( \sigma \) is admissible and for every player \( i \) and every strategy \( \tau^i \) of player \( i \), we have

\[
\gamma^i_{\tilde{w}}(z, \sigma) \geq \gamma^i_{\tilde{w}}(z, \sigma^{-i}, \tau^i) - \epsilon,
\]

(1)

where

\[
\gamma^i_{\tilde{w}}(z, \sigma) := E_{P_{\sigma}}^{\tilde{w}} \int_{[0, \infty)} g^i_t dw^i(t), \quad \gamma^i_{\tilde{w}}(z, \sigma^{-i}, \tau^i) := E_{P_{\sigma^{-i},\tau^i}}^{\tilde{w}} \int_{[0, \infty]} \bar{g}^i_t dw^i(t),
\]

\[
g^i_t := \bar{g}^i_t := g^i_t(z_t, x_t) \quad \text{for } t < \infty,
\]

\[
g^i_{\infty} := \lim_{s \to \infty} \frac{1}{s} \int_0^s g^i_t dt, \quad \bar{g}^i_{\infty} := \lim_{s \to \infty} \frac{1}{s} \int_0^s \bar{g}^i_t dt.
\]

Obviously, if \( \tilde{w} \) is supported on \([0, \infty)\), as in the case of stationary discounting or finite-horizon games, \( \gamma^i_{\tilde{w}}(z, \sigma) = \gamma^i_{\tilde{w}}(z, \sigma) \), and then we write \( \gamma^i_{\tilde{w}}(z, \sigma) \) for short.

In the remainder of this section we introduce the continuous-time equilibrium conditions, which are analogous to those in the discrete-time model, for three special cases: the time-independent discounting game, the finite horizon game, and the limiting-average game.

We show that each one of these equilibrium conditions corresponds to inequality (1) for a suitable profile of measures \( \tilde{w} \).

Fix a profile \( \tilde{\rho} = (\rho_t)_{t \in N} \) of discounting rates. The strategy profile \( \sigma \) is an \( \epsilon \)-equilibrium, \( \epsilon \geq 0 \), of the normalized \( \tilde{\rho} \)-discounted game if \( \sigma \) is an admissible strategy profile and for every state \( z \), every player \( i \), and every strategy \( \tau^i \) of player \( i \), we have

\[
E_{\sigma}^{\tilde{w}} \int_0^{\infty} g^i_t \rho_t e^{-\rho t} dt \geq E_{\sigma^{-i},\tau^i}^{\tilde{w}} \int_0^{\infty} \bar{g}^i_t \rho_t e^{-\rho t} dt - \epsilon.
\]

This inequality corresponds to inequality (1) with the profile \( \tilde{w} = (w^i)_{i \in N} \) of measures where \( w^i((t, \infty)) = e^{-\rho t} \) or, equivalently, where \( w^i \) is the probability measure on \([0, \infty)\) with density \( \rho_t e^{-\rho t} \). Indeed, the left-hand side of the above inequality corresponds to \( \gamma^i_{\tilde{w}}(z, \sigma) \) and the right-hand side of the above inequality corresponds to \( \gamma^i_{\tilde{w}}(z, \sigma^{-i}, \tau^i) \).
Fix a finite horizon \( s > 0 \). The strategy profile \( \sigma \) is an \( \varepsilon \)-equilibrium, \( \varepsilon \geq 0 \), of the normalized \( s \)-horizon game if \( \sigma \) is an admissible strategy profile and for every state \( z \), every player \( i \), and every strategy \( \tau^i \) of player \( i \), we have

\[
\gamma^i_s(z, \sigma) := E_z^\sigma \frac{1}{s} \int_0^s g^i_t \, dt \geq E^\sigma_{\sigma^{-i}, \tau^i} \frac{1}{s} \int_0^s g^i_t \, dt - \varepsilon.
\]

This inequality corresponds to inequality (1) with the profile \( \bar{w} = (w^i)_{i \in N} \) of measures where \( w^i \) is the uniform probability distribution on \([0, s]\), namely, where \( w^i \) is the probability measure on \([0, s]\) with density \( \frac{1}{s} \). Indeed, for this profile of measures the left-hand side of the above inequality corresponds to \( \gamma^i_w(z, \sigma) \) and the right-hand side of the above inequality corresponds to \( \gamma^i_w(z, \sigma^{-i}, \tau^i) \).

The strategy profile \( \sigma \) is an \( \varepsilon \)-equilibrium, \( \varepsilon \geq 0 \), of the limiting-average game if \( \sigma \) is an admissible strategy profile and for every state \( z \), every player \( i \), and every strategy \( \tau^i \) of player \( i \), we have

\[
E^\sigma \liminf_{s \to \infty} \frac{1}{s} \int_0^s g^i_t \, dt \geq E^\sigma_{\sigma^{-i}, \tau^i} \limsup_{s \to \infty} \frac{1}{s} \int_0^s g^i_t \, dt - \varepsilon.
\]

This inequality corresponds to inequality (1) with the profile \( \bar{w} = (w^i)_{i \in N} \) of probability measures on \([0, \infty]\) where \( w^i(\infty) = 1 \). Indeed, for this profile of probability measures the left-hand side of the above inequality corresponds to \( \gamma^i_w(z, \sigma) \) and the right-hand side of the above inequality corresponds to \( \gamma^i_w(z, \sigma^{-i}, \tau^i) \).

### 1.8. Uniform equilibrium

In this section we define a uniform equilibrium payoff of a continuous-time stochastic game. The definition is the natural concept that is analogous to the much-studied uniform equilibrium payoffs of discrete-time stochastic games.

A uniform \( \varepsilon \)-equilibrium payoff of the non-zero-sum continuous-time stochastic game is a vector \( u = (u^i(z))_{i \in N, z \in S} \in \mathbb{R}^{S \times N} \) such that there is an admissible strategy profile \( \sigma_{\varepsilon} \) and a sufficiently long finite horizon \( s_{\varepsilon} \) such that for every initial state \( z \), every player \( i \in N \), every \( s \geq s_{\varepsilon} \), and every strategy \( \sigma^{-i} \) of player \( i \), we have

\[
\gamma^i_s(z, \sigma_{\varepsilon}^{-i}, \sigma^i) - \varepsilon \leq u^i(z) \leq \gamma^i_s(z, \sigma_{\varepsilon}) + \varepsilon.
\]

A uniform equilibrium payoff of the non-zero-sum continuous-time stochastic game is a vector \( u = (u^i(z))_{i \in N, z \in S} \in \mathbb{R}^{S \times N} \) that is a uniform \( \varepsilon \)-equilibrium payoff for every \( \varepsilon > 0 \).

A uniform \( \varepsilon \)-equilibrium strategy profile (or equilibrium strategy, for short) of the non-zero-sum continuous-time stochastic game is an admissible strategy profile \( \sigma_{\varepsilon} \) such that there is a sufficiently long finite horizon \( s_{\varepsilon} \) such that for every initial state \( z \), every player \( i \in N \), every \( s \geq s_{\varepsilon} \), and every strategy \( \sigma^i \) of player \( i \), we have

\[
\gamma^i_s(z, \sigma_{\varepsilon}^{-i}, \sigma^i) - \varepsilon \leq u^i(z) \leq \gamma^i_s(z, \sigma_{\varepsilon}) + \varepsilon,
\]

where \( u \) is a uniform equilibrium payoff.

It is important to note that a strategy profile \( \sigma_{\varepsilon} \) that is an \( \varepsilon \)-equilibrium strategy in all sufficiently long finite-horizon games – for which there is a sufficiently long finite-horizon \( s_{\varepsilon} \) such that for every initial state \( z \), every player \( i \in N \), every \( s \geq s_{\varepsilon} \), and every strategy \( \sigma^i \) of player \( i \), we have \( \gamma^i_s(z, \sigma_{\varepsilon}^{-i}, \sigma^i) - \varepsilon \leq \gamma^i(z, \sigma_{\varepsilon}) \) – need not be a uniform \( \varepsilon \)-equilibrium strategy. The reason is that the payoff \( \gamma^i(z, \sigma_{\varepsilon}) \) need not converge as the horizon \( s \) goes to \( \infty \).

The uniform equilibrium and the limiting-average equilibrium concepts are both tailored for the long-horizon undiscounted games. However, they are two different concepts.

A uniform \( \varepsilon \)-equilibrium strategy need not be a limiting-average \( 2\varepsilon \)-equilibrium strategy, and, vice versa, a limiting-average \( \varepsilon \)-equilibrium strategy need not be a uniform \( 2\varepsilon \)-equilibrium strategy.

A third equilibrium concept, which extends simultaneously both of the above, is the uniform-limiting-average, or uniform-l-average for short.

A uniform-l-average \( \varepsilon \)-equilibrium payoff of the non-zero-sum continuous-time stochastic game is a vector \( u = (u^i(z))_{i \in N, z \in S} \in \mathbb{R}^{S \times N} \) such that there is an admissible strategy profile \( \sigma_{\varepsilon} \) and a sufficiently long finite horizon \( s_{\varepsilon} \) such that for every initial state \( z \), every player \( i \in N \), every \( s \geq s_{\varepsilon} \), and every strategy \( \tau^i \) of player \( i \), we have

\[
\gamma^i_s(z, \sigma_{\varepsilon}^{-i}, \sigma^i) - \varepsilon \leq u^i(z) \leq \gamma^i(z, \sigma_{\varepsilon}) + \varepsilon \quad \text{and} \quad \bar{\gamma}^i(z, \sigma_{\varepsilon}^{-i}, \sigma^i) - \varepsilon \leq u^i(z) \leq \bar{\gamma}^i(z, \sigma_{\varepsilon}) + \varepsilon.
\]

A uniform-l-average \( \varepsilon \)-equilibrium strategy of the non-zero-sum continuous-time stochastic game is an admissible strategy profile \( \sigma_{\varepsilon} \) for which there is a sufficiently long finite horizon \( s_{\varepsilon} \) such that for every initial state \( z \), every player \( i \in N \), every \( s \geq s_{\varepsilon} \), and every strategy \( \tau^i \) of player \( i \), we have
\[ γ^I_i(z, σ_e, γ^I_i) - \varepsilon \leq γ^I_i(z, σ_e) \leq γ^I_i(z, σ_e) + \varepsilon \quad \text{and} \]
\[ γ^I_i(z, σ_e, γ^I_i) - \varepsilon \leq γ^I_i(z, σ_e) \leq γ^I_i(z, σ_e) + \varepsilon. \]

A uniform-\(l\)-average equilibrium payoff of the non-zero-sum continuous-time stochastic game is a vector \( u = (u^I_i(z))_{i \in N, z \in \mathbb{S}} \in \mathbb{R}^{2 \times N} \) that is a uniform-\(l\)-average \( \varepsilon \)-equilibrium payoff for every \( \varepsilon > 0 \).

Obviously, a uniform-\(l\)-average equilibrium payoff is both a uniform equilibrium payoff and a limiting-average equilibrium payoff.

1.9. Robust equilibria

Our main objective is the study of approximate\(^3\) equilibria (of discounted continuous-time stochastic games) that are insensitive to a small imprecision in the specification of players' evaluations of streams of payoffs.

The interest in such a study stems from the fact that in most real-life interactions, the exact duration and/or the exact discounting rate is unknown (and definitely not commonly known).

Moreover, equilibrium analysis of game models whose duration is known, but not commonly known, to all players leads to results that are completely different from those of the model with a fixed finite duration; see Neyman (1999).

For each continuous-time stochastic game \( \Gamma \), we wish to identify a large family of sequences of profiles \( \bar{w}_k \) of measures for which the following condition holds.

For every \( \varepsilon > 0 \) there is a pair consisting of a strategy profile \( \sigma \) and a vector payoff \( \nu \) such that for \( k \) sufficiently large \( \sigma \) is an \( \varepsilon \)-equilibrium of \( \bar{w}_k \) with a payoff within \( \varepsilon \) of \( \nu \). If the condition holds for a sequence of profiles \( \bar{w}_k \) we say that \( \Gamma \) has \( (\bar{w}_k) \)-robust-equilibria.

The sequence of profiles \( \bar{w}_k \) of measures might be interpreted as an attempt to construct “better and better” approximations to a desired profile of measures, or to a small imprecision in the profile of measures that define the time-separable payoffs of the players.

The “better and better” approximations are subject to the usual caveats of converging sequences. The convergence can be with respect to a metric or a topology. Note that the smaller the metric, respectively, the weaker the topology, the larger the set of converging sequences of profiles of measures.

Before turning to our definition of converging sequences and the corresponding concept of robust equilibrium, we introduce a simple example of robustness.

Let \( \bar{w}_k \) be a sequence of profiles of measures on \([0, \infty)\) such that \( \bar{w}^I_i \) is, for every \( i \), a Cauchy sequence in the norm distance.\(^4\) Note that the sequence of payoffs \( γ^I_i(z, \sigma) \) converges uniformly in \( \sigma \) and that it is easy to prove that for any profile of measures \( \bar{w} \) on \([0, \infty)\) and \( \varepsilon > 0 \), a continuous-time stochastic game \( \Gamma \) (with finitely many states and actions) has an \( \varepsilon \)-equilibrium. Therefore, for sufficiently large \( k \), an \( \varepsilon \)-equilibrium \( \sigma \) of \( \bar{w}_k \) is, for every \( l \geq k \), a \( 2\varepsilon \)-equilibrium of \( \bar{w}_l \) with a payoff within \( \varepsilon \) of \( γ^I_i(z, \sigma) \). Therefore \( \Gamma \) has \( (\bar{w}_k) \)-robust equilibria.

We continue with the introduction of a few examples of sequences of profiles of measures. Let us state upfront that our results prove that a continuous-time stochastic game with finitely many states and actions has \( (\bar{w}_k) \)-robust equilibria for each one of the sequences \( (\bar{w}_k) \) that appear in Examples 1–5 below.

**Example 1.** Fix a profile \( \bar{w} \) of probability measures that are supported on \([0, \infty)\). Consider the family of all sequences of profiles of probability measures \( \bar{w}_k \) such that \( \bar{w}_k([0, t]) \to \bar{w}([0, t]) \) for every point \( t \) with \( \bar{w}(t) = 0 \).

**Example 2.** The sequence of profiles \( \bar{w}_k \) of measures, where \( w^j_k \) is the uniform probability measure on \([0, k)\). Note that for this sequence of profiles of measures, the statement that \( \Gamma \) has a uniform equilibrium payoff is equivalent to the statement that \( \Gamma \) has \( (\bar{w}_k) \)-robust equilibria.

**Example 3.** The family of sequences of profiles \( \bar{w}_k \) of measures, \( \rho^j_k \downarrow 0 \) \( \forall i \), where \( \rho^j_k \) \( (t, \infty)) = e^{-\rho^j_k t} \).

**Example 4.** The family of sequences of profiles \( \bar{w}_k \) of measures such that \( \forall i \), \( w^j_k([a + c, b + c]) \) is nonincreasing in \( c \) (\( a < b < \infty \)) and \( w^j_k([0, 1]) \to 0 \).

Note that each one of the sequences of the profiles of measures in Examples 2 and 3 is in the family of sequences described in Example 4.

\(^3\) It is impossible to find a strategy profile (or a vector payoff) that is an exact equilibrium (or an exact equilibrium payoff) in all sufficiently small changes in the payoff valuations.

\(^4\) The norm distance between two probability measures \( P \) and \( Q \) on a measurable space \((X, X')\) is defined by \( \|P - Q\| := \sup_{B \in X'} |P(B) - Q(B)| \). If \( X \) is finite (or countable) and \( X' \) consists of all subsets of \( X \), then \( \|P - Q\| = \sum_{x \in X} |P(x) - Q(x)| \).
Example 5. The family of sequences of profiles $\tilde{W}_k$ of measures that are supported on $[0, \infty)$ and such that $\forall i, w^i_{k}(\alpha + c, b + c)$ is nonincreasing in $c$ ($\forall 0 \leq a < b < \infty$) and $w^i_{k}(t, \infty) \to k \to \infty \theta^i e^{-\rho t}$, where $\rho$ is a fixed profile of nonnegative numbers $\rho^i \geq 0$ and $\theta$ is a fixed profile with $0 \leq \theta^i \leq 1$.

Note that each one of the sequences of the profiles of measures in Example 3 is in the family of sequences described in Example 5.

Our robustness results will require that the sequence $\tilde{w}_k$ $w^*$-converge; i.e., for every continuous function $[0, \infty] \ni t \mapsto g(t) \in \mathbb{R}$ the integrals $\int_0^\infty g(t) dw^i_k(t)$ converge as $k \to \infty$.

In a $w^*$-converging sequence of measures $\tilde{w}_k$ on $[0, \infty)$, part of the mass of $w^i_k$ may be pushed to infinity. Therefore, there need not be a profile $\tilde{w}$ of measures on $[0, \infty)$ that represent the limit.

The limit is represented by a profile $\tilde{w}$ of measures on $[0, \infty]$ such that for every player $i$ and every continuous function $g : [0, \infty) \to \mathbb{R}$, we have $\int_0^\infty g(t) dw^i_k(t) \to k \to \infty \int_0^\infty g(t) dw^i(t)$.

We equip the space of measures $W$ on $[0, \infty]$ with the minimal topology (called the $w^*$-topology) such that for every continuous function $[0, \infty] \ni t \mapsto f_t$, the function $W \ni w \mapsto \int_0^\infty f_t dw(t)$ is continuous. $W := W^N$ is equipped with the product topology.

Given $\tilde{w} \in W$, we say that $\Gamma$ has $\tilde{w}$-robust equilibria if there is a vector payoff $v$ (called a $\tilde{w}$-time-separable-robust equilibrium payoff) such that for every $\varepsilon > 0$ there is a strategy profile $\sigma$ (called a $\tilde{w}$-robust $\varepsilon$-equilibrium strategy) and a neighborhood $U$ of $\tilde{w}$, such that for every $u \in U$, $\sigma$ is an $\varepsilon$-equilibrium of $\Gamma_{\tilde{w}}$ with a payoff within $\varepsilon$ of $v$.

A discounting measure is a measure $w \in W$ with $w((a + c, b + c)$ nonincreasing in $c \geq 0$ for every $0 \leq a < b < \infty$. A discounting measure $w$ is absolutely continuous on $(0, \infty)$ with $\frac{dw}{dt}(t)$ nonincreasing. It may have atoms at 0 and $\infty$. Small values of $w((0, 1))/w((0, \infty))$ correspond to valuations of patient players. The space of all discounting measures is denoted by $W_d$.

Given $\tilde{w} \in W_d := (W_d)^N$, we say that $\Gamma$ has $\tilde{w}$-discounting-robust equilibria if there is a vector payoff $v$ (called a $\tilde{w}$-discounting-robust equilibrium payoff) such that for every $\varepsilon > 0$ there is a strategy profile $\sigma$ called a $\tilde{w}$-discounting-robust $\varepsilon$-equilibrium (strategy) and a neighborhood $U$ of $\tilde{w}$, such that for every $u \in U \cap W_d$, $\sigma$ is an $\varepsilon$-equilibrium of $\Gamma_{\tilde{w}}$ with a payoff within $\varepsilon$ of $v$.

The continuous-time stochastic game $\Gamma$ has discounting-robust equilibria if for every $\tilde{w} \in W_d$, $\Gamma$ has $\tilde{w}$-discounting-robust equilibria.

1.10. The results

The main result is that every continuous-time stochastic game with finitely many states and actions has discounting-robust equilibria.

The set $W^*_d$ of all profiles of probability discounting measures is infinite. However, as it is a compact subset of $\tilde{W}$, the existence of discounted-robust equilibria implies (and is equivalent to) the following finiteness result.

For every $\varepsilon > 0$, there is a finite open cover $(U_{\alpha})_{\alpha \in A}$ of $W^*_d$, a finite list of strategy profiles $(\sigma_{\alpha})_{\alpha \in A}$, and a finite list of vector payoffs $(v_{\alpha})_{\alpha \in A}$, such that for every profile of discounting measures $w$ in $U_{\alpha}$, $\sigma_{\alpha}$ is an $\varepsilon$-equilibrium of $\Gamma_{\tilde{w}}$ with a payoff within $\varepsilon$ of $v_{\alpha}$.

The discounting measure can vary among players. Therefore the main result applies also to cases where some of the players are extremely patient (e.g., a government) and other players’ patience can vary (e.g., as a function of each player’s uncertain life horizon). The discounting measure of player $i$, $w^i$, may have a positive mass at infinity, representing the weight of the discounting measure on the distant future.

If both $w^i((0, \infty))$ and $w^i((\infty))$ are positive, the discounting measure $w^i$ represents a valuation that is an average of a classical (not necessarily stationary) discounting valuation and a valuation of an extremely patient player.

The existence of $\tilde{w}$-discounting-robust equilibria, where each discounting measure $w^i$ is supported on $[\infty]$, implies the existence of a uniform-$l$-average equilibrium payoff.

The existence of a uniform (or a limiting-average) equilibrium payoff in any continuous-time stochastic game with finitely many states and actions is in sharp contrast to our knowledge about uniform equilibrium in discrete-time stochastic games, where it is still unknown whether all discrete-time stochastic games with finitely many states and actions have a uniform (or a limiting-average) equilibrium payoff.

A continuous-time stochastic game with finitely many states and actions need not have a $\tilde{w}$-robust equilibrium payoff for every $\tilde{w} \in W$. However, if the profile of measures $\tilde{w} \in W$ is supported on $[0, \infty)$, then it has a $\tilde{w}$-robust equilibrium.

Earlier studies of continuous-time stochastic games include Zachrisson (1964), Lai and Tanaka (1982), and Guo and Hernandez-Lerma (2003, 2005). An important difference between the continuous-time game model of the present paper and those of the other papers is that in the present paper we consider all strategies in the continuous-time game, while each one of the other papers confines the strategy space to a subclass of all strategies.

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5 The assumption of finitely many states is essential even in the case of a single player; see, e.g., Neyman (2003b, Section 1.4). The assumption of finitely many actions is essential even in the two-person zero-sum case with finitely many states; see Vigeral (2012).
On the other hand, the present paper focuses on the model with finitely many states and actions, while the earlier papers include also more general action and state spaces. Below, we briefly mention the contributions of the above-mentioned papers to continuous-time stochastic games with finitely many states and actions.

Zachrisson (1964) studies two-person zero-sum finite-horizon stochastic games with a terminal payoff, where players are confined to Markov strategies, and proves the existence of the value and optimal strategies.

Lai and Tanaka (1982) study non-zero-sum discounted games, where players are confined to Markov strategies, and prove the existence of stationary equilibria.

Guo and Hernandez-Lerma (2003) study a subclass of two-person zero-sum stochastic games with the average payoff criterion, where players are confined to continuous Markov strategies (see Section 2.1 for the definition), and prove the existence of the value and optimal strategies.

Guo and Hernandez-Lerma (2005) study two-person non-zero-sum discounted stochastic games, where players are confined to continuous6 Markov strategies, and prove the existence of stationary equilibria.

For completeness, we derive the existence of stationary equilibrium in the continuous-time non-zero-sum discounted stochastic game (without the restriction to continuous Markov strategies and more generally without the restriction to (memoryless) Markov strategies).

2. The model

A continuous-time stochastic game (with finitely many states and actions) is defined by: a finite set of players \( N \); a finite set of states \( S \); for each \( z \in S \) and each player \( i \in N \) a finite set of actions \( A_i(z) \); a (vector-valued) payoff function \( g : A \to \mathbb{R}^N \), where \( A = \{(z, a) : a \in A(z)\} \) and \( A(z) = \times_{i \in N} A_i(z) \); and for each \( z' \neq z \in S \) and \( a \in A(z) \) a real-valued transition rate \( \mu(z', z, a) \geq 0 \). The \( i \)-th component of \( g \) is denoted by \( g_i \).

For notational convenience we set \( A^i = \cup_{z \in S} A^i(z) \) and \( A = \times_{i \in N} A^i \). The \( i \)-th coordinate, \( i \in N \), of \( a \in A(z) \) is denoted by \( a^i \).

The set \( A^i(z) \) represents the set of feasible actions of player \( i \) when the state is \( z \). An element \( a \in A(z) \), \( a = (a^i)_{i \in N} \) with \( a^i \in A^i(z) \), is called an action profile.

The interpretation of the payoff function and of the transition rates is that when the state is \( z \in S \) and players play the action profile \( a \in A(z) \) during the infinitesimal time \( dt \), then the payoff to player \( i \) is \( g^i(z, a) dt \) and the state moves to state \( z' \neq z \) with probability \( \mu(z', z, a) dt \) and stays in state \( z \) with probability \( 1 + \mu(z, z, a) dt \), where \( \mu(z, z, a) := -\sum_{z' \neq z} \mu(z', z, a) \).

A pure play of the continuous-time stochastic game is a measurable function \( h : [0, \infty) \to S \times A \), \( t \mapsto h(t) = (z_t, a_t) \), with \( a_t \in A(z_t) \), and \( t \mapsto z_t \) right continuous.

A pure play \( h \) is identified with a pair of functions \( h^5 : [0, \infty) \to S \), which is the restriction of \( h \) to the first coordinate, \( h^5 : [0, \infty) \to A \), which is the restriction of \( h \) to the second coordinate.

Given a pure play \( h \) we define \( h_t \) as the restriction of the first coordinate of \( h \) to the time interval \([0, t]\) and the restriction of the second coordinate to \([0, t]\). \( h_t \) is identified with the pair \((h_{[0,t]}^5, h_{A_{[0,t]}}^5)\), where \( h_{[0,t]}^5 \) is the restriction of \( h^5 \) to the time interval \([0, t]\) and \( h_{A_{[0,t]}}^5 \) is the restriction of \( h^5 \) to the time interval \([0, t]\).

The unnormalized, respectively, normalized, \( \rho \)-discounted payoff \( (\rho > 0) \) of a pure play \( h \) is \( \int_0^\infty e^{-\rho t} g(z_t, a_t) dt \), respectively, \( \int_0^\infty e^{-\rho t} g(z_t, a_t) dt \). The \( s \)-horizon normalized payoff of a pure play\(^7 \) \( h \) is \( \frac{1}{s} \int_0^s g(z_t, a_t) dt \).

A pure play \( h \) of the continuous-time stochastic game is a (measurable) function \( h : [0, \infty) \to S \times \Delta(A) \), \( t \mapsto h(t) = (z_t, x_t) \), with \( x_t \in \Delta(A(z_t)) \), where \( \Delta(\star) \) denotes all probabilities on the set \( \star \).

Given a play \( h \) we define \( h_t \), as we did for the pure play, i.e., as the restriction of the first coordinate of \( h \) to the time interval \([0, t]\) and the restriction of the second coordinate to \([0, t]\).

The \( \rho \)-discounted payoff of a play \( h \) is \( \int_0^\infty e^{-\rho t} g(z_t, x_t) dt \), where \( g(z, x) = \sum_{a \in A(z)} x(a) g(z, a) \) is the linear extension of \( g \).

The \( s \)-horizon normalized payoff of a play \( h \) is \( \frac{1}{s} \int_0^s g(z_t, x_t) dt \).

2.1. Strategies in continuous-time games

In this section we define and discuss several classes of strategies: stationary strategies, Markov strategies, discretized strategies, and general strategies.

Each one of these classes of strategies is a proper subclass of the next one in the list. In each equilibrium existence result we desire that the equilibrium strategies be as simple as possible.

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6 The confinement there is more severe, as it requires the strategy of a player to be a Markov strategy such that for any Markov strategy of the other player the transition rates are continuous. See Guo and Hernandez-Lerma (2005, Definition 3.1). However, unless the game is degenerate, there are no such strategies. Nevertheless, the proofs there hold also for continuous Markov strategies.

7 Much of the theory developed remains intact even if the integral \( \int_0^t f(t) dt \) of a real-valued bounded function \( f \) refers to a fixed monotonic linear functional (on the space of real-valued bounded functions), with \( \int_0^t f(t) dt = c(b - a) \) and \( \int_0^t f(t) dt = \int_0^a f dt + \int_a^b f dt \) for all \( 0 \leq a, b, c \). This defines the discounted and \( s \)-horizon payoffs over all plays, not necessarily measurable ones.
However, one has to recall that the equilibrium strategy profile should be such that there is no unilateral beneficial deviation, and a player can deviate to any general strategy. Therefore, an equilibrium strategy profile need be such that, in particular, any unilateral deviation to a general strategy yields a strategy profile that defines a probability distribution on plays.

2.1.1. Markov strategies

Let $X^i(z), X(z), X^i, and X$ denote all probability distributions over $A^i(z), A(z), A^i, and A$, respectively.

A Markov strategy of player $i$ is defined by a measurable function $\sigma^i : S \times [0, \infty) \to X^i$ with $\sigma^i(z, t) \in X^i(z)$. A profile $\sigma$ of Markov strategies $(\sigma^i)_{i \in N}$ defines a Markov strategy profile $\sigma : S \times [0, \infty) \to X$ with $\sigma(z, t) \in X(z)$ by $\sigma(z, t)[a] = \prod_{i \in N} \sigma^i(z, t)[a^i]$.

A Markov correlated strategy is defined by a measurable function $\sigma : S \times [0, \infty) \to X$ with $\sigma(z, t) \in X(z)$. Note that a profile of Markov strategies is a special case of a Markov correlated strategy.

A Markov strategy $\sigma^i$ of player $i$ is continuous if $\sigma^i(z, t)$ is continuous in $t$. Similarly, a Markov correlated strategy $\sigma$ is continuous if $\sigma(z, t)$ is continuous in $t$.

A stationary strategy of player $i$ is a Markov strategy $\sigma^i$ where $\sigma^i(z, t)$ is independent of $t$.

A profile $\sigma$ of Markov strategies, or more generally, a Markov correlated strategy $\sigma$, and an initial state $z_0$ define on the space of right-continuous functions $t \mapsto z_t \in S$ a unique probability distribution $P_\sigma$ that for every $z \neq z_t$ satisfies the equality

$$P_\sigma(z_{t+\delta} = z \mid h_t) = \int t^{\delta} \mu(z, z_t, \sigma(z_t, s)) ds + o(\delta),$$

where $\mu(z', z, \chi) := \sum_{z' \in A} \mu(z', z, \chi)(z')$.

This unique distribution is supported on the space $H^k$ of all functions $t \mapsto z_t$ that have left limits and are right continuous at every $t \geq 0$. (The fact that with $P_\sigma$ probability 1, the function $t \mapsto z_t$ has left and right limits at any $t$ follows from condition (2) and the right continuity follows from our assumption.)

In order to define the probability $P_\sigma$ on the space of right continuous functions $t \mapsto z_t$, it suffices to define the distribution of $(s_k, z_{sk})_{k \geq 0}$ where $s_k$ is the time of the $k$-th state change.

Conditional on $(s_k, z_{sk})_{k \leq k}$ the distribution of $s_{k+1} - s_k$ is given by

$$P_\sigma(s_k + 1 = s_k \geq a \mid (s_k, z_{sk})_{k \leq k}) = \frac{\mu^a(s_k, z_{sk}, \sigma(z_{sk}, s)) ds}{\mu(s_k, z_{sk}, \sigma(z_{sk}, s))},$$

and $(\mu(z', z, \chi)(a))$ the Lebesgue density theorem conditional on $(s_k)_{k \leq k+1}$ and $(z_{sk})_{k \leq k}$, the distribution of $z_{sk+1}$ is given by

$$P_\sigma(z_{sk+1} = z \mid (s_k)_{k \leq k+1}, (z_{sk})_{k \leq k}) = \frac{-\mu(z, z_{sk}, \sigma(z_{sk}, s)) ds}{\mu(z_{sk}, z_{sk}+1), \sigma(z_{sk}, s)) ds} \ a.e.$$

The space of mixed strategies that are mixtures of Markov strategies depends on the measurable structure on the space of Markov strategies. Obviously, there are many measurable structures on the space of Markov strategies. One of them is derived from the minimal topology for which, for every real-valued continuous function $f$ on $S \times X^i$ and all $0 \leq a < b < \infty$, the function $\sigma^i \mapsto \int_a^b f(z, \sigma^i(z, t)) dt$ is continuous.

2.1.2. Discretized strategies

Fix a strictly increasing sequence of times $T = (t_k)_{k \geq 0}$ with $t_0 = 0$ and $\mathbb{R} \ni t_k \to \infty$ as $k \to \infty$.

A $T$-discretized strategy of player $i$ specifies for each $k$ and $h_k \in H_k$ a function $\sigma^i_{h_k} : S \times [t_k, t_{k+1}) \to X^i$ with $\sigma^i_{h_k}(z, t) \in X^i(z)$.

The interpretation of this description of $\sigma^i$ is that at time $t_k \leq t < t_{k+1}$ the strategy $\sigma^i$ selects an action $\sigma^i_{h_k}(z, t) \in X^i(z)$ as a function of the state $z_t$ at time $t$ and the history of play $h_k \in H_k$ up to time $t_k$.

A profile $\sigma = (\sigma^i)_{i \in N}$ of $T$-discretized strategies defines for each $k$ and $h_k \in H_k$ a function $\sigma_{h_k} : S \times [t_k, t_{k+1}) \to X$ with $\sigma_{h_k}(z, t) := \bigotimes_{i \in N} \sigma^i_{h_k}(z, t) \times \mathbb{X} \in \mathbb{X}(X(z))$ (where $\otimes$ stands for the product of measures).

$\sigma_{h_k}(z, t)$ specifies for each $k$ and $h_k \in H_k$, a function $\sigma_{h_k} : S \times [t_k, t_{k+1}) \to X$ with $\sigma_{h_k}(z, t) \in X(z)$. Note that a profile of $T$-discretized strategies is a special case of a $T$-discretized correlated strategy.

A $T$-discretized correlated strategy (with suitable measurability assumptions) defines (for each initial state $z_0$) a (unique) probability distribution $P_{\sigma}$ on plays (with right continuous state function $t \mapsto z_t$) such that for $t_k \leq t < t_{k+1}$ and $z \in S$,

$$x_t = \sigma_{h_k}(z_t, t)$$

and

$$P_{\sigma}(z_{t+\delta} = z \mid h_t) = 1_{z = z_0} + \int t^{\delta} \mu(z, z_t, \sigma_{h_k}(z_t, s)) ds + o(\delta),$$

where $1_{z = z_t} = 1$ if $z = z_t$, and $1_{z = z_t} = 0$ if $z \neq z_t$. 

In all our results, the profile of equilibrium strategies will be a profile of \( T \)-discretized strategies (and recall that \( T \) is a sequence of real numbers).

However, a deviation of a player (even in the single-player case) need not be a \( T^* \)-discretized strategy with \( T^* \) being a sequence of real numbers.

For example, the strategy that for \( t \geq s_1 \) plays one action \( a_t = a \), and for \( t < s_1 \) plays another action \( a_t = b \neq a \), is not a \( T \)-discretized strategy w.r.t. a strictly increasing sequence \( T = (t_k)_{k \geq 0} \) of real numbers.

Let us show, informally, that a profile of strategies that is obtained by a unilateral deviation from a profile of \( T \)-discretized strategies defines a unique probability on plays.

The derivation is not formal as we have not yet introduced the definition of general strategies. However, aside from the needed suitable measurability assumptions, the general form of the behavior of a strategy of player \( i \) when confronted with \( T \)-discretized strategies of the other players should be clear.

Following our definition (in the next section) of general strategies, it will be clear that the assumed behavior of a general unilateral deviation holds for any general strategy, and therefore, with the suitable measurability assumptions, the informal arguments below turn into a formal proof.

Let \( \sigma \) be a profile of \( T \)-discretized strategies, where \( T = (t_k)_{k \geq 0} \) is an increasing sequence of real numbers with \( t_0 = 0 \) and \( t_k \to k \to \infty \). 

Let \( s_0 = 0 \) and let \( s_l \) be the time of the \( l \)-th state change. If there are fewer than \( l \) state changes then \( s_l = \infty \). Note that \( s_l \) is a function of the play.

Define \( t^l_k = s_l \lor (s_l+1 \land t_k) \), where \( a \lor b \) stands for \( \max(a, b) \) and \( a \land b \) stands for \( \min(a, b) \). Note that \( s_l \leq t^l_k \leq s_{l+1} \) and \( t^l_k \to k \to \infty s_{l+1} \).

Assume that player \( i \) is deviating to a strategy \( \tau^i \). By induction on \( l \) assume that the distribution of \( h_{s_l} \) is uniquely defined by the strategy profile \((\sigma^{-i}, \tau^i)\). We will show that this assumption implies that the distribution of \( h_{s_{l+1}} \) is uniquely defined.

The first step shows that for \( s_l \leq t < s_{l+1} \), the action profile \( x_t \) is a well-defined function of \( h_{s_l} \) and the strategy profile \((\sigma^{-i}, \tau^i)\).

For any \( j \neq i \), the strategy \( \sigma^j \) is a \( T = (t_k)_{k \geq 0} \)-discretized strategy. Therefore, for any \( t^l_k \leq t < t^l_{k+1} \), \( x^l_t \) is well defined as a function of \( h^l_k \) and \( \sigma^j \).

Therefore, for any \( t^l_k \leq t < t^l_{k+1} \), \( x^l_t = (x^l_t)^{(j) \neq i} \) is well defined as a function of \( h^l_k \) and \( \sigma^{-i} = (\sigma^j)^{(j) \neq i} \).

As \( x_t^{\tau^i} \) is well defined for all \( t^l_k \leq t < t^l_{k+1} \) as a function of \( h^l_k \) and \( \sigma^{-i} \), player \( i \) does not learn any new information from the actions \( x_t^{\tau^i} \leq t \leq t^l_{k+1} \). Therefore, \( x_t^{\tau^i} \) is well defined as a function of \( h^l_k \).

Therefore, for any \( t^l_k \leq t < t^l_{k+1} \), the action profile \( x_t \) is well defined as a function of \( h^l_k \) and the strategy profile \((\sigma^{-i}, \tau^i)\).

By induction on \( k \), the action profile \( x_t \) is well defined for every \( s_l \leq t < s_{l+1} \) as a function of \( h_{s_l} \) and the strategy profile \((\sigma^{-i}, \tau^i)\).

For \( s_l \leq t \), we denote by \( x_t \) the action profile defined by the strategy profile \((\sigma^{-i}, \tau^i)\) at time \( t \) conditional on \( s_{l+1} > s_l \). Note that the action profile \( x_t \) is a well-defined function of \( h_{s_l} \) and the strategy profile \((\sigma^{-i}, \tau^i)\).

This enables us to define the action profile \( x_t \) at time \( s_l \leq t < s_{l+1} \) and the conditional distribution of \((s_{l+1}, z_{s_{l+1}})\) as a function of \( h_{s_l} \) and the strategy profile \( \sigma := (\sigma^{-i}, \tau^i) \), as follows.

For \( s_l \leq t < s_{l+1} \),

\[
x_t = x_t^{s_l},
\]

\[
P_{\sigma} (s_{l+1} \geq s_l + \theta | h_{s_l}) = e^{-\mu (z_{s_l}, x_t^{s_l}, x_t^{s_l})} d\theta,
\]

and (using the Lebesgue density theorem) for \( z' \neq z_{s_l} \),

\[
P_{\sigma} (z_{s_{l+1}} = z' | h_{s_l}, s_{l+1}) = \frac{-\mu (z', x_{s_l}, x_{s_{l+1}})}{\mu (z', z', x_{s_{l+1}})} \text{ a.e.}
\]

By induction on \( l \) we deduce, using Tulcea’s extension theorem (see, e.g., Struik and Varadhan, 2006, Theorem 1.1.9), that the strategy profile \((\sigma^{-i}, \tau^i)\) defines a unique probability distribution on plays.

2.1.3. General strategies

The space of all plays is denoted by \( H \). The space of all (measurable) functions \( h : [0, \infty) \to h(t) = (z_t, x_t^{i-}), \) with \( x_t^{i-} \in \Delta(\times_{N \neq i} A^i) \) and \( z_t \in S \), is denoted by \( H^{-i} \).

A strategy \( \sigma^i \) of player \( i \) is a (measurable) function \( \sigma^i : H^{-i} \times [0, \infty) \to X^i \) with \( \sigma^i(h, t) \in X^i(z_t) \) and such that for all triples \((h', h, t) \in H^{-i} \times H^{-i} \times [0, \infty) \) with \( h'_i = h_i \) we have \( \sigma^i(h, t) = \sigma^i(h', t). \) This last condition asserts that a player cannot
decide his current mixed action based on future events. Note that a discretized strategy is a special case of a (general) strategy.

A play \( h \in H \) is compatible with the strategy profile \( \sigma = (\sigma^t)_{t \in \mathbb{N}} \) if \( x_t^i = \sigma^t(h^{-1}, t) \) for every \( t \geq 0 \).

A correlated strategy \( \sigma \) is a (measurable) function \( \sigma : H^S \times [0, \infty) \rightarrow X \) with \( \sigma(h, t) \in X(z_t) \) and such that for all triples \((z_t^i, z_t^j, z_t^k) \in H^S \times H^S \times [0, \infty)\) with \( z_t^i, z_t^j, z_t^k \) we have \( \sigma(t) = \sigma((z_t^i, z_t^j, z_t^k)). \)

A strategy profile need not define unambiguously a probability distribution over plays; see, e.g., Proposition 1.

An admissible profile of strategies \( \sigma = (\sigma^t)_{t \in \mathbb{N}} \) is a strategy profile \( \sigma \) such that for every player \( i \) and every strategy \( t^i \) of player \( i \) the strategy profile \( (\sigma^{-i}, t^i) \) defines a (unique and thus unambiguously) probability distribution \( \rho_{\sigma^{-i}, t^i} \) on plays (with right continuous state function \( t \mapsto z_t \)).

Throughout, we have used measurability assumptions in our definitions. The meaning of a play and of a strategy depends on the \( \sigma \)-algebras that define measurability. However, we skipped the spelling out of the \( \sigma \)-algebras for the following reason.

The space of measurable functions from a measurable space \((X, \mathcal{X})\) to a measurable space \((Y, \mathcal{Y})\) is monotonically increasing in \( \mathcal{X} \) (namely, for each fixed space \( X \) and measurable space \((Y, \mathcal{Y})\), the larger the \( \sigma \)-algebra \( \mathcal{X} \), the larger is the space of measurable functions from the measurable space \((X, \mathcal{X})\) to the measurable space \((Y, \mathcal{Y})\) and monotonic decreasing in \( \mathcal{Y} \).

In constructing a strategy profile in an existence result, it is desirable that this strategy be in a space of strategies whose meaning, implementation, and interpretation are as simple as possible. To this end, we wish to take a minimal reasonable \( \sigma \)-algebra on the space of plays in the domain of a strategy, and a maximal reasonable \( \sigma \)-algebra in the image of a strategy.

On the other hand, when we wish to demonstrate that there are no unilateral beneficial deviations, it is desirable to demonstrate it for the most general deviation. To this end, we wish to take a maximal reasonable \( \sigma \)-algebra on the space of plays in the domain of a strategy, and a minimal reasonable \( \sigma \)-algebra in the image of a strategy.

2.2. Discussion of strategies

The realization of an increasing sequence of real numbers is a well-ordered subset of the reals. This enables us to define unambiguously the distribution on plays that is defined by a profile of strategies that is obtained by a unilateral deviation from a profile of discretized strategies.

In this section we illustrate the difficulties that arise when the set of decision times is not well ordered.

2.2.1. Prelude: the set-theoretic setup

Consider a totally ordered set \( T \). For every element \( t \in T \) let \( T_t := \{t' \in T : t' < t\} \). Fix a set \( A \) with at least two elements.

If we interpret the set \( T \) as the set of action times, and \( A \) as the set of single-stage action profiles, then a play is an element \( h \in A^T \), and the value of \( h \) at time \( t \), \( h(t) \), is interpreted as the action profile at time \( t \). A history of play up to (and not including) time \( t \) is an element \( A^{T_t} \), and given a play \( h \) we denote by \( h_t \) its restriction to \( A^{T_t} \).

A local pure strategy profile \( \sigma \) is a list of functions \( \sigma_t : A^{T_t} \rightarrow A, t \in T \). The integral of a local pure strategy profile \( \sigma \) is the set of all plays \( h \) such that \( h(t) = \sigma_t(h_t) \) for every \( t \). A local pure strategy profile is integrable if its integral is nonempty; i.e., there is at least one play \( h \in A^T \) such that for every \( t \in T \) we have \( h(t) = \sigma_t(h_t) \).

The following simple proposition demonstrates that when the set of times is not well ordered then there are local strategies that are not integrable, and there are local strategies whose integral contains more than one element. On the other hand, if the set of times is well ordered, the integral contains a single element.

**Proposition 1.** The following conditions are equivalent: 1) every local pure strategy profile \( \sigma \) is integrable, 2) the set of times \( T \) is well ordered, and 3) the integral of every local pure strategy contains at most one element.

**Proof.** Assume that \( T \) is well ordered. We define for every \( t \in T \) the action profile \( h(t) \) by transfinite induction: (1) for the first element \( t^* \) of \( T \), \( h_{t^*} = \emptyset \) and therefore we set \( h(t^*) = \sigma_{t^*}(\emptyset) \); (2) if \( h(t^i) \) has been defined for every \( t^i < t \), then \( h_t \) has been defined and we set \( h(t) = \sigma_t(h_t) \). As \( T \) is well ordered the action profile \( h(t) \) is defined uniquely for every \( t \in T \).

Therefore, the play \( h \) is the unique element in the integral of \( \sigma \).

If \( T \) is not well ordered, there is an infinite decreasing sequence \( t_1 > t_2 > \ldots \) of times in \( T \). Let \( a, b \) be two distinct elements of \( A \). Define the local pure strategy profile \( \sigma \) by \( \sigma_t(a) = a \) if \( t \not\in \{t_j : j \geq 1\} \), \( \sigma_t(b) = b \) if \( \limsup_{k \to \infty} |i < j \leq i + k : h(t_j) = a|/k < 1/2 \), and \( \sigma_t(a) = a \) if \( \limsup_{k \to \infty} |i < j \leq i + k : h(t_j) = a|/k \leq 1/2 \). Note that for every play \( h \), \( \limsup_{k \to \infty} |i < j \leq i + k : h(t_j) = a|/k \) is independent of \( i \). Therefore, if \( h \) is in the integral of \( \sigma \), \( h(t_j) \) is a constant and \( \sigma_t(h_t) = h(t) \), contradicting the local definition of \( \sigma \).

Define the local pure strategy \( \tau \) by \( \tau_t(h_t) = b \) if \( |j : i < j \land h(t_j) = a| < \infty \), and \( \tau_t(a) = a \) otherwise. The play that plays \( a \) everywhere, and the play that plays \( b \) at time \( t = t_j \), \( j \geq 1 \), and \( a \) elsewhere, i.e., at times \( t \not\in \{t_j : j \geq 1\} \), are in the integral of \( \tau \).

□
A delayed local pure strategy profile is a local strategy $\sigma$ such that there is a well-ordered\(^8\) subset $T^*$ of $T$ such that for every $t \in T$ that is not a maximal element of $T$ there is $t^* \in T^*$ such that $t^* \leq t < t^*$, where $t^*$ is the least element in $T^*$ that is $> t^*$, and for $t^* \leq t < t^*$ we have

$$\sigma_t(h_t) = \sigma_t(h^*_t)$$

whenever $h_{t^*} = h^*_{t^*}$.

**Proposition 2.** A delayed local pure strategy profile is integrable and its integral contains a unique element.

**Proof.** The restriction of the function $h(t)$ to the interval $t^* \leq t < t^*$ is defined by transfinite induction on $t^* \in T^*$. \hfill $\Box$

**Proposition 2** is standard and classical, and is used in the theory of differential games, when the set of decision times $T$ is either a finite interval $[0, T]$ or the nonnegative real numbers $[0, \infty)$.

2.3. Strategies that observe and select mixed actions

In discrete-time games with perfect monitoring we focus on strategies that observe the pure past actions and select a pure action. Below we discuss our choice of the model, where players observe past mixed actions and strategies select mixed actions as a function of past observed variables.

These assumptions are conceptually innocuous in the case where each player is a continuum of agents and the observable variables are the statistics of actions of the continuum of agents. In the continuous-time model, we find these assumptions natural even in the case where each player represents a single decision maker. This is motivated in part by the limitations on the players’ perception.

The perception of players is not without limitation. A person watching a sequence of blue and yellow pictures will be unable to tell the exact times when the blue ones were presented if the switching rate crosses some threshold. In fact, he will observe a greenish picture, and the greenness will change as a function of the fraction of time in which the blue pictures were presented. It is therefore desirable to model the observation of the past by the observation of time averages of pure actions, and time averages of pure actions correspond to mixed actions. Given this limitation on perception, it is also natural to model the interaction by assuming that the players select mixed actions.

These assumptions – of observing and selecting mixed actions – are also technically convenient. They result in a tractable analytic model. In addition, the analysis of this analytic model leads to analogous results in the asymptotic theory of discrete-time stochastic games with short-stage durations and in the (discretized approach) to continuous-time stochastic games where players are subject to a short delay in observing other players’ actions; see Neyman (2012, 2013).

3. Non-zero-sum continuous-time stochastic games

3.1. Useful inequalities

Fix a continuous-time stochastic game $\Gamma = (N, S, A, g, \mu)$. In this section we introduce several inequalities that are often used in the paper.

Recall that $H^S$ is the space of all functions $z : \Sigma \rightarrow S$ that are right continuous and have left limits everywhere. Let $\mathcal{H}^S$, respectively, $\mathcal{H}^Z$, be the $\sigma$-algebra of subsets of $H^S$ that is generated by all the $S$-valued random variables $z_s$, $s \leq t$, respectively, $s < \infty$.

The concept of a $T$-discretized strategy with respect to an increasing sequence of real numbers, which was defined earlier, extends naturally to the case where $T$ is an increasing sequence of $(H^S_k)_{k \geq 0}$-stopping times.

Fix a strictly increasing sequence of $(H^S_k)_{k \geq 0}$-stopping times $T = (t_k)_{k \geq 0}$ with $t_0 = 0$ and $t_k \rightarrow \infty$ as $k \rightarrow \infty$.

A special case of interest is where $T = (s_k)_{k \geq 0}$, where $s_0 = 0$ and $s_k \leq \infty$ is the time of the $k$-th state change.

In Section 2.1.2 we showed that if $\sigma$ is a profile of $T$-discretized strategies, where $T$ is an increasing sequence of real numbers, and $\tau^i$ is a strategy of player $i$, then, for $s_j \leq t < s_{j+1}$, the action profile $x_t$ is a well-defined function of $h_{s_j}$ and the strategy $\sigma^{-1}(\tau^i)$. Therefore, there is an $(s_k)_{k \geq 0}$-discretized correlated strategy $\tilde{\sigma}$ such that $\tilde{\sigma}$ and $(\sigma^{-1}, \tau^i)$ define the same distribution on plays.

The same property holds also in the case where each strategy $\sigma^i$ is a discretized strategy with respect to a strictly increasing sequence (or even a well-ordered countable set) of $(H^S_k)_{k \geq 0}$-stopping times.

The lemmas in this section assume that $\sigma$ is a $(s_k)_{k \geq 0}$-discretized correlated strategy. Therefore, the conclusions of the lemmas hold for any profile that is obtained by a unilateral deviation from a profile of discretized strategies.

For such a correlated strategy, the mixed action-profile at time $t < s_1$ is well defined as a function of $z_0$ (and $\sigma$) and is denoted by $x_{0,t}$. Note that $x_{0,t}$ is the mixed action-profile that is selected by the correlated strategy $\sigma$ at time $t$ conditional on no state change in the interval $[0, t]$. Note that $x_{0,t}$ is defined for every $t \geq 0$ as a function of $z_0$ (and $\sigma$).

---

\(^8\) A well-ordered subset of an ordered set is a subset such that every nonempty subset of it has a minimal element.
Using the notations that are used in the definition of a general correlated strategy, \( x_{0,t} = \sigma(h, t) \), where \( h_t \) is the unique play that is compatible with \( \sigma \) and \( z_s = z_0 \) for every \( s \leq t \).

Similarly, the mixed action-profile at time \( s_1 \leq t < s_2 \) is well defined as a function of \( (s_1, z_{s_1}) \) (and \( \sigma \)) and is denoted by \( x_{1,t} \). In other words, \( x_{1,t} \) is the mixed action-profile that is selected by the correlated strategy \( \sigma \) at time \( t \) conditional on \( t \geq s_1 \) and no state change in the interval \([s_1, t] \). Note that \( x_{1,t} \) is defined for every \( t \geq s_1 \) as a function of \( (s_1, z_{s_1}) \) and \( \sigma \).

Using the notations that are used in the definition of a general correlated strategy, for every \( (s_1, z_{s_1}) \) and \( t > s_1 \), \( x_{1,t} = \sigma(h, t) \), where \( h_t \) is unique play up to time \( t \) that is (1) compatible with \( \sigma \), (2) \( z_t = z_0 \) for every \( s < s_1 \), and (3) \( z_s = z_{s_1} \) for every \( s_1 < s \leq t \).

Let \( \|\mu\| := \max\{\mu(x, z, a)\} \). Obviously, \( \|\mu\| \) is finite as it is the maximum over finitely many real numbers.

The following Lemma shows that the probability of at least one state change in a time interval is bounded by a constant \( \|\mu\| \) times the length of the time interval, and that the probability of at least two state changes in a time interval is bounded by a constant \( \|\mu\|^2 \|\mu\| \) times the square of the length of the time interval.

The second bound implies that for any (finite or) countable collection of time intervals the probability of at least two state changes in one of these time intervals is bounded by a constant \( \|\mu\|^2 / 2 \) times the sum of the squares of the lengths of the time intervals.

**Lemma 1.** For every \( \theta > 0 \),

\[
P_{\sigma}(s_1 \geq \theta) \geq e^{-\|\mu\|\theta} \geq 1 - \theta \|\mu\| \tag{8}
\]

and

\[
P_{\sigma}(s_2 \leq \theta) \leq \theta^2 \|\mu\|^2 / 2. \tag{9}
\]

**Proof.** As \( P^z_{\sigma}(s_1 \geq \theta) = e^{\int_0^1 \mu(x, z, x_0, t) dt} \) and \( 0 \geq \int_0^1 \mu(x, z, x_0, t) dt \geq -\|\mu\| \theta \), \( P^z_{\sigma}(s_1 \geq \theta) \geq e^{-\|\mu\|\theta} \geq 1 - \theta \|\mu\| \), proving inequality (8).

The conditional probability, given \( (s_1, z_{s_1}) \), that \( s_2 \leq \theta - s_1 \) is equal, on \( s_1 \leq \theta \), to \( 1 - e^{\int_0^1 \mu(x, z, x_0, t) dt} \leq \|\mu\| \|\theta - s_1\| \).

The density of \( s_1 \) at \( \eta \in \mathbb{R}^+ \) is given by \( -\mu(x, z, x_1, \eta)e^{\int_0^1 \mu(x, z, x_0, t) dt} \leq \|\mu\| \).

Therefore, \( P_{\sigma}(s_1 \leq \theta) \leq \int_0^1 \|\mu\| \|\theta - s\| \|\mu\| ds = \|\mu\|^2 \theta^2 / 2, \) proving inequality (9). \( \square \)

Several proofs in the paper use backward induction arguments. In these arguments we will need to consider the expectation of the sum of the accumulation of payoffs in a (short) time interval and a function of the state at the end of the interval.

Without loss of generality we can assume that the time interval starts at time 0 and ends at time \( \delta \). Therefore, we are led to estimate the expectation of \( \int_0^T g_t^i dw(t) + u(z) \), where \( w \) is a nonnegative measure on \([0, \delta] \) and \( u \) is a real-valued function that is defined on the state space. The expectation depends (obviously) on the strategy profile, or, more generally, the correlated strategy \( \sigma \).

The next lemma uses the inequalities of Lemma 1 to approximate each one of the terms \( E_{\sigma} \int_0^T g_t^i dw(t) \) and \( E_{\sigma} u(z) \), and therefore also their sum, by functions that depend on the choice of the mixed action-profile of the correlated strategy \( \sigma \) conditional on no state change in the interval \([0, \delta] \). Recall that this choice is denoted by \( x_{0,t} \).

Recall that \( \|g\| = \max_{x, z, a} |g^i(x, z, a)| \).

**Lemma 2.** Let \( \delta > 0 \), \( w \) be a finite nonnegative measure on \([0, \delta] \), and \( u : S \rightarrow \mathbb{R} \). Let \( \sigma \) be a correlated strategy. Then, for every \( z \in S \),

\[
|E_{\sigma} g_t^i dw(t)| - \int_0^T g_t^i (z, x_0, t) dw(t)| \leq 2\|\mu\|\|g\|w([0, \delta]),
\]

and, therefore, by setting \( \varepsilon = \varepsilon(\delta, \|\mu\|, w) = 2\|\mu\|\|g\|w([0, \delta]) + 2\|\mu\|^2\|u\|, \)

\[
\varepsilon + E_{\sigma} g_t^i dw(t) + u(z) - u(z) - \sum_{z' \in S} u'(z') \int_0^T \mu(z', z, x_0, t) dt \leq 2\|\mu\|^2\|u\|,
\]

and, therefore, by setting \( \varepsilon = \varepsilon(\delta, \|\mu\|, w) = 2\|\mu\|\|g\|w([0, \delta]) + 2\|\mu\|^2\|u\|, \)

\[
\varepsilon + E_{\sigma} g_t^i dw(t) + u(z) \geq \int_0^T g_t^i (z, x_0, t) dw(t) + u(z) + \sum_{z' \in S} u(z') \int_0^T \mu(z', z, x_0, t) dt
\]
\[
\geq E^Z_\sigma \left( \int_{[0,\delta]} g^i_t \, dw(t) + u(z_t) \right) - \varepsilon.
\]

**Proof.** As \( x_t = x_{t,\delta} \) on \( t < t_\delta := \delta \land s_1 \), \( |g^i_t(z_t, x_t) - g^i_t(z, x_{t,\delta})| \leq 2\|g\| \mathbb{I}(t \geq s_1) \), where \( \mathbb{I}(t \geq s_1) \) is the indicator of the event \( \{t \geq s_1\} \). Therefore,

\[
\left| E^Z_\sigma \int_{[0,\delta]} g^i_t \, dw(t) - \int_{[0,\delta]} g^i_t(z, x_{t,\delta}) \, dw(t) \right| \\
\leq E^Z_\sigma \left| \int_{[0,\delta]} |g^i_t - g^i_t(z, x_{t,\delta})| \, dw(t) \right| \leq E^Z_\sigma \int_{[0,\delta]} 2\|g\| \mathbb{I}(t \geq s_1) \, dw(t) \\
\leq \int_{[0,\delta]} 2\|g\| P^Z_\sigma(s_1 < \delta) \, dw(t) \leq 2\|g\|\|\delta\|\mu\|w([0,\delta]).
\]

The event \( \{z_{t,\delta} \neq z_\delta\} \) is a subset of the event \( \{s_2 \leq \delta\} \). Therefore, by inequality (9), \( P^Z_\sigma(z_{t,\delta} \neq z_\delta) \leq \delta^2\|\mu\|^2/2 \). For every \( z' \in S \),

\[
P(z_{t,\delta} = z') = 1_{z' = z} + \int_0^\delta \mu(z', z, x_{t,\delta}) e^{\int_0^t \mu(z', z, x_{t,\delta}) \, ds} \, dt.
\]

Therefore, as \( |e^{\int_0^t \mu(z, x_{t,\delta}) \, ds} - 1| \leq t\|\mu\| \) and thus \( \sum_{z' \in S} \int_0^\delta |\mu(z', z, x_{t,\delta})| \, dt \leq \delta^2\|\mu\|^2 \),

\[
| \sum_{z' \in S} P(z_{t,\delta} = z') u(z') - u(z) - \sum_{z' \in S} u(z') \int_0^\delta \mu(z', z, x_{t,\delta}) \, dt | \leq \|u\|\|\delta^2\|\mu\|^2,
\]

which together with the inequality \( |E^Z_\sigma u(z_\delta) - E^Z_\sigma u(z_{t,\delta})| \leq 2\|u\| P^Z_\sigma(z_{t,\delta} \neq z_{t,\delta}) \leq \|u\|\|\delta^2\|\mu\| \) completes the proof of the second part of the lemma.

The last part of the lemma is obtained by summing the inequalities of the first two parts. \( \Box \)

### 3.2. Stationary discounting games

Fix a profile \( \bar{\rho} = (\rho_i)_{i \in N} \) of discount rates. We say that the strategy profile \( \sigma \) is an \( \varepsilon \)-equilibrium, \( \varepsilon \geq 0 \), of the \( \bar{\rho} \)-discounted continuous-time stochastic game \( \Gamma_{\bar{\rho}} \), if \( \sigma \) is an admissible strategy profile such that for every player \( i \) and every strategy \( \tau^i \) of player \( i \) we have

\[
E^\infty_\sigma \int_0^\infty e^{-\rho_i t} \sum_{z \in S} \mu(z, x_t) \, dt \geq E^\infty_{\sigma^{-i}, \tau^i} \int_0^\infty e^{-\rho_i t} \sum_{z \in S} \mu(z, x_t) \, dt - \varepsilon
\]

where \( E^\infty_\sigma \) is the expectation with respect to the probability distribution \( P_\sigma \) defined by \( \sigma \) on plays, and \( \sigma^{-i}, \tau^i \) is the strategy profile whose \( i \)-th component is \( \tau^i \) and for \( j \neq i \) the \( j \)-th component is \( \sigma^j \).

The strategy profile \( \sigma \) is an equilibrium if it is a 0-equilibrium.

**Theorem 1.** Every \( \bar{\rho} \)-discounted continuous-time stochastic game \( \Gamma_{\bar{\rho}} \) (with finitely many states and actions) has a profile of stationary strategies that is an equilibrium of \( \Gamma_{\bar{\rho}} \).

**Proof.** Define \( \|g^i\| := \max_{z \in S} \max_{a \in A(z)} |g^i(z, a)| \), \( J^i(z) = [-\|g^i\|/\rho_i, \|g^i\|/\rho_i] \), and \( J = \times_{(i,z) \in N \times S} J^i(z) \). The \( (i,z) \)-th coordinate of \( v \in \mathbb{R}^{N \times S} \) is \( v^i(z) \). Recall that \( X^i(z) = \Delta(A^i(z)) \). Let \( Y^i = \times_{z \in S} X^i(z) \) and \( Y = \times_{i \in N} Y^i \).

For every \( i \in N, z \in S, a \in A(z), v \in \mathbb{R}^{N \times S}, \) and \( x \in Y \), \( G^i_x(v)(x) \), and \( f_{2,1} \) are the real-valued functions defined on \( Y \times J \) as follows.

\[
G^i_x(v)(a) = \frac{1}{\|\mu\| + \rho_i} \left( g^i(z, a) + \sum_{z' \in S} \mu(z', z, a) v^i(z') + \|\mu\| v^i(z) \right).
\]

where \( \|\mu\| = \max_{z,a} |\mu(z, a)| \).
\[ G_z^i[v](x) = \sum_{a \in A(z)} G_x^i[v](a) \prod_{j \in N} x^j(a^j), \text{ and} \]
\[ f_{z,i}(x, v) = \max_{y^i \in X^i(z)} G_z^i[v](x, y^i), \]

where \((x^{-i}(z), y^i)\) is the profile of mixed actions whose \(j\)-th coordinate for \(j \neq i\) is \(x^j(z)\) and whose \(i\)-th coordinate is \(y^i\). Let \(F_{z,i}\) be the correspondence from \(Y \times J\) to \(X^i(z)\) given by
\[ F_{z,i}(x, v) = \arg \max_{y^i \in X^i(z)} G_z^i[v](x^{-i}, y^i). \]

The cartesian product \(Y \times J\) is a product of nonempty convex compact sets and therefore it is nonempty, convex, and compact.

The function \((x, v) \mapsto G_z^i[v](x)\) is a polynomial in the coordinates of \((x, v)\) and therefore continuous, and therefore \(f_{z,i}(x, v)\) is also a continuous function of \((x, v)\).

If \(|v| \leq \|g^i\|/\rho_i\) then \(|G_z^i[v](a)| \leq \frac{1}{\|\mu\|+\rho_i} (\|g^i\|+\|\mu\|\|g^i\|/\rho_i) \leq \|g^i\|/\rho_i\). Therefore \(|G_z^i[v](x)| \leq \|g^i\|/\rho_i\) and therefore also \(|f_{z,i}(x, v)| \leq \|g^i\|/\rho_i\).

We deduce that \(f_{z,i}\) is a continuous function from \(X \times J\) to \(J^i(z)\).

The correspondence \(F_{z,i}\) from \(Y \times J\) to \(X^i(z)\) is nonempty, convex-valued, and upper semicontinuous. Therefore the correspondence \(F\) defined on \(Y \times J\) by \(F(x, v) = \times_{z,i}(F_{z,i}(x, v) \times \{f_{z,i}(x, v)\})\) is a nonempty, convex-valued, upper semicontinuous correspondence from the nonempty convex compact set \(Y \times J\) to itself, and therefore has a fixed point.

Let \((x, V)\) be a fixed point of \(F\). We claim that the profile of stationary strategies \(\sigma^i\) with \(\sigma^i(h_t) = x^i(z_t)\) is an equilibrium of the \(\rho\)-discounted game with equilibrium payoff \(V\).

Fix a player \(i \in N\), a strategy \(\tau^i\) of player \(i\), and an initial state \(z = z_0\). First we prove that
\[ E_{\sigma^{-i}, \tau^i}^Z \int_0^\infty e^{-\rho \delta t} g(z_t, x_t) dt \leq V^i(z). \tag{10} \]

For every correlated strategy, or a strategy profile that defines a unique distribution on plays, \(\tau\), state \(z \in S\), and \(s \geq 0\), define \(f_{s}^z(s) = E_{\sigma^{-i}, \tau^i}^Z \int_0^\infty e^{-\rho \delta t} g(z_t, x_t) dt + e^{-\rho \delta} V^i(z_s)\).

Set \(f(s) = f^2(s) = f_{\sigma^{-i}, \tau^i}(s)\). Note that \(f(0) = V^i(z)\) and that \(f(s) \rightarrow s \rightarrow \infty E_{\sigma^{-i}, \tau^i}^Z \int_0^\infty e^{-\rho \delta t} g(z_t, x_t) dt\). Therefore, in order to prove (10), it suffices to prove that \(f\) is Lipschitz and that the upper-right derivative of \(f\) is nonpositive a.e.

Let \(y_t := y_{0,i}(x, \sigma^{-i}, \tau^i)\), namely, \(y_t = (\sigma^{-i}, \tau^i)(h_t, t)\), where \(h\) is a play that is (1) compatible with \(\sigma\) and (2) \(z_t = z\) for every \(0 \leq s < \delta\). By Lemma 2,
\[
\begin{align*}
    f(\delta) - f(0) &= E_{\sigma^{-i}, \tau^i}^Z \left( \int_0^\delta e^{-\rho \delta t} g(z_t, x_t) dt + e^{-\rho \delta} V^i(z_s) \right) \\
    &\leq \int_0^\delta e^{-\rho \delta t} g(z_0, y_t) dt + e^{-\rho \delta} \left( V^i(z_0) + \sum_{z' \in S} V^i(z') \int_0^\delta \mu(z', z_0, y_t) dt \right) + O(\delta^2) \\
    &\leq \int_0^\delta \left( g(z_0, y_t) \sum_{z' \in S} V^i(z') \mu(z', z_0, y_t) - \rho_i V^i(z_0) \right) dt + O(\delta^2).
\end{align*}
\]

Therefore, the function \(s \mapsto f(s)\) is Lipschitz \((|f(s + \delta) - f(s)| \leq \delta e^{-\rho \delta} (2\|\mu\|\|V^i\| + \|g\| + O(\delta))\)). Therefore, in order to prove (10), it suffices to prove that the upper-right derivative of \(f\) is nonpositive a.e.

The choice of \(\sigma\) implies that \(g(z_0, y_t) + \sum_{z' \in S} \mu(z', z_0, y_t) V^i(z') - \rho_i V^i(z_0) \leq 0\) and therefore the upper-right derivative of \(f\) at 0, \(\lim_{s \rightarrow 0^+} f(s)/s\), is nonpositive.

Similarly, for a.e. \(s > 0\), the upper-right derivative of \(f\) at \(s\) is nonpositive. We conclude that \(f\) is nonincreasing.

Now setting \(\tau^i = \sigma^i\), we can change \(\leq \) in the above-derived inequalities to \(\geq\) (and replace the terms \(+ O(\delta^2)\) by \(\pm O(\delta^2)\)) in order to deduce that \(f^2_s(x)\) is nondecreasing and thus
\[ E_{\sigma}^Z \int_0^\infty e^{-\rho \delta t} g(z_t, x_t) dt \geq V^i(z), \]
which together with inequality (10) completes the proof of the theorem. □
Covariance properties. In the case where \( \rho_i = \rho \) for every \( i \), an alternative interpretation of the \( \rho \)-discounted game is as a game with an uncertain duration \( d \) in which each player maximizes his expected total payoff. Under this interpretation, \( \rho \) is a parameter of the distribution of the random duration \( d \), \( P(d \geq t) = e^{-\rho t} \). Multiplying \( \rho \) by \( \alpha \) means that the game terminates faster at a rate of \( \alpha \). This can be compensated by multiplying all transitions and payoffs by \( \alpha \) as well. Therefore, any stationary \( \rho \)-discounted equilibrium in the game \( (N, S, A, \mu, g) \) is also a stationary \( \alpha \rho \)-discounted equilibrium in the game \( (N, S, A, \alpha \mu, \alpha g) \). In addition, if \( \bar{p} \) is the transition matrix such that \( \bar{p}(z') | z, a) = \frac{1}{1-\rho} \mu(z, z, a) \) for all \( z' \neq z \) and \( \bar{p}(z | z, a) = 1 + \frac{1}{1-\rho} \bar{\mu}(z, z, a) \), then equations defining a fixed point \( (x, V) \) of the correspondence \( F \) (used in the proof of the theorem) are equivalent\(^9\) to the equations defining a stationary equilibrium of the discrete-time stochastic game with a stage payoff function \( g \) and transition probabilities \( \bar{p} \).

Formally, and more generally, fix \( \alpha > 0 \). Note that the auxiliary function \( G_\alpha^1[V] \) of the \( \bar{p} \)-discounted game \( \Gamma = (N, S, A, \mu, g) \) and the \( \alpha \bar{p} \)-discounted game \( \Gamma = (N, S, A, \alpha \mu, \alpha g) \) are the same. Therefore, a point \( (x, V) \) of the \( \alpha \bar{p} \)-discounted game \( \Gamma = (N, S, A, \alpha \mu, \alpha g) \) is a stationary equilibrium (strategies and payoffs) of the continuous-time \( \bar{p} \)-discounted game \( \Gamma = (N, S, A, \mu, g) \) if and only if \( (x, V) \) is a stationary equilibrium of the continuous-time \( \alpha \bar{p} \)-discounted game \( \Gamma = (N, S, A, \alpha \mu, \alpha g) \).

The discounting minmax. The next theorem is used in the proof of Theorem 3, which is used (in the punishing phases) in the proof of the main result. It asserts that (1) the minmax \( \bar{V}_\rho^i \) of player \( i \) in the unnormalized \( \rho \)-discounted game exists, (2) the function \( \rho \mapsto \bar{V}_\rho^i \) is semialgebraic,\(^10\) and (3) there are minimaxing strategies that are stationary.

**Theorem 2.** Fix a \( \rho \)-discounted continuous-time stochastic game (with finitely many states and actions) \( \Gamma_\rho \) and a player \( i \in N \).

a) The set of equations in the real variables \( v(z), z \in S \),

\[
\rho v(z) = \min_{x^{-i}} \max_{x^i} \left( g^i(z, x^{-i} \otimes x^i) + \sum_{z' \in S} \mu(z', z, x^{-i} \otimes x^i) v(z') \right),
\]

(11)

where the minimum is over all \( x^{-i} = (x^j)_{j \neq i} \in \times_{j \neq i} \Delta(A^j(z)) \) and the maximum is over all \( x^i \in \Delta(A^i(z)) \), has a unique solution \( \bar{V}_\rho^i \).

b) The function \( \rho \mapsto \bar{V}_\rho^i \) is semialgebraic and \( \| \bar{V}_\rho^i \| \leq \| g^i \| / \rho \).

c) There are stationary strategies \( \sigma^j, j \neq i \), such that for every strategy \( \sigma^i \) of player \( i \) and every initial state \( z \) we have

\[
E_{\sigma^j, \sigma^j}^Z \left( \int_0^t e^{-\rho^i \bar{V}_\rho^i(z_i)} dt + e^{-\rho \bar{V}_\rho^i(z_j)} \right) \leq \bar{V}_\rho^i(z) \quad \forall s \geq 0, \text{ and}
\]

\[
E_{\sigma^i, \sigma^j}^Z \left( \int_0^\infty e^{-\rho^i \bar{V}_\rho^i(z_i)} dt \leq \bar{V}_\rho^i(z) \right).
\]

(12)

**Proof.** Let \( Q \) be the map from \( \mathbb{R}^5 \) to itself that is given by \( Q v(z) = \min_{x^{-i}} \max_{x^i} G_\rho^i[z^{-i}, x^i] \), where

\[
G_\rho^i[z^{-i}, x^i] = \frac{1}{\| \mu \| + \rho} \left( g^i(z, x^{-i} \otimes x^i) + \sum_{z' \in S} \mu(z', z, x^{-i} \otimes x^i) v(z') + \| \mu \| v(z) \right).
\]

\( Q \) is a strict contraction and therefore has a unique fixed point \( V \).

A vector \( V = (\bar{V}_\rho^i(z))_{z \in S} \) is a solution of \( Q v = v \) if and only if it is a solution of equality (11). Therefore, equality (11) has a unique solution. This completes the proof of part a).

Observe that any solution \( (\bar{V}_\rho(z))_{z \in S} \) of equality (11) obeys \( \| \bar{V}_\rho(z) \| \leq \| g^i \| / \rho \), and the semialgebraicity of equality (11) guarantees that the function that maps \( \rho \) to the unique solution \( (\bar{V}_\rho(z))_{z \in S} \) is semialgebraic. This completes the proof of part b).

Let \( x^{-i}(h_i) = (x^j(h_i))_{j \neq i} \) be the minimizer of the right-hand side of (11) for \( z = z_i \).

For every \( j \neq i \) we define the strategy \( \sigma_j^i \) by \( \sigma_j^i(h) = x^i(h_i) \). As in the proof of Theorem 1, it follows that for every strategy \( \sigma^i \) we have

---

\(^9\) This equivalence can be used to provide an alternative proof of the existence of a fixed point of \( F \).

\(^10\) See, e.g., Neyman (2003a) for the definition of semialgebraic functions, and their applications to stochastic games.
\[
E^2_{\sigma^i,\sigma^j} \int_0^\infty e^{-\rho t} g^i(z_t, x_t) dt \leq V(z),
\]

which completes the proof of the theorem. □

Theorem 2 implies that the function \( \rho \mapsto v^i_{\rho}(z) := \rho \bar{V}^i_{\rho} \) is a bounded semialgebraic function, and therefore it converges to a limit \( v^i(z) \) as \( \rho \to 0^+ \).

3.3. The undiscounted minimax

Fix a continuous-time stochastic game \( \Gamma = (N, S, A, \mu, g) \) with finitely many states and actions.

A correlated strategy \( \sigma \) defines for every state \( z \in S \) a unique probability distribution \( P^z_\sigma \) on the plays of \( \Gamma \). The expectation with respect to the probability distribution \( P^z_\sigma \) is denoted by \( E^z_\sigma \).

For every player \( i \in N \), we set

\[
\gamma^i_s(z, \sigma) := E^z_\sigma \frac{1}{S} \int_0^S g^i_t dt, \text{ where } g^i_t := g^i(z_t, x_t),
\]

\[
\bar{\gamma}^i_s(z, \sigma) := E^z_\sigma \limsup_{s \to \infty} \frac{1}{S} \int_0^S g^i_t dt, \text{ and } \underline{\gamma}^i_s(z, \sigma) := E^z_\sigma \liminf_{s \to \infty} \frac{1}{S} \int_0^S g^i_t dt.
\]

Recall that the profile of strategies that is obtained by a unilateral deviation from a profile of admissible strategies (e.g., from a profile of discretized strategies) is equivalent, in terms of the probability distribution that it defines on the plays, to a correlated strategy.

In the previous section, we showed that there are stationary minimaxing strategies in the \( \rho \)-discounted games.

There are continuous-time stochastic games (as in the discrete-time case) for which Markov strategies can serve neither as approximate minimaxing strategies in the limiting-average game, nor as approximate minimaxing strategies in all sufficiently long finite-horizon games.

The next result shows that such minimaxing strategies exist in the class of discretized strategies. Moreover, the discretization can be done with the sequence of nonnegative integers. An \( \mathbb{N} \)-discretized strategy is a \( (t_k)_{k \geq 0} \)-discretized strategy where \( t_k = k \).

Theorem 3. For every \( \epsilon > 0 \) and every player \( i \) there are \( \mathbb{N} \)-discretized strategies \( \sigma^{-i}, \sigma^i \), \( j \neq i \), and a time \( s \) such that for every strategy \( \sigma^i \) of player \( i \) and every \( s > s \) we have

\[
\gamma^i_s(z, \sigma^{-i}, \sigma^i) \leq v^i(z) + \epsilon \quad \text{and} \quad \bar{\gamma}^i_s(z, \sigma^{-i}, \sigma^i) \leq v^i(z) + \epsilon.
\]

Proof. For every nonnegative integer \( k \geq 0 \), let \( r_k := \int_0^{k+1} g^i_t dt \).

Fix a player \( i \in N \). For every sequence \( \rho_k, k \geq 0 \), where \( \rho_k > 0 \) is a function of \( h_{k+1} \) (e.g., a function of \( \rho_k, z_{k+1} \), and \( t_k \)), there is by Theorem 2 an \( \mathbb{N} \setminus \{0\} \) profile of strategies \( \sigma^{-i} = (\sigma^j)_{j \neq i} \), so that for every strategy \( \sigma^i \) of player \( i \) we have

\[
E^{h_k}_{\sigma^i} \left( \int_k^{k+1} \rho_k e^{-\rho_k(t-k)} g^i_t dt + e^{-\rho_k} v^i_{\rho_k}(z_{k+1}) \right) \leq v^i_{\rho_k}(z_k),
\]

where \( \sigma \) is the strategy profile \( (\sigma^{-i}, \sigma^i) \), and \( E^{h_k}_{\sigma^i} f \) is the conditional expectation, given \( h_k \), of the random variable \( f \) w.r.t. the probability on plays defined by the strategy profile \( \sigma \).

As \( |g^i_t| = |g^i_t(z_t, y_t)| \leq \|g^i\| \) and \( \|v^i_{\rho_k}\| := \max_{z \in S} |v^i_{\rho_k}(z)| \leq \|g\| \), the inequality \( \int_0^1 |1 - e^{-\rho t}| |f| dt + |1 - \rho - e^{-\rho}| = 2|1 - \rho - e^{-\rho}| \leq O(\rho^2) \) implies that

\[
E^{h_k}_{\sigma^i} (\rho_k r_k + (1 - \rho_k) v^i_{\rho_k}(z_{k+1})) \leq E^{h_k}_{\sigma^i} \left( \int_k^{k+1} \rho_k e^{-\rho_k(t-k)} g^i_t dt + e^{-\rho_k} v^i_{\rho_k}(z_{k+1}) \right) + O(\rho_k^2).
\]

Therefore, for sufficiently small \( \theta > 0 \), the inequality \( \rho_k \leq \theta \) implies that

\[
E^{h_k}_{\sigma^i} (\rho_k r_k + (1 - \rho_k) v^i_{\rho_k}(z_{k+1})) \leq v^i_{\rho_k}(z_k) + \epsilon(\theta) \rho_k,
\]

where \( \epsilon(\theta) \to 0 \) as \( \theta \to 0^+ \).
By the proof in Mertens and Neyman (1981) (or using the statement of Neyman, 2003b, Lemma 1) and the inequality
\[ \frac{1}{T} \int_0^T g^t(z, x_t) dt \leq \frac{1}{|T|} \int_0^{|T|} g^t(z, x_t) dt + O(\varepsilon), \]
we deduce that there is a sequence of history-dependent (hence, random) discount rates \( \rho_k \) (where \( \rho_0 \) is a constant and for a positive integer \( k \) the discount rate \( \rho_k \) is a function of \( z_k, \rho_{k-1}, \) and \( r_{k-1} \)) such that the corresponding strategy profile \( \sigma^{-i} \) satisfies, for all \( s \) sufficiently large and all strategies \( \sigma^i \) of player \( i \),
\[ \gamma^{\sigma^i}_s(z, \sigma^{-i}, \sigma^i) \leq \nu^i(z) + O(\varepsilon) \quad \text{and} \quad \tilde{\gamma}^{\sigma^i}_s(z, \sigma^{-i}, \sigma^i) \leq \nu^i(z) + O(\varepsilon). \]

The choice of \( N \) in the construction of the \( N \)-discretized minimaxing strategies is for convenience. In fact, for any increasing sequence \( (t_k) \) of reals with \( t_0 = 0 \) and, e.g., bounded difference \( t_{k+1} - t_k \), there are undiscounted minimaxing strategies that are \( (t_k) \)-discretized.

3.4. Uniform \( l \)-average equilibria

One of the central open problems in stochastic games is the existence of a uniform (or a limiting-average) equilibrium payoff in (discrete-time) stochastic games with finitely many states and actions.

The present paper proves the existence of uniform and limiting-average equilibrium payoffs in all continuous-time stochastic games with finitely many states and actions.

**Theorem 4.** Let \( \Gamma = \langle N, S, A, g, \mu \rangle \) be a continuous-time stochastic game with finitely many states and actions. There exists a vector payoff \( u \in \mathbb{R}^{S \times N} \) such that for every \( \varepsilon > 0 \) there are (discretized) strategies \( \sigma^i, i \in N, \) and a positive number \( s_0 > 0 \), such that for every \( s > s_0 \), every player \( i \), every state \( z \in S \), and every strategy \( \tau^i \) of player \( i \), we have
\[
-\varepsilon + \gamma^{\sigma^i}_s(z, \sigma^{-i}, \tau^i) \leq u^i(z) \leq \nu^i(z, \sigma) + \varepsilon, \quad \text{and} \quad
-\varepsilon + \tilde{\gamma}^{\sigma^i}_s(z, \sigma^{-i}, \tau^i) \leq u^i(z) \leq \nu^i(z, \sigma) + \varepsilon.
\]

Such a vector payoff \( u \) is called a uniform \( l \)-average equilibrium payoff, and the corresponding strategy \( \sigma \) is called a uniform \( l \)-equilibrium strategy.

A vector payoff \( u \) for which \( \forall \varepsilon > 0 \exists \sigma, s_0 \) s.t. \( \forall s > s_0, i, \tau^i \) the system of inequalities (13) holds is called a uniform equilibrium payoff, and a vector payoff \( u \) for which \( \forall \varepsilon > 0 \exists \sigma \) s.t. \( \forall i, \tau^i \) the system of inequalities (14) holds is called a limiting-average equilibrium payoff.

The corresponding strategies are called a uniform \( \varepsilon \)-equilibrium strategy and a limiting-average \( \varepsilon \)-equilibrium strategy.

The proof of **Theorem 4** is given in Section 4. The proof follows a similar outline to that used in Solan and Vieille (2002) for the proof of the existence of uniform extensive-form correlated equilibria in discrete-time stochastic games.

3.5. Robust equilibria

Given a correlated strategy profile \( \sigma \) (which, as said earlier, defines for every state \( z \in S \) a unique probability distribution \( P^z_\sigma \)) on the plays of the continuous-time stochastic game \( \Gamma = \langle N, S, A, g, \mu \rangle \), a profile of discounting measures \( \tilde{w} \in W^N \), and a player \( i \in N \), we set
\[
\gamma^{\sigma^i}_\tilde{w}(z, \sigma) := E^\tilde{w}_\sigma \int_{[0, \infty)} g^t_i dw^i(t) + w^i(\infty) \gamma^i(z, \sigma) = E^\tilde{w}_\sigma \int_{[0, \infty)} g^t_i dw^i(t),
\]
where \( g^t_i = g^t_i \) if \( t < \infty \) and \( g^t_i = \lim \inf_{s \to \infty} \frac{1}{s} \int_0^s g^t_i dt \), and
\[
\tilde{\gamma}^{\sigma^i}_\tilde{w}(z, \sigma) := E^\tilde{w}_\sigma \int_{[0, \infty)} g^t_i dw^i(t) + w^i(\infty) \tilde{\gamma}^i(z, \sigma) = E^\tilde{w}_\sigma \int_{[0, \infty)} g^t_i dw^i(t),
\]
where \( g^t_i = g^t_i \) if \( t < \infty \) and \( g^t_i = \lim \sup_{s \to \infty} \frac{1}{s} \int_0^s g^t_i dt \).

**Theorem 5.** Let \( \Gamma = \langle N, S, A, g, \mu \rangle \) be a continuous-time stochastic game with finitely many states and actions. For every \( \varepsilon > 0 \), there is a finite set \( A \) and an open cover (\( U_A \)) of \( W^N \), such that for every \( \alpha \in A \) there is a vector payoff \( u_\alpha \in \mathbb{R}^{S \times N} \) and a profile of (discretized) strategies \( \sigma_\alpha \), such that for every \( i \in U_\alpha \), \( \sigma_\alpha \) is an \( \varepsilon \)-equilibrium of \( \Gamma_\alpha \) with a payoff \( \varepsilon \)-close to \( u_\alpha \); alternatively, such that for every player \( i \), every state \( z \in S \), and every strategy \( \tau^i \) of player \( i \) we have
\[
-\varepsilon + \tilde{\gamma}^{\sigma_\alpha^i}_\tilde{w}(z, \sigma^{-i}, \tau^i) \leq u^i(z) \leq \nu^{\sigma_\alpha^i}(z, \sigma^{-i}) + \varepsilon.
\]
The concept of robust equilibria depends on the topology defined on \((W, \text{ hence on }) W_d\). Note that the coarser the topology is, the stronger the corresponding robustness result is.

A basis for the \(w^\ast\)-neighborhoods of \(W^N\) are sets of the form \(\{\hat{w} = (w^i)_{i \in \mathbb{N}} \in W^N : |\int_{[0,\infty)} f^j_i(t) \, dw^k_i(t) - c^j_i| < \delta \text{ } \forall j \in \mathbb{N}, \text{ where } j \text{ is a finite set, } f^j_i \in \mathbb{C}(0, \infty[) \text{ with } 0 \leq f^j_i \leq 1, \text{ and } c^j_i, d^j_i \in [0, 1]\).

Therefore, the robustness result is equivalent to the following statement.

For every \(\varepsilon > 0\), there is \(\delta > 0\) and a finite set \(\mathcal{A}\), where \(\alpha \in \mathcal{A}\) is a list \((f^i_{a\alpha}, c^i_{a\alpha}, v^i_{a\alpha}, \sigma^i_{a\alpha})\) with \(f^i_{a\alpha} = (f^i_{a\alpha})_{i \in \mathbb{N}}\) a vector of \([0, 1]\)-valued continuous functions on \([0, \infty[)\), \(\{c^i_{a\alpha} : \alpha \in \mathcal{A}\}\) being a finite \(\delta\) net of \([0, 1]^N\), \(v^i_{a\alpha} \in \mathbb{R}^{3 \times N}\), and \(\sigma^i_{a\alpha}\) an admissible strategy profile, such that for every \(\hat{w} \in W^N_d\) with \(|\int_{[0,\infty)} f^j_i(t) \, dw^k_i(t) - c^j_i| < \delta, \text{ } \sigma^i_{a\alpha}\) is an \(\varepsilon\)-equilibrium of \(\Gamma_{\hat{w}}\) with a payoff \(\varepsilon\)-close to \(u_{a\alpha}\).

A crucial step in the proof of Theorem 5 is Theorem 4, which is of independent interest.

4. Proof of Theorem 4

Theorem 4 is straightforward if \(\|\mu\| = 0\). Indeed, if \(\|\mu\| = 0\), there are no state changes in a play, i.e., conditional on the initial state \(z\) we have a continuous-time supergame. Therefore, if \(\sigma\) is a stationary strategy profile with \(\alpha\) an equilibrium of the single-stage game with payoff function \(a \rightarrow g(z, a)\) and \(u(z) = g(z, \sigma(z))\), then inequalities (13) and (14) hold for every \(\varepsilon \geq 0\). Therefore, we assume that \(\|\mu\| > 0\).

Without loss of generality we can assume that \(0 \leq g^i \leq 1\) and \(\|\mu\| = 1\). Indeed, the continuous-time stochastic game \(\Gamma = (N, S, A, \mu, \bar{g})\) has a uniform-l-average equilibrium payoff if and only if \(\tilde{\Gamma} = (N, S, A, \tilde{\mu}, \bar{g})\) does, where \(\tilde{\mu} = \alpha \mu\), and \(\tilde{g}^i = \beta\bar{g} + \alpha g^i\) for some (and therefore for all) \(\alpha, \alpha_1, \ldots, \alpha_n > 0\) and \(\beta_1, \ldots, \beta_n \in \mathbb{R}\). Therefore, we assume that \(0 \leq g^i(z, a) \leq 1\) and \(\|\mu\| = 1\).

In order to prove Theorem 4, it suffices to prove it with the \(\varepsilon\)-dependent vector payoff \(u\) being replaced by an \(\varepsilon\)-dependent \(u_{\varepsilon}\). Namely, it suffices to prove that \(\forall \varepsilon > 0\), \(\exists u_{\varepsilon}, s_{\varepsilon}, \sigma_{\varepsilon}\) s.t. \(\forall \tau_{\varepsilon} \leq s < \infty\), (13) and (14) hold when \(u = u_{\varepsilon}\) and \(\sigma\) are replaced there by \(u_{\varepsilon}, s_{\varepsilon}\), and \(\sigma_{\varepsilon}\), respectively. Indeed, let \(u\) be a limit point of \(u_{\varepsilon}\); then \(\exists \tau_{\varepsilon} \rightarrow k \rightarrow \infty\) 0+ s.t. \(u_{\varepsilon} \rightarrow k \rightarrow \infty\) \(u\). For \(\varepsilon > 0\), (13) and (14) hold with \(s_0 = s_{\varepsilon}\) and \(\sigma = \sigma_{\varepsilon}\) and \(k\) sufficiently large so that \(\varepsilon_k \varepsilon / 2\) and \(\|u - u_{\varepsilon}\| < \varepsilon / 2\).

4.1. An informal outline of the proof

The proof constructs a profile of discretized pure-action strategies \(\hat{\tau}\) that depends on \(\varepsilon > 0\) and satisfies several properties. The discretization is with respect to a refinement of a sequence of real numbers \(T = (t_j)_{j \geq 0}\), where \(t_0 = 0\) and \(t_j < t_{j+1} \rightarrow j \rightarrow \infty\).

The approximate equilibrium strategy profile \(\sigma\) consists of following \(\hat{\tau}\) as long as no player deviates from playing according to \(\tau\). If no player deviates before time \(t_{j-1}\) and a single player deviates in the time interval \([t_{j-1}, t_j]\), then all the other players switch at time \(t_j\) to punishing the deviating player in the uniform-l-average game.

The constructed strategy profile \(\hat{\tau}\) satisfies the following property. For any nonnegative integer \(j\) and \(s \geq t_j\), the following inequality holds.

\[ E^\hat{\tau}_s \left( \frac{1}{s-t_j} \int_{t_j}^s g^i \, dt \right) \geq E^\hat{\tau}_s \left( z(t_j) \right) - O(\varepsilon) - O \left( \frac{1}{s-t_j} \right). \] (16)

A deviation from \(\hat{\tau}\) in the time interval \([t_{j-1}, t_j]\) cannot change the expectation of \(\psi^i(z(t_j))\) or the payoff of a player in this time interval by more than \(O(t_j - t_{j-1})\).

Therefore, if \(\sup_j (t_{j+1} - t_j) \leq \varepsilon\), then for any strategy \(\tau\) of player \(i\),

\[ \gamma_s^i(z, \tau^{i-1}, \tau^i) = E^\tau_{\sigma^{i-1}, \tau^i} \left( \frac{1}{s} \int_0^s g^i \, dt \right) \geq \gamma_s^i(z, \hat{\tau}) + O(\varepsilon) + O \left( \frac{1}{s} \right). \] (17)

Therefore, if \(\sup_j (t_{j+1} - t_j) \) is sufficiently small, \(\sigma\) is an approximate equilibrium of the \(s\)-horizon game with a payoff within \(O(\varepsilon)\) of \(\gamma_s^i(z, \hat{\tau})\) for all sufficiently large \(s\).

A priori, the limit of \(\gamma_s^i(z, \hat{\tau})\), as \(s \rightarrow \infty\), need not exist. Moreover, the difference between the limit superior and the limit inferior of \(\gamma_s^i(z, \hat{\tau})\) need not be small. Therefore, condition (17) by itself does not guarantee the existence of a uniform equilibrium payoff.

In order to prove the existence of a uniform equilibrium payoff, the constructed strategy \(\hat{\tau}\) will be such that in addition to obeying (16) it satisfies

\[ \gamma_s^i(z, \hat{\tau}) = E^\hat{\tau}_s \left( \frac{1}{s} \int_0^s g^i \, dt \right) \text{ converges as } s \rightarrow \infty. \] (18)
The existence, for every $\varepsilon > 0$, of a profile of $T = (t_j)_{j \geq 0}$-discretized pure-action strategies $\tilde{\tau}$ (with $\sup_j (t_{j+1} - t_j)$ sufficiently small) that satisfies (16) and (18) guarantees the existence of a uniform equilibrium payoff.

In order to prove the existence of a uniform-1-average equilibrium payoff, the constructed strategy $\tilde{\tau}$ will be such that, in addition to satisfying (16) and (18), it satisfies

$$\frac{1}{s} \int_0^s g_t \, dt \text{ converges as } s \to \infty \text{ with } P_\tau^s - \text{probability } 1,$$

and

$$\lim_{s \to \infty} \gamma_t^s(z, \tilde{\tau}) = E_z^s \lim_{t \to \infty} \frac{1}{s} \int_0^s g_t \, dt.$$

(20)

The pure-action strategy $\tilde{\tau}$ is derived by a “purification” of a $T$-discretized correlated strategy $\tau$.

If $\tilde{\tau}$ satisfies conditions (16)–(20), then, by Lemma 2, if in each time interval $[t_{j-1}, t_j)$, the empirical distribution of the actions of the pure Markov strategy $\tilde{\tau}$ is sufficiently close to the average of the mixed action profile of $\tilde{\tau}$, then $\tilde{\tau}$ will satisfy these conditions with a possible error of $O(\sum_j (t_j - t_{j-1})^2)$.

Therefore, the existence of a $T$-discretized correlated strategy $\tilde{\tau}$ that satisfies conditions (16)–(20) will imply the existence of a uniform-1-average equilibrium.

Let us mention that the existence of a continuous-time strategy that obeys conditions (16)–(20) follows easily from the existence of a stationary correlated strategy in the discrete-time game that obeys the discrete-time analogous conditions (16*)–(20*).

Moreover, if $\tilde{\tau}$ is a discrete-time stationary correlated strategy, then conditions (18*), (19*), and (20*) hold, and condition (16*) holds if and only if condition (17*) holds.

However, the existence of a continuous-time correlated Markov strategy that obeys conditions (16)–(20) does not follow easily from the existence of a correlated Markov strategy (in the discrete-time game) that obeys conditions (16*)–(20*).

Therefore, we cannot use the result of Solan and Vieille (2002) directly. However, we follow similar ideas to those used in Solan and Vieille (2002).

The strategy $\tilde{\tau}$ is constructed in a sequence of steps.

First, one defines a strategy profile $\tau$ that is discretized w.r.t. the sequence $T_\ell = (j/\ell)_{j \geq 0}$, where $\ell$ is a sufficiently large integer, and a sufficiently large positive integer $N_0$, such that

$$E_z^\ell v^\ell(zT) \geq v^\ell(z) - \varepsilon_1/2 \text{ for every finite } (H_\ell)_{\ell \geq 0}-\text{stopping time } T$$

and

$$\gamma_{N_0}^\ell(z, \tau) \geq v^\ell(z) - \varepsilon_1/2,$$

(21)

where $\varepsilon_1$ is sufficiently small. See Section 4.4.

The second step defines a correlated strategy $\tilde{\tau}$ that is discretized w.r.t. the sequence $T_\ell = (j/\ell)_{j \geq 0}$ and (1) coincides with $\tau$ if the initial state is in some subset $S_2$ of states, where the set $S_2$ depends on $\tau$, (2) coincides with a profile of stationary strategies $\gamma$ if the initial state is in some subset $S_3$ of states, where $\gamma$ and $S_3$ depend on the stochastic game (but not on $\tau$), and (3) coincides with a discretized correlated strategy whose mixed action is close to that of $\tau$ but avoids certain possible transitions if the initial state is in the complement $S_1$ of $S_2$ and $S_3$. See Section 4.7.

The third step “approximates” the correlated strategy $\tilde{\tau}$ by a pure-action strategy $\tilde{\tau}$. The “approximation” is so that in any time interval of the form $[t_{j-1}, t_j)$, where $k$ is a nonnegative integer, the empirical distribution of the pure actions of $\tilde{\tau}$ is within $O(1/\ell)$ of the average of the correlated mixed actions of $\tilde{\tau}$. The “approximation” enables us to use Lemma 2 in order to derive properties of $\tilde{\tau}$ from the corresponding properties of $\tilde{\tau}$. See Section 4.8.

Finally, one defines the pure-action strategy $\tilde{\tau}$ that is played in blocks of play-dependent random duration that are $\leq N_0$. At the start of each block, $\tilde{\tau}$ restarts with the strategy $\tilde{\tau}$. The periodicity of $\tilde{\tau}$ guarantees that $\tilde{\tau}$ obeys properties (18)–(20). See Section 4.8.

The approximate equilibrium strategy $\sigma$ follows $\tilde{\tau}$ until the first time $t = j/\ell$, with $j$ a positive integer, where a deviation by a single player was observed, and thereafter all other players punish the deviating player in the uniform-1-average game. See Section 4.10.

4.2. A few auxiliary functions of $\Gamma$

In this subsection we define $v_\rho$, $v$, $C_z$, $w$, and $d > 0$, as a function of $\Gamma$. In short, $v_\rho^i(z)$ is a payoff level where the other players can force the expected payoff of player $i$ in the normalized $\rho$-discounted game that starts at state $z$ to be no more than this level; $v(z)$ is the limit of $v_\rho^i(z)$ as $\rho \to 0$; $C_z$ is the subset of states $z'$ with $v(z') = v(z)$ (namely, $v(z') = v(z)$
for all $i$; $w \in \mathbb{R}^5$ is a positive linear combination of the vectors $v^j_i \in \mathbb{R}^5$, with $0 \leq w(z) \leq 1$ and such that $v(z) \neq v(z')$ implies that $w(z) \neq w(z')$ (and therefore $w(z) = w(z')$ whenever $v(z) = v(z')$); and $d > 0$ is a positive constant such that $w(z) \neq w(z')$ implies that $|w(z) - w(z')| > d$.

**Definition of $\nu^\rho$ and $v$.** Let $V^\rho = (V^\rho_i(z))_{z \in S}$ be the (unique) solution $V \in \mathbb{R}^5$ of the system of equations

$$\rho V(z) = \max_{x^i} \left( g^i(z, x^{-i} \otimes x^i) + \sum_{z' \in S} \mu(z, z', x^{-i} \otimes x^i) V(z') \right), \quad z \in S,$$

where the minimum ranges over all $x^i = (x^i)_{i \neq 0} \in X^{-i}(z) := x_{j \neq i} \Delta(A^j(z))$, and the maximum ranges over all $x^i \in X^i(z)$. Set $V^\rho_i = \rho V^\rho_i$. The function $\rho \mapsto V^\rho_i(z)$ is a bounded semialgebraic function, and therefore it converges to a limit $v(z)$ as $\rho \rightarrow 0^+$. The assumption $0 \leq \rho^i \leq 1$ implies that $0 \leq v^\rho_i(z) \leq 1$, and therefore $0 \leq v^\rho_i(z) \leq 1$, for every $\rho > 0, i \in N$, and $z \in S$.

**Definition of $C_d$.** For every state $z \in S$ let $C_d = \{z' \in S \text{ s.t. } v(z') = v(z)\}$.

**Definition of $w$ and $d$.** For $i \in N$, with $\max_{z \in S} v^i(z) > \min_{z \in S} v^i(z)$, we set $d_i = \min_{z \in S} v^i(z) - v^i(z') : z, z' \in S, \text{ and } v^i(z) > v^i(z')$. For $i \in N$ with $v^i(z) = v^i(z')$ for every $z, z' \in S$, we set $d_i = 0$. By renaming the players, we can assume without loss of generality that $N = \{1, 2, \ldots, n\}$ and $d_i \geq d_{i+1}$ for $1 \leq i < n$. We will see that if $d_1 = 0$ (and then $v(z) = v(z')$ for every $z, z' \in S$) the theorem follows easily. So assume that $d_1 > 0$ and let $i_0$ be the maximal positive integer that is less than or equal to $n$ and with $d_{i_0} > 0$.

For $z \in S$, we denote by $v(z)$ the vector $(v^i(z))_{i \in N} \in \mathbb{R}^n$. We define a $[0, 1]$-valued function $w$ on $S$ and a positive number $d > 0$ such that for every two states $z, z' \in S$ and player $i_1 \in N$, we have

$$v(z) \neq v(z') \implies |w(z) - w(z')| \geq d, \quad \text{and}$$

$$v^i(z) = v^i(z') \quad \forall i < i_1 \text{ and } v^i_1(z) > v^i_1(z') \implies w(z) > w(z').$$

For example, the function $w(z) = \sum_{i=1}^n 2^{-i} (\prod_{j \neq i} d_j) v^i(z)$ is $[0, 1]$-valued and satisfies (24), and if $d = \min_{z \in S} w(z) - w(z') : z, z' \in S, \text{ and } w(z) > w(z')$, then $d \geq 2^{-n} \prod_{j \in S} d_j > 0$, and the function $w$ and the positive constant $d$ satisfy (23) and (24).

### 4.3. The auxiliary mixed actions and their basic properties

In this subsection we first select profiles of mixed actions $x(z, \tilde{\rho}) \in X(z), z \in S$, that depend on the vector $\tilde{\rho} = (\rho^i)_{i \in N}$ of individual discount rates, and $X^*(z) \subset X(z)$ consists of all limit points of $x(z, \tilde{\rho})$ as $\tilde{\rho} \rightarrow 0$.

The subset of states $\tilde{S}$ is defined as all states $z$ for which there is a stationary strategy $y, z \mapsto y(z) \in X^*(z)$, such that $P_{y}^\rho(z_1 \notin \tilde{C}) > 0$.

We select a stationary strategy $y : z \mapsto y(z) \in X^*(z)$ and a positive number $0 < \delta < 1$ that satisfy

$$P_{y}^\rho(z_1 \notin \tilde{C}) > 4\delta \quad \forall z \in \tilde{S}. \quad (25)$$

**Definition of $x(z, \tilde{\rho})$.** For every $\tilde{\rho} = (\rho^i)_{i \in N} \in (0, \infty)^N$, let $x(z, \tilde{\rho}), z \in S$, be an equilibrium of the single-stage game with the payoff function to player $i$

$$a \mapsto \rho^i g^i(z, a) + \sum_{z \in S} \mu(z, z', a) v^i_\rho(z').$$

Then, using the definition of $v^i_\rho$ and the fact that $x(z, \tilde{\rho})$ is a profile of mixed actions and therefore $x^{-i}(z, \tilde{\rho}) \in X^{-i}(z)$, we have

$$\rho^i v^i_\rho(z) = \min_{x^i} \max_{x^{-i}} \left( \rho^i g^i(z, x^{-i} \otimes x^i) + \sum_{z' \in S} \mu(z, z', x^{-i} \otimes x^i) v^i_\rho(z') \right)$$

$$\leq \max_{x^i} \left( \rho^i g^i(z, x^{-i}(z, \tilde{\rho}) \otimes x^i) + \sum_{z' \in S} \mu(z, z', x^{-i}(z, \tilde{\rho}) \otimes x^i) v^i_\rho(z') \right)$$

$$= \rho^i g^i(z, x(z, \tilde{\rho})) + \sum_{z' \in S} \mu(z, z, x(z, \tilde{\rho})) v^i_\rho(z').$$

where the minimum is over all $x^{-i} \in X^{-i}(z) := x_{j \neq i} \Delta(A^j(z))$ and the maximum is over all $x^i \in X^i(z) := \Delta(A^i(z))$.

**Definition of $X^*(z)$.** For every $\theta > 0$ let $X(z, \theta) \subset X(z)$ be the closure of the set $\{(x(z, \tilde{\rho})) : \max_{i \in N} \rho^i < \theta\}$. Let $X^*(z) \subset X(z)$ be the intersection of the sets $X(z, \theta)$, namely, $X^*(z) := \cap_{\theta, \theta} X(z, \theta)$.

As the sets $X(z, \theta)$ and $X(z, \theta)$ are nonempty and compact, and $X(z, \theta) \subseteq X(z, \theta')$ whenever $0 < \theta < \theta'$, the set $X^*(z)$ is nonempty and compact.
Note that if \( y \in X^*(z) \) then there is a sequence \( \{\tilde{\rho}(k)\} \) s.t. \( \rho^i(k) \to_{k \to \infty} 0 \) and \( x(z, \tilde{\rho}(k)) \to_{k \to \infty} y \). Therefore, by passing to the limit in inequality (26), for every \( z \in S, y \in X^*(z) \), and \( i \in N \), we have
\[
\sum_{z \in S} \mu(z', z, y) v^i(z') \geq 0. \tag{27}
\]

**Definition of \( \tilde{S}, \ y, \ \text{and} \ \delta.** The set of states \( \tilde{S} \), the stationary strategy profile \( y : z \mapsto y(z) \in X^*(z) \), and \( 0 < \delta < 1 \) are such that
\[
P^*_y(z_1 \notin C_z) > 4\delta \ \forall z \in \tilde{S},
\]
\[
\mu((S \setminus C_z) \cup \tilde{S}, z, x) = 0 \ \forall z \notin \tilde{S}, x \in X^*(z).
\]
An explicit construction of \( \tilde{S}, y, \ \text{and} \ \delta \) is as follows.

The set \( \tilde{S} \) is the set of all states \( z \) for which there is a sequence \( z = z_0, y_1, z_1, \ldots, z_k, \bar{z}_k \) such that \( v(z_k) \neq v(z), y_j \in X^*(z_{j-1}), \) and \( \mu(z_j, z_{j-1}, y_j) > 0 \), and \( y : z \mapsto y(z) \in X^*(z) \) is a stationary strategy such that for every state \( z \in \tilde{S} \) there is a sequence \( z = z_0, z_1, \ldots, z_k \) such that \( v(z_k) \neq v(z) \) and \( \mu(z_j, z_{j-1}, y(z_{j-1})) > 0 \).

An alternative longer definition, but somehow more explicit, can be obtained by defining the sequence of disjoint subsets of states \( S_k \) with \( z \in S_k \) iff \( k \) is the minimal positive integer for which there is a sequence \( z = z_0, y_1, z_1, \ldots, y_k, z_k \) such that \( v(z_k) \neq v(z), y_j \in X^*(z_{j-1}), \) and \( \mu(z_j, z_{j-1}, y_j) > 0 \). In that case,
\[
\tilde{S}_1 := \{ z : \exists y \in X^*(z) \ \text{s.t.} \ \mu(S \setminus C_z, z, y) > 0 \}
\]
and for every \( z \in \tilde{S}_1 \) we select an element \( y(z) \in X^*(z) \) with \( \mu(S \setminus C_z, z, y) > 0 \). Inductively, for \( k \geq 1 \),
\[
\tilde{S}_{k+1} := \{ z \notin \bigcup_{j \leq k} \tilde{S}_j : \exists y \in X^*(z) \ \text{s.t.} \ \mu(\tilde{S}_k, z, y) > 0 \}
\]
and for every \( z \in \tilde{S}_{k+1} \) we select an element \( y(z) \in X^*(z) \) with \( \mu(\tilde{S}_k, z, y(z)) > 0 \). We set
\[
\tilde{S} = \bigcup_{k=1}^{\infty} \tilde{S}_k.
\]
Next, we show that \( \tilde{S} \) is a proper subset of \( S \). Let \( z \in \text{arg max}_{x \in S} w(z) \), i.e., \( w(z) \geq w(z') \ \forall z' \in S \). By inequality (27) and implication (24), we deduce that (1) \( w(z) < \text{max}_{x \in S} w(z') \) for every \( z \in \tilde{S}_1 \), and (2) \( w(z) < \text{max}_{x \in \tilde{S}_k} w(z') \) for every \( z \in \tilde{S}_{k+1} \). Therefore, any state \( z \in \text{arg max}_{x \in S} w(z) \) (i.e., \( w(z) \geq w(z') \ \forall z' \in \tilde{S} \) is not in \( \tilde{S} \).

For every \( z \in S \setminus \tilde{S} \) we select an (arbitrary) element \( y(z) \in X^*(z) \), and define the stationary strategy \( y \) by \( y : z \mapsto y(z) \).

As the probability \( P^*_y \) that the state \( z_1 \) at time 1 is not in \( C_z \) is > 0, i.e., \( P^*_y(z_1 \notin C_z) > 0 \), for every \( z \in \tilde{S} \) we can select \( \delta > 0 \) such that \( P^*_y(z_1 \notin C_z) > 4\delta \) for all \( z \in \tilde{S} \).

Note that as \( y(z) \in X^*(z) \) for every \( z \in S \), inequality (27) implies that
\[
E^z_y v(z_1) \geq v(z) \ \forall z \in \tilde{S}.
\]

**Definition of \( B(z) \).** For every \( z \in S \) let
\[
B(z) := \{ a \in A(z) \ \text{s.t.} \ \mu((S \setminus C_z) \cup \tilde{S}, z, a) > 0 \}.
\]
Fix \( z \notin \tilde{S} \). By the definition of \( \tilde{S} \), for every \( x \in X^*(z) \), we have \( \mu((S \setminus C_z) \cup \tilde{S}, z, x) = 0 \). Therefore, for any action profile \( a \) in the support of \( x \in X^*(z) \), we have \( \mu((S \setminus C_z) \cup \tilde{S}, z, a) = 0 \). As \( X^*(z) \) is nonempty, we deduce that for any \( z \notin \tilde{S} \), \( B(z) \) is a proper subset of \( A(z) \).

Therefore, for \( z \notin \tilde{S} \), for every \( x \in X(z) \) there is a point \( x^* \in X(z) \) with \( x^*(B(z)) = 0 \) and \( x^*(a) \geq x(a) \) for all \( a \in A(z) \setminus B(z) \).

The norm distance between \( x \) and \( x^* \), \( \| x - x^* \| := \sum_{a \in A(z)} |x^*(a) - x(a)| \) (\( = \sum_{a \in A(z) \setminus B(z)} |x^*(a) - x(a)| \)), is equal to \( 2\delta(B(z)) \).

**Definition of \( \xi \).** Let \( \xi > 0 \) be a positive constant such that \( \mu(z', z, a) > 0 \) implies that \( \mu(z', z, a) > \xi \).

### 4.4. The strategy \( \tau \) and the time \( N_0 \)

**Lemma 3.** For every \( 1 > \varepsilon_1 > 0, \) and a positive integer \( \ell > 8/\varepsilon_1 \), there are functions \( (z_k, \tilde{\rho}_j)_{k=0}^{j=0} \rightarrow \tilde{\rho}_j \in (0, \alpha_0)^N \) (where \( j \) and \( j \) are nonnegative integers), and a sufficiently large positive integer \( N_0 \), such that if \( \tau \) is the profile of strategies with \( \tau(h, t) = x(z_1, \tilde{\rho}_j(t)) \), where \( \{\ell t\} \) is the largest integer that is \( \leq \ell t \), then for all \( z \in S \) and \( i \in N \),
\[
E^z_{\tau_i} v^i(z_\ell) \geq v^i(z) - \varepsilon_1/2 \quad \text{for every } (H_i)_{i \geq 0} \text{-stopping time } T
\]
and
\[
y^i_{N_0}(z, \tau) \geq v^i(z) - \varepsilon_1/2.
\]
\textbf{Proof.} For every fixed vector $\tilde{\rho} = (\rho_i^j)_{i\in N} \in (0,1)^N$, the stationary strategy $\alpha(\tilde{\rho}) : z \mapsto x(z, \tilde{\rho})$ satisfies

$$E^*_\alpha\left( \int_0^t \rho^i e^{-\rho^i t} g^i(z_t, x_t) dt + e^{-\rho^i t} v^i_{\rho^i}(z_t) \right) \geq v^i_{\rho^i}(z) \quad \forall t \geq 0.$$  \hspace{1cm} (30)

Let $\varepsilon_2 \leq \varepsilon_1/2$; e.g., $\varepsilon_2 = \varepsilon_1/2$. Following the proof in Mertens and Neyman (1981), we select an integrable function $\psi : (0,1] \to \mathbb{R}_+$ such that $\|dv_\psi(x)\| \leq \psi(x) \quad \forall x \in (0,1]$ (equivalently, $\int_{\rho^i}^{\rho^i_{\bar{\mu}}} \psi(x) dx \geq \|v^i_{\rho^i} - v^i_{\rho^i_{\bar{\mu}}}\| \quad \forall 0 < \rho, \tilde{\rho} \leq 1$ and $\forall i \in N$), and define the function $s : (0,1] \to [1,\infty)$ by

$$s(y) = \frac{12}{\varepsilon_2} \int_y^1 \frac{\psi(x)}{x} dx + y^{-1/2}.$$  

The function $s$ is strictly increasing and onto and its inverse function is denoted by $\lambda$.

Letting $\tilde{z}_j = z^j_1$ and $\check{H}_j = H^j_1$, we set

- $s_0^i \geq M$ with $M$ sufficiently large,
- $\rho^j_i = \lambda(s^j_i)$,
- $u^j_i = v^j_{\rho^j_i}$ and $y_j = \ell E^*_\alpha(\int_0^\infty g^i(z_t, x_t) dt \mid \check{H}_j)$, and
- $s^j_{i+1} = \max\{M, s^j_i + y_j - u^j_i(\tilde{z}_{j+1}) + 4\varepsilon_2\}$.

The inequalities that follow depend on a fixed player $i$ and for ease of notation we suppress the superscript $i$.

Inequality (30) implies, as in the proof of Theorem 3, that for every $\theta > 0$ there is $\alpha_0$ sufficiently small such that for $\rho^j_i < \alpha_0$ and $\lambda_j = \rho^j_i/\ell$, we have

$$E^*_\alpha(\lambda_j y_j + (1 - \lambda_j) u_j(\tilde{z}_{j+1}) \mid \check{H}_j) \geq u_j(\tilde{z}_j) - \theta \rho^j_i.$$  

Equivalently,

$$E^*_\alpha(u_j(\tilde{z}_{j+1}) - u_j(\tilde{z}_j) + \lambda_j(y_j - u_j(\tilde{z}_{j+1})) \mid \check{H}_j) \geq -\theta \lambda_j.$$  \hspace{1cm} (31)

Note the similarity between this last inequality and inequality (2.1) in Mertens and Neyman (1981). One minor difference, however, is the added term of $-\theta \lambda_j$ on the right-hand side of the inequality. Another one is that $y_j$ depends $\tilde{\rho}_j$ and is thus measurable with respect to $\check{H}_j$, while in inequality (2.1) in Mertens and Neyman (1981) the corresponding term is not a function of the past.

We first observe the following:

$$|s^j_{i+1} - s^j_j| \leq 6$$  \hspace{1cm} (32)

$$y_j - u_j(\tilde{z}_{j+1}) + 4\varepsilon_2 \leq s^j_{i+1} - s^j_j \leq y_j - u_j(\tilde{z}_{j+1}) + 4\varepsilon_2 + 21s_{j+1}=M.$$  \hspace{1cm} (33)

Since for every $1 > \eta > 0$, $s((1-\eta)y) - s(y) \to y \to 0$, $\infty$, (32) implies that there is $M_0$ such that for every $M \geq M_0$

$$|\rho^j_{i+1} - \rho^j_{j}| \leq \varepsilon_2 \rho^j_j/6.$$  \hspace{1cm} (34)

Thus for $M \geq M_0$,

$$6 \geq |s^j_{i+1} - s^j_j| \geq \frac{12}{\varepsilon_2} \int_{\rho^j_j}^{\rho^j_{i+1}} \frac{\psi(x)}{x} dx \geq \frac{6}{\varepsilon_2 \rho^j_j} \int_{\rho^j_j}^{\rho^j_{i+1}} \psi(x) dx \geq \frac{6}{\varepsilon_2 \rho^j_j} \|v^j_{\rho^j_{i+1}} - v^j_{\rho^j_j}\|.$$  

Hence,

$$\|v^j_{\rho^j_{i+1}} - v^j_{\rho^j_j}\| \leq \varepsilon_2 \rho^j_j.$$  \hspace{1cm} (35)

By the mean value theorem and inequality (33),

$$\int_{s^j_j}^{s^j_{i+1}} \lambda(y)dy = \int_{s^j_j}^{s^j_{i+1}} \lambda(y)dy - \int_{s^j_j}^{s^j_{i+1}} \lambda(y)dy \geq \rho^j_j(s^j_{i+1} - s^j_j) - \varepsilon_2 \rho^j_j \geq \rho^j_j(y_j - u_j(\tilde{z}_{j+1})) + 3\varepsilon_2 \rho^j_j.$$  \hspace{1cm} (36)
Replacing, in (31), \( u_j(\tilde{z}_{j+1}) \) by the larger number \( u_{j+1}(\tilde{z}_{j+1}) + \varepsilon_2 \rho_j^1 \) (see (35)), \( \rho_j^1(y_j - u_j(\tilde{z}_{j+1})) \) by the larger number \( \int_{y_j}^{\infty} \lambda(y)dy - \int_{y_j}^{\infty} \lambda(y_j)dy = 3 \varepsilon_2 \rho_j^1 \) (see (36)), and letting \( Y_j = u_j(\tilde{z}_j) - \int_{y_j}^{\infty} \lambda(y)dy \), we have (for \( \theta < \varepsilon_2 \))

\[
E_T(Y_j + Y_j \mid \mathcal{H}_j) \geq \varepsilon_2 \rho_j^1.
\]

(37)

\( Y_j \) is thus a bounded \((|Y_j| \leq 1)\) submartingale (by (37)), and therefore for any \( \mathcal{H}_j \)-stopping time \( T \), \( E_T^x Y_T \geq Y_0 \).

By summing the right-hand side inequalities of (33) for \( 0 \leq j < \ell n \), we have

\[
\sum_{j < \ell n} y_j \geq \sum_{j < \ell n} u_j(\tilde{z}_{j+1}) + s_{\ell n} - s_0 - 2 \sum_{j < \ell n} 1_{s_{j+1} = M} - 4 \ell n \varepsilon_2.
\]

(38)

For sufficiently large \( M \), we have \( E_T^x \sum_{j < \ell n} 1_{s_{j+1} = M} \leq E_T^x \sum_{j < \ell n} E_{2\lambda(\lambda_0)}^{\varepsilon_2} \rho_j^1 \leq \frac{1}{E_{2\lambda(\lambda_0)}^{\varepsilon_2}} E_T^x(Y_{\ell n} - Y_0) \).

Therefore, for all sufficiently large \( n \),

\[
E_T^x \int_0^n g^i(z_t, x_t) dt = \frac{1}{\ell n} E_T^x \sum_{j < \ell n} y_j \geq v^i(z) - 6 \varepsilon_2.
\]

This completes the proof of inequality (29).

For every real number \( x \), let \( [x] \) denote the smallest integer that is \( \geq x \).
Note that if \( T \) is an \( \mathcal{H}_\ell \)-stopping time, then \( \varepsilon(T) \) is an \( \mathcal{H}_j \)-stopping time. As \( Y_j \) is a submartingale, we have \( E_T^x Y_{\varepsilon(T)} \geq Y_0 \).

Therefore, as \( |Y_j - v^i(\tilde{z}_j)| < \varepsilon_2/16 \) for \( M \) sufficiently large, we deduce that

\[
E_T^x v^i(\tilde{z}_{\varepsilon(T)}) \geq v^i(z) - \varepsilon_2/8.
\]

By Lemma 1 and the assumption \( ||\mu|| = 1 \), we have \( E |v^i(z_T) - v^i(\tilde{z}_{\varepsilon(T)})| \leq 1/\ell \leq \varepsilon_2/8 \).

Therefore,

\[
E_T^x v^i(z_T) \geq v^i(z) - \varepsilon_2/4 \geq v^i(z) - \varepsilon_1/2.
\]

This completes the proof of inequality (28). \( \square \)

4.5. An auxiliary lemma

**Lemma 4.** Let \( C \subset S \) be a subset of states. Let \( T = T_C = \min_{s \in C} \{ s : z_s \in C \} \). Then for every \((s_k)_{k \geq 0}\)-discritematized correlated strategy \( \sigma \) and \( n > t \),

\[
P_\sigma(t < T \leq n \mid \mathcal{H}_t) = E_\sigma \left( \int_t^{T \wedge n} \mu(C, z_s, x_s) ds \mid \mathcal{H}_t \right)
\]

(39)

where \( T \wedge n := \min(T, n) \).

**Proof.** Equality (39) holds trivially on \( z_t \in C \). Assume that \( z_t \notin C \). It suffices to prove the lemma for \( t = 0 \).

We provide two proofs. The first one approximates the probability that \( T \leq n \) by summing the approximations of the probabilities that \( s < T \leq s + \delta \), where \( \delta > 0 \) is small, and \( s \) ranges over a grid of \( n/\delta \) points. The error term of each approximation is \( O(\delta^2) \). The sum of the approximations is the desired formula and the cumulative error term is \( O(\delta) \). The second proof derives the formula without an approximation.

The mixed action profile \( x_t \) that the \((s_k)_{k \geq 0}\)-discretematized correlated strategy \( \sigma \) selects at time \( s \) is a function of the random list \((s_k \wedge S, \mathcal{H}_{s_k}, k \geq 0)\).

For \( s_j \leq s < s_{j+1} \) we denote this action profile by \( x_{j,s} \). Note that \( x_{j,s} \) is measurable w.r.t. \( \mathcal{H}_{s_j} \) and is defined for every \( s_j \leq s \). It coincides with \( x_s \) on \( s_j \leq s < s_{j+1} \), but may differ from \( x_s \) on \( s_{j+1} \leq s \).

Given \( s_j \leq s < s_{j+1} \) and \( T > s \), the conditional \( P_\sigma^s \) probability, given \( \mathcal{H}_{s_j} \), that \( s < T = s_{j+1} \leq s + \delta \) equals

\[
\int_s^{s+\delta} \mu(C, z_{j,t}, x_{j,t}) \mu^i(z_{j,t}, x_{j,t}, u) du dt.
\]

If there is no state change in the interval \([s, s + \delta]\), then \( x_{j,t} = x_t \) for every \( s \leq t \leq s + \delta \). Therefore, given \( s_j \leq s < s_{j+1} \), the \( P_\sigma^s \) probability that \( x_t = x_{j,t} \) for every \( s \leq t \leq s + \delta \) is at least \( 1 - \delta \cdot ||\mu|| \).
Therefore, the above conditional probability, given $\mathcal{H}_s^S$, is within $O(\delta^2)$, the conditional probability, given $\mathcal{H}_s^T$, of $\int_{s_j}^{s_{j+1}} \mu(C, z_t, x_t) dt$.

Therefore, the $P^S_\alpha$ probability that $T \leq n$ is, within $O(\delta)$, the expectation w.r.t. $P^S_\alpha$ of $\int_0^T \mu(C, z_t, x_t) dt$. As this holds for any $\delta > 0$ the result follows. \hfill $\square$

Now we present the alternative proof.

Note that on $s_j < n$ the conditional expectation of the indicator of the event $s_j < T = s_{j+1} \leq n$, given the history up to time $s_j$, equals $\int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t}) e^{\int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t})} du$ in symbols,

$$E_\sigma(\mathbb{1}(s_j < T = s_{j+1} \leq n) | \mathcal{H}_{s_j}) = \int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t}) e^{\int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t})} du,$$

where $\mathbb{1}(\ast)$ denotes the indicator of the event $\ast$.

The term $e^{\int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t})} du$ equals the conditional probability of the event $t < s_{j+1}$, given the history up to time $s_j$.

Therefore,

$$\int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t}) e^{\int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t})} du = \int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t}) E_\sigma(\mathbb{1}(t < s_{j+1}) | \mathcal{H}_{s_j}) dt.$$

On $s_j \leq t < s_{j+1}$, $z_{s_j} = z_t$ and $x_{j,t} = x_t$. Therefore,

$$\int_{s_j}^{s_{j+1}} \mu(C, z_{s_j}, x_{j,t}) E_\sigma(\mathbb{1}(t < s_{j+1}) | \mathcal{H}_{s_j}) dt = E_\sigma(\int_{s_j}^{s_{j+1}} \mu(C, z_t, x_t) dt | \mathcal{H}_{s_j}).$$

We conclude that

$$E_\sigma(\mathbb{1}(s_j < T = s_{j+1} \leq n) | \mathcal{H}_{s_j}) = E_\sigma(\int_{s_j}^{s_{j+1}} \mu(C, z_t, x_t) dt | \mathcal{H}_{s_j}).$$

Taking the expectation on both sides and summing over all nonnegative integers $j$, we have

$$P^S_\alpha(T \leq n) = E_\sigma^S \sum_{j=0}^{\infty} \mathbb{1}(s_j < T = s_{j+1} \leq n) = E_\sigma^S(\int_0^T \mu(C, z_t, x_t) dt). \hfill \square$$

4.6. The partitioning of the state space: the sets $S_1$ and $S_2$

The subset $S_1$ of $S \setminus \bar{S}$ is defined below so as to guarantee the existence of a correlated strategy $\tau$ such that for $z \in S_1$ (1) the distance between the $P^T_\tau$-distribution and the $P^T_{\bar{\tau}}$-distribution of the state process up to time $N_0$ is small, (2) the difference between the expectation w.r.t. $P^T_\tau$ and the expectation w.r.t. $P^T_{\bar{\tau}}$ of the accumulated payoff up to time $N_0$ is small, and (2) the state process remains in $C_z \setminus \bar{S}$ with $P^T_{\bar{\tau}}$-probability 1. See Section 4.7.

Fix $\varepsilon_2 > 0 (\varepsilon_2 = O(\varepsilon))$; e.g., $\varepsilon_2 = \varepsilon_1 / 2$ or $\varepsilon_2 = \varepsilon$. The subset $S_1$ of $S \setminus \bar{S}$ depends on the constant $\varepsilon_2 > 0$ and the strategy $\tau$ as follows. Define the $\mathcal{H}_t^T_{t \geq 0}$-stopping time $T$ by $T = \min[t : z_t \in (S \setminus C_z) \cup \bar{S}]$. Set

$$S_1 = \{z \in S : P^T_\tau(T \leq N_0) \leq 2\varepsilon_2\}$$

and

$$S_2 = S \setminus (\bar{S} \cup S_1) = \{z \in S \setminus \bar{S} : P^T_\tau(T \leq N_0) > 2\varepsilon_2\}.$$

Note that $T = 0$ on $z_0 \in \bar{S}$. Therefore, by the definition of $S_1$, we have $S_1 \cap \bar{S} = \emptyset$.

**Lemma 5.** For every $z \in S_1$,

$$E^S_\tau \left( \int_0^{N_0} x(z_t, \bar{\sigma}_t) B(z_t) dt \right) \leq 2\varepsilon_2 / \varepsilon. \quad (40)$$
Proof. It follows from the selection of $\xi > 0$ that
\[
\mu(\{S \cap C_z \cup \tilde{S}, z_t, x(z_t, \tilde{\rho}_{t[1]})) \geq \xi x(z_t, \tilde{\rho}_{t[1]}|B(z_t)) \text{ on } t < T.
\] (41)
Therefore (using Lemma 4), if $z \in S_1,$ then
\[
2\varepsilon_2 \geq P^z_\varepsilon(T \leq N_0) = E^z_\varepsilon \left( \int_0^{T \wedge N_0} \mu(\{S \cap C_z \cup \tilde{S}, z_t, x(z_t, \tilde{\rho}_{t[1]})) \) dt \geq E^z_\varepsilon \left( \int_0^{T \wedge N_0} \xi x(z_t, \tilde{\rho}_{t[1]}|B(z_t)) dt \right). \tag{42}
\]

4.7. Definition of the correlated strategy $\bar{\tau}$

We start by defining an auxiliary correlated strategy $\bar{\tau}.$

\[
\bar{\tau}(h, t) = \begin{cases} 
    y(z_t) & \text{if } z_0 \in \tilde{S} \\
    \tau(h, t) & \text{if } z_0 \in S_2 \\
    x^*(z_t, \tilde{\rho}_{t[1]}) & \text{if } z_0 \in S_1,
\end{cases}
\]
where $x^*(z_t, \tilde{\rho}_{t[1]})$ is a point in $\{x \in X(z_t) : x(B(z_t)) = 0\}$ with $\|x^*(z_t, \tilde{\rho}_{t[1]}) - x(z_t, \tilde{\rho}_{t[1]})\| \leq 2x(z_t, \tilde{\rho}_{t[1]}|B(z))$ if $x(z_t, \tilde{\rho}_{t[1]}|B(z)) > 0,$ and $x^*(z_t, \tilde{\rho}_{t[1]}) = x(z_t, \tilde{\rho}_{t[1]})$ otherwise. Recall that for $z \notin \tilde{S}, B(z)$ is a proper subset of $A(z),$ and therefore $\{x \in X(z) : x(B(z)) = 0\}$ is nonempty.

Note $\tau(h, t) = x(z_t, \tilde{\rho}_{t[1]})$ is a profile of mixed actions. However, $x^*(z_t, \tilde{\rho}_{t[1]})$ need not be a profile of mixed actions. Therefore, the correlated strategy $\bar{\tau}$ need not be a profile of strategies.

Recall that by the definition of $\tilde{S}$ it follows that for every $z \in S \setminus \tilde{S}$ and every $x \in X^*(z),$ we have $\mu(\tilde{S}, z, x) = 0$ and $\sum_{z' \in S} \mu(z', z, x) \nu(z') = 0.$

Lemma 6. For every $z \in S_1,$ $P^z_\varepsilon(z_t \in C_z \setminus \tilde{S}, \forall 0 \leq t \leq N_0) = 1$ and therefore $P^z_\varepsilon(z_{N_0} \in C_z \setminus \tilde{S}) = 1.$

Proof. Follows from the definition of $\bar{\tau}$ on $z_0 \in S_1.$

Lemma 7. For every $z \in S_1,$ the norm distance between the $P^z_\varepsilon$-distribution and the $P^z_\varepsilon$-distribution of the state process $(z_t)_{0 \leq t \leq T \wedge N_0}$ is $\leq 4\varepsilon_2 / \xi.$

Proof. By the definition of $\bar{\tau}, \|\bar{\tau}(h, t) - \tau(h, t)\| \leq 2x(z_t, \tilde{\rho}_{t[1]}|B(z))$ on $z_0 = z \in S_1.$ Therefore, the result follows from Lemmas 12 (in Section 6) and 5.

Lemma 8. For every $z \in S_1,$
\[
E^z_\varepsilon \int_0^{N_0} g^i_t dt \geq E^z_\varepsilon \int_0^{N_0} g^i_t dt - 2(2N_0 + 2)\varepsilon_2 / \xi \geq N_0 \left( v^i(z) - \varepsilon_1 / 2 - 0(\varepsilon_2) \right).
\]

Proof. In the following sequence of equations, equality (43) follows from the fact that on the plays that are compatible with the strategy $\bar{\tau}$ we have $g^i_t = g^i(z_t, x^*(z_t, \tilde{\rho}_{t[1]})).$

Inequality (44) follows from:

(i) the inequality $0 \leq g^i(z_t, x(z_t, \tilde{\rho}_{t[1]})) \leq 1,$
(ii) $\tilde{\rho}_{t[1]}$ and (therefore also) $g^i(z_t, x(z_t, \tilde{\rho}_{t[1]}))$ depend only on the state process in time $s \in [0, N_0],$ and
(iii) half the norm distance between the distributions defined by $\tau$ and $\bar{\tau}$ on this state process is (by Lemma 7) less than or equal to $2\varepsilon_2 / \xi.$

Inequality (45) follows from the inequality
\[
g^i(z_t, x^*(z_t, \tilde{\rho}_{t[1]})) \geq g^i(z_t, x(z_t, \tilde{\rho}_{t[1]})) - \|x(z_t, \tilde{\rho}_{t[1]}) - x^*(z_t, \tilde{\rho}_{t[1]})\|.
\]

Inequality (46) follows from the equality
\[
E^z_\varepsilon \int_0^{N_0} \|x(z_t, \tilde{\rho}_{t[1]}) - x^*(z_t, \tilde{\rho}_{t[1]})) \| dt = E^z_\varepsilon \int_0^{T \wedge N_0} \|x(z_t, \tilde{\rho}_{t[1]}) - x^*(z_t, \tilde{\rho}_{t[1]})\| dt + E^z_\varepsilon \int_{T \wedge N_0}^{N_0} \|x(z_t, \tilde{\rho}_{t[1]}) - x^*(z_t, \tilde{\rho}_{t[1]})\| dt
\]
and the inequality
\[
E_{\mathcal{F}}^{\infty} \int_{T \wedge N_0} \|x(z_t, \tilde{\rho}_{[t]}^-) - x^*(z_t, \tilde{\rho}_{[t]}^+)| dt \leq P_{\mathcal{F}}^z(T < N_0)2N_0.
\]

Inequality (47) follows from:

(i) the inequality \( P_{\mathcal{F}}^z(T < N_0) \leq 2\varepsilon_2 \) for \( z \in S_1 \),

(ii) the inequality \( \|x(z_t, \tilde{\rho}_{[t]}^-) - x^*(z_t, \tilde{\rho}_{[t]}^+)| \leq 2\varepsilon_2 \), and

(iii) the inequality \( E_{\mathcal{F}}^{\infty} \int_{T \wedge N_0} 2x(z_t, \tilde{\rho}_{[t]}^+)|B(z_t)| dt \leq 4\varepsilon_2 / \xi \).

Finally, inequality (48) follows from inequality (29).

\[
E_{\mathcal{F}}^{\infty} \int_{0}^{N_0} g_t^1 dt = E_{\mathcal{F}}^{\infty} \int_{0}^{N_0} g_t^1(z_t, x^*(z_t, \tilde{\rho}_{[t]})) dt
\]
\[
\geq E_{\mathcal{F}}^{\infty} \int_{0}^{N_0} g_t^1(z_t, x^*(z_t, \tilde{\rho}_{[t]})) dt - 2N_0\varepsilon_2 / \xi
\]
\[
\geq E_{\mathcal{F}}^{\infty} \int_{0}^{N_0} g_t^1(z_t, x(z_t, \tilde{\rho}_{[t]})) dt - 2N_0\varepsilon_2 / \xi
\]
\[
- E_{\mathcal{F}}^{\infty} \int_{0}^{N_0} \|x(z_t, \tilde{\rho}_{[t]}^-) - x^*(z_t, \tilde{\rho}_{[t]}^-)| dt
\]
\[
\geq E_{\mathcal{F}}^{\infty} \int_{0}^{N_0} g_t^1 dt - 4\varepsilon_2 N_0 - \varepsilon_2 / \xi (4 + 2N_0)
\]
\[
\geq N_0 \left(v^1(z) - \varepsilon / 2 - O(\varepsilon_2)\right). \quad \Box
\]

4.8. The discretization \( \tilde{\tau} \) of \( \bar{\tau} \) and the strategy \( \bar{\tau} \)

Recall that conditional on \( h_j \), the strategy \( \bar{\tau} \) on the time-interval \([j, j + 1)\) coincides with a stationary strategy, denoted by \( \bar{\tau}(h_j) : z \mapsto \bar{\tau}(h_j)(z) \).

We discretize \( \bar{\tau} \) into a pure-action strategy \( \tilde{\tau} \) as follows. Let \( \ell \) be a sufficiently large positive integer, and conditional on \( z_j := (z_0, z_1, \ldots, z_j) \), \( j \) a nonnegative integer. The average of \( \bar{\tau} \) on time intervals \([j + k/\ell, j + (k + 1)/\ell)\), \( k \) a positive integer with \( 0 \leq k < \ell \), is within \( O(1/\ell \alpha) \) of \( \bar{\tau}(h_j) \) - the average mixed action of \( \bar{\tau} \) over that interval. For example, define nonnegative integers \( \ell(z, a) \geq \ell \bar{\tau}(h_j)(z)(a) - 1 \), where \( z \in S \) and \( a \in A(z) \), such that \( \sum_{a \in A(z)} \ell(z, a) = \ell \). For \( B \subset A(z) \) we set \( \ell(z, B) := \sum_{a \in B} \ell(z, a) \). Let \( \tilde{\tau} \) be a pure-action strategy that conditional on \( h_j \), (1) \( \bar{\tau}(h_j)(z, t) = \bar{\tau}(h_j)(z, t + i/\ell^2) \) if \( j + k/\ell \leq t < j + (k + 1)/\ell \leq j + 1 \), (2) for every \( 0 \leq k < \ell \), \( z \in S \), and \( a \in A(z) \), \( \sum_{0 \leq i < \ell} \bar{\tau}(h_j)(z, j + k/\ell + i/\ell^2)(a) = \ell(z, a) \), and (3) \( \bar{\tau}(h_j)(B(z_j)) = 0 \) on \( z_j \notin \mathcal{S} \). Therefore, conditional on \( z_j \), for every \( z \in S \), \( B \subset A(z) \), and \( 0 \leq k < \ell \), we have

\[
E_{\mathcal{F}}^{\infty} \int_{j + k/\ell}^{j + (k + 1)/\ell} \bar{\tau}(h_j)(z, t) dt = \frac{1}{\ell^2} \left| 0 \leq i < \ell : \bar{\tau}(h_j)(z, j + k/\ell + i/\ell^2)(B) = 1 \right| \geq \ell(z, B)/\ell^2 \geq \bar{\tau}(h_j)(z)(B)/\ell - 1/\ell^2.
\]

Therefore, for \( z \notin \mathcal{S} \),

\[
\left\| \frac{1}{\ell} \bar{\tau}(h_j)(z) - \frac{1}{\ell^2} \bar{\tau}(h_j)(z, t) dt \right\| = 2 \max_{B \subset A(z), B(z)} \frac{\bar{\tau}(h_j)(z)(B) - \ell(z, B)/\ell}{\ell} \leq 2|A(z)|/\ell^2 = O(1/\ell^2). \quad (49)
\]
For two nonnegative integers \( k, \ell \) we denote by \( z^k_\ell \) the sequence of states \( z_0, z_1/\ell, z_2/\ell, \ldots, z_k/\ell \). By Lemma 12 in Section 6, conditional on \( z_0 \in S_1 \),

\[
\sum_{z^k_\ell_{N_0}} |P^k_\ell (z^k_\ell_{N_0}) - P^k_\ell (z^k_{\ell N_0})| \leq N_0(4 + 2|A|)/\ell = O(1/\ell),
\]

where the sum is over all realizations \( z^k_{\ell N_0} \) of the random sequence \( z^k_{\ell N_0} = z_0, z_1/\ell, z_2/\ell, \ldots, z_{N_0} \).

**Lemma 9.** For every \( z \in S_1 \), \( P^k_\ell (z_t | z_{\ell} \in S_2 \setminus \tilde{S}) \) \( 0 \leq t \leq N_0 \) 1, and

\[
\gamma^k_{\ell N_0} (z, \tilde{\tau}) \geq v^k(z) - \varepsilon_1 - O(\varepsilon_2) \text{ for all } \ell \text{ sufficiently large.}
\]

**Proof.** The first assertion follows from \( \tilde{\tau} \) being a pure-action strategy and from the imposition of \( \tilde{\tau}(h, t) \notin B(z_t) \) on \( z_t \notin \tilde{S} \).

For every nonnegative integer \( k \) (for every strategy profile and every initial state), the conditional probability, given \( h_{k/\ell} \), that \( z_t = z_{k/\ell} \) for all \( k/\ell \leq t < (k + 1)/\ell \) is greater than or equal to \( 1 - 1/\ell \). On \( z_t = z_{k/\ell} \) for all \( k/\ell \leq t < (k + 1)/\ell \), we have (by (49))

\[
\int_{k/\ell}^{(k+1)/\ell} g^k_t(z_t, \tilde{\tau}(h, t)) dt \geq \int_{k/\ell}^{(k+1)/\ell} g^k_t(z_t, \tilde{\tau}(h, t)) dt - |A|/\ell^2.
\]

Therefore, for \( z \in S_1 \) and \( z_{k/\ell} \in C_2 \setminus \tilde{S} \),

\[
E^k_\ell (\int_{k/\ell}^{(k+1)/\ell} g^k_t dt | z^k_{\ell}) \geq E^k_\ell (\int_{k/\ell}^{(k+1)/\ell} g^k_t dt | z^k_{\ell}) - (|A| + 1)/\ell^2.
\]

Therefore, using (51) and (50),

\[
N_0 \gamma^k_{\ell N_0} (z, \tilde{\tau}) = \sum_{k=0}^{\ell N_0 - 1} E^k_\ell E^k_\ell (\int_{k/\ell}^{(k+1)/\ell} g^k_t dt | z^k_{\ell})
\]

\[
\geq \sum_{k=0}^{\ell N_0 - 1} E^k_\ell E^k_\ell (\int_{k/\ell}^{(k+1)/\ell} g^k_t dt | z^k_{\ell}) - (|A| + 1)N_0 \ell
\]

\[
= \sum_{z^k_{\ell N_0}} P^k_\ell (z^k_{\ell N_0}) \sum_{k=0}^{\ell N_0 - 1} E^k_\ell (\int_{k/\ell}^{(k+1)/\ell} g^k_t dt | z^k_{\ell}) - (|A| + 1)N_0 \ell
\]

\[
\geq \sum_{z^k_{\ell N_0}} P^k_\ell (z^k_{\ell N_0}) \sum_{k=0}^{\ell N_0 - 1} E^k_\ell (\int_{k/\ell}^{(k+1)/\ell} g^k_t dt | z^k_{\ell}) - (|A| + 1)N_0 \ell
\]

\[
\geq N_0 \gamma^k_{\ell N_0} (z, \tilde{\tau}) - (|A| + 1)N_0 \ell - (2|A| + 4)N_0^2 \ell.
\]

Using Lemma 8 in Section 4.7 we conclude that for all sufficiently large \( \ell \), we have \( \gamma^k_{\ell N_0} (z, \tilde{\tau}) \geq v^k(z) - \varepsilon_1 - O(\varepsilon_2) \). \( \square \)

The strategy \( \tilde{\tau} \) plays in blocks of random sizes that are \( \leq N_0 \), and in each block it follows \( \tilde{\tau} \) as if the game were starting at the state of the start of the block.

The duration of a block that starts at a state \( z \in S_1 \) is \( N_0 \).

The duration of a block that starts at a state \( z \in \tilde{S} \) is 1.

A block that starts at time \( t \) in state \( z \in S_2 \) continues until the first time \( t + j/\ell \), with \( j \) being a positive integer and \( z \notin C_2 \setminus \tilde{S} \) for some \( t < s \leq t + j/\ell \), if this time is \( \leq t + N_0 \). Otherwise it continues until time \( t + N_0 \).

Let \( \tilde{t}_k \) be the time at the end of the \( k \)-th block of \( \tilde{t}_0 = 0 \), and \( \tilde{z}_k \) be the state at the end of the block, i.e., \( \tilde{z}_k = z_{\tilde{t}_k} \).

The strategy \( \tilde{\tau} \) is given by \( \tilde{\tau}(h, t) = \tilde{\tau}(\tilde{z}_k) \), whenever \( \tilde{t}_k \leq t < \tilde{t}_{k+1} \), and where \( \tilde{z}_k \) is the left translation of the play \( \tilde{h} \) by \( \tilde{t}_k \); i.e., if \( \tilde{h} = (z_t, x_t)_{t \geq 0} \), then \( \tilde{t}_k \) \( \tilde{h} = (z_{\tilde{t}_k}, x_{\tilde{t}_k})_{\tilde{t}_k \geq 0} \).
Let us formally define the stopping times $\tilde{t}_k$ inductively. $\tilde{t}_{k+1}$ is a function of $\tilde{t}_k$ and $z_k$, and of the process that follows time $\tilde{t}_k$. If $z_{k+1} \in S_1$ then $\tilde{t}_{k+1} = \tilde{t}_k + N_0$. If $z_{k+1} \in S_2$ then $\tilde{t}_{k+1} = N_0 \wedge \inf\{ j/\ell : N \ni j, \exists j/\ell \geq s > \tilde{t}_k \text{ s.t. } z_s \in (S \setminus C_2) \cup \bar{S} \}$. If $z_{k+1} \in \bar{S}$ then $\tilde{t}_{k+1} = \tilde{t}_k + 1$.

**Lemma 10.** The discrete-time process $\tilde{z}_0, \tilde{z}_1, \ldots$, with the probability $P_{\tilde{t}_k}$ is a homogeneous Markov chain process with a transition matrix $F$ that obeys the following properties.

1. $F_{\tilde{z}_0, \tilde{z}_1} := \sum_{\tilde{z} \in C_1} F_{\tilde{z}, \tilde{z}'} = 1$ for $\tilde{z} \in S_1$.
2. $F_{\tilde{z}, S_1 \cup \bar{S}} > 4\delta - O(1/\ell) \geq 3\delta$ for $\tilde{z} \in \bar{S}$ and $\ell$ sufficiently large.
3. $F_{\tilde{z}, S_1 \cup \bar{S}} \geq 2\delta z - O(1/\ell) \geq \delta z$ for $\tilde{z} \in S_2$ and $\ell$ sufficiently large, and
4. $\sum_{\tilde{z} \in S} F_{\tilde{z}, \tilde{z}'}(z') \geq v^l(z) - \epsilon_1(z \in S_2 \cup \bar{S})$ for every $z \in S$, $i \in N$, and $\ell$ sufficiently large.

**Proof.** By the definition of $\tilde{t}_k$, the $P_{\tilde{t}_k}$ conditional distribution of $\tilde{z}_{k+1}$, given $\tilde{z}_0, \ldots, \tilde{z}_k$, depends only on $\tilde{z}_k$. Therefore, the stochastic $\tilde{z}_0, \ldots, \tilde{z}_k, \ldots$, defined by $\tilde{t}_k$, is a homogeneous Markov chain.

Equality 1) follows from the definition of $\tilde{t}_k$ on $z_0 \in S_1$. Recall that the norm distance between the $P_{\tilde{t}_k}$ and the $P_{\tilde{t}_k}$ distribution of $z_0, z_1, \ldots, z_{N_0}$ is $\leq O(1/\ell)$. For $z \in \bar{S}$, $F_{\tilde{z}, S_1 \cup \bar{S}} = P_{\tilde{t}_k}(z_1 \in S \cap C_2) = P_{\tilde{t}_k}(z_1 \in S \setminus C_2) \geq P_{\tilde{t}_k}(z_1 \in S \setminus C_2) - O(1/\ell) > 4\delta - O(1/\ell) \geq 3\delta$ for sufficiently large $\ell$. Therefore, 2) holds. For $z \in S_2$, $F_{\tilde{z}, S_1 \cup \bar{S}} \geq P_{\tilde{t}_k}(z \notin S_1 \text{ s.t. } j/\ell \in (S \setminus C_2) \cup \bar{S}) - O(1/\ell) \geq 2\delta z - O(1/\ell) \geq \delta z$ (where the last inequality uses the fact that for every stopping time $T$ and every strategy $\sigma$, $P_{\tilde{t}_k}(z_t = z_T \forall T \leq s \leq T + 1/\ell) \geq 1 - O(1/\ell)$). Therefore, as $P_{\tilde{t}_k}(T \leq N_0) \geq 2\delta z$, 3) follows.

By 1), for $z \in S_1$ we have $\sum_{\tilde{z} \in S} F_{\tilde{z}, \tilde{z}'}(z') = v^l(z)$; for $z \in \bar{S}$ we have $\sum_{\tilde{z} \in S} F_{\tilde{z}, \tilde{z}'}(z') = \sum_{\tilde{z} \in S} P_{\tilde{t}_k}(z_1 = z') v^l(z') \geq \sum_{\tilde{z} \in S} P_{\tilde{t}_k}(z_1 = z') v^l(z') - O(1/\ell) \geq v^l(z) - O(1/\ell)$, and for $z \in S_2$, using (28), we have $\sum_{\tilde{z} \in S} F_{\tilde{z}, \tilde{z}'}(z') \geq \sum_{\tilde{z} \in S} F_{\tilde{z}, \tilde{z}'}(z) - \epsilon_1/2 - O(1/\ell)$. Therefore, 4) follows.

4.9. A probabilistic lemma

The following lemma appears implicitly in Solan and Vieille (2002).

**Lemma 11.** Assume that $0 < \epsilon := \epsilon_2 < 1$ and $\epsilon_1 < \delta d^2/\ell$. Then

(a) All ergodic classes of the Markov chain $\tilde{z}_j$ are subsets of $S_1$.
(b) $E[\{ 0 \leq j < \ell : \tilde{z}_j \notin S_1 \}] \ll \infty$, and
(c) $w_j := w(\tilde{z}_j) \to j \to \infty w_\infty$, and therefore $v^l_j := v^l(\tilde{z}_j) \to j \to \infty v^l$. \forall i \in N,

and, if $\epsilon_1 < \delta d^2/\ell$.

(d) $E[\{ 0 \leq j < \ell : \tilde{z}_j \in \bar{S} \subseteq S_1 \}] \ll \frac{\ell}{\ell^2} \delta d$, and
(e) $\forall i \in N, E[v^l_j \mid z_0] \to v_0^l - \epsilon$ and $E(v_\infty \mid z_0) \to v_0^l - \epsilon$.

**Proof.** Let $\tilde{F}_j$, $j \geq 0$, be the algebra generated by $\tilde{z}_0, \ldots, \tilde{z}_j$. Consider the sequences $w_j$ and $Y_j$ of real-valued random variables $w_j = w(\tilde{z}_j)$ and $Y_j = w_\infty + \delta_j$, where $\delta_j = \delta(\tilde{z}_j)$ and $\epsilon \geq \delta(z) \geq 0$ is defined as follows. For $z \in \bar{S}$, we set $\delta(z) = \epsilon d^2 = \delta(0)$, for $z \in S_2$ we set $\delta(z) = \delta(2) = \delta d^2/4 - \epsilon_1$, and for $z \in S_1$ we set $\delta(z) = 0$.

We claim that

$$E(Y_{j+1} \mid \tilde{F}_j) \geq Y_j + \delta(\tilde{z}_j) \geq Y_j + \delta(\tilde{z}_j) I(\tilde{z}_j \in S_2 \cup \bar{S}).$$

As the process $\tilde{z}_j$ is a stationary Markov chain, it suffices to prove inequality (52) for $j = 0$.

On $z_0 \in S_1$, we have $Y_0 = w_0^l$, $w_1 = w_0$, and $Y_1 = w_0 + \delta(\bar{z}_0)$ a.e. Therefore, inequality (52) holds on $z_0 \in S_1$.

On $\bar{Z}_0 \in \bar{S}$, we have $E(w_1 \mid z_0) = w_0$ and the probability that $|w_1 - z_0| \geq d$ is $\geq 3\delta^2$. Therefore, $E(w_1^2 \mid z_0) - w_0^2 \geq 3\delta^2$, implying that $E(Y_1 \mid z_0) \geq w_0^2 + 3\delta^2 = Y_0 + 2\delta(0) \geq Y_0 + \delta(0)$.

We partition $S_2$ into two subsets. $S_2(\epsilon) := \{ z \in S_2 : P(w(\tilde{z}_1) \neq w(\tilde{z}_0) \mid \tilde{z}_0 = z) \geq \epsilon/2 \}$ and $S_2(\neq) := S_2 \setminus S_2(\epsilon)$. The inequality $F_{\tilde{z}, C_1} + F_{\tilde{z}, \bar{S}} \geq \epsilon$ for $z \in S_2$ implies that for $z \in S_2 \setminus S_2(\epsilon)$ we have $P(\tilde{z}_1 \in \bar{S} \mid \tilde{z}_0 = z) \geq \epsilon/2$.

On $\bar{Z}_0 \in S_2(\epsilon)$, the (conditional) probability that $|w_1 - z_0| \geq d$ is $\geq \epsilon/2$, and $E(w_1 \mid z_0) \geq w_0 - \epsilon_1$. Therefore, on $\bar{Z}_0 \in S_2(\epsilon)$,

$$E(w_1^2 \mid \bar{z}_0) = E((w_1 - w_0)^2 + 2w_0 w_1 - w_0^2 \mid \bar{z}_0) \geq d^2\epsilon/2 + 2w_0(w_0 - \epsilon_1) - w_0^2 \geq w_0^2 + \delta(2) = Y_0 + \delta(2),$$

where the last inequality follows from $2\delta(2) = \delta d^2/2 - \epsilon_1 \leq d^2\epsilon/2 - 2\epsilon_1$. 

Everywhere, and thus also on \( z_0 \in S_2 \setminus S_2(\neq) \), \( E(w_1 | z_0) \geq w_0 - \epsilon_1 \). The conditional probability that \( \bar{z}_1 \in \bar{S} \), given \( z_0 \in S_2 \setminus S_2(\neq) \), is \( \geq \epsilon/2 \). Therefore, on \( z_0 \in S_2 \setminus S_2(\neq) \), using the convexity of \( x \mapsto x^2 \),

\[
E(Y_1 | \bar{z}_0) \geq (E(w_1 | \bar{z}_0))^2 + \delta(0)\epsilon/2 \geq (w_0 - \epsilon_1)^2 - \epsilon_1^2 + \delta^2\epsilon/2 = w_0^2 - 2\epsilon_1 + \delta^2\epsilon/2 = w_0^2 + 2\delta(2) = Y_0 + \delta(2).
\]

This completes the proof of inequality (52).

Inequality (52) shows that \( Y_1 \) is a bounded \( (0 \leq Y_1 \leq 2) \) submartingale, and therefore it converges a.e. to \( Y_\infty \). By taking the expectation in inequality (52), and summing over all integers \( 0 \leq j < \ell \), we have \( 2 \geq EY_\ell \geq Y_0 + \delta(2)E|\{0 \leq j < \ell : z_j \in \bar{S} \cup \bar{S}_2\}| \), implying that \( E[|0 \leq j < \ell : z_j \in \bar{S} \cup \bar{S}_2|] \) is bounded by \( 2/\delta(2) \). This completes the proof of (b). Claim (b) implies that with probability 1, \( |0 \leq j < \infty : z_j \notin S_1| < \infty \), which together with part 1 of Lemma 10 proves part (c).

By selecting \( \epsilon_1 = \frac{\delta^2\epsilon}{16} \), we have \( \delta(2) \geq \epsilon^2\delta/8 \). Therefore, \( 2/\delta(2) \leq \frac{16}{\epsilon^2\delta} \), proving part (d). For every \( i \in N \) and \( j \geq 0 \), by Lemma 10 (part 4), we have

\[
E(\psi^j_{i+1} | \mathcal{F}_j) \geq w_j - \epsilon_1\bar{I}(z_j \in S_2 \cup \bar{S}).
\]

Summing these inequalities over all \( 0 \leq j < \ell \), we have

\[
E(\psi^j_i | z_0) \geq \psi^j_i(z_0) - \epsilon_1 E \sum_{0 \leq j < \ell} \bar{I}(z_j \in S_2 \cup \bar{S}) \geq \psi^j_i(z_0) - \epsilon, \quad \text{and therefore}
\]

\[
E(\psi^j_i | z_0) \geq \psi^j_i(z_0) - \epsilon.
\]

This completes the proof of part (e). □

4.10. The uniform-l-average equilibrium strategy \( \sigma \)

For every player \( i \in N \) let \( \sigma^i_{\epsilon^{-1}} \) be an \( N \setminus \{i\} \) strategy profile and \( s_\epsilon \) a positive number such that for every \( s \geq s_\epsilon \) and every strategy \( \tilde{\sigma}^i \) we have \( \gamma^i_j(\tilde{\sigma}^i, \sigma^i_{\epsilon^{-1}}, \tilde{\sigma}^i) \leq \psi^i(z) + \epsilon \). Therefore, for every \( s \geq 0 \) we have

\[
\gamma^i(z, \sigma^i_{\epsilon^{-1}}, \tilde{\sigma}^i) \leq \psi^i(z) + \epsilon + \frac{s_{\epsilon}}{s}.
\]

The strategy \( \sigma \) follows \( \bar{t} \) until the first time \( t = j/\ell \), with \( j \) a positive integer, where by a deviating by a single player, say player \( i \), in the time interval \( (1 - 1)/\ell, j/\ell \) is observed. Thereafter, all players \( i' \neq i \) start playing the \( N \setminus \{i\} \) strategy profile \( \sigma^i_{\epsilon^{-1}} \).

Let \( s > 0 \) and \( 0 \leq T \leq s \) be a \( (\mathcal{H}_\ell^i)_{i=0} \)-stopping time. Recall that \( \bar{t}_k \) denotes the time of the end of the \( k \)-th block and that \( \bar{z}_k = z_{\ell k} \). Let \( k_T \) be the smallest integer \( k \) s.t. \( \bar{t}_k \geq T \).

Note that \( k_T \leq T + N_0 \) and if \( k_T \leq k \leq k_T + \frac{T - \bar{t}_{k_T}}{N_0} - 2 \) then \( \bar{t}_k + N_0 \leq s \).

Note that for every \( T \leq t \leq s \) we have

\[
g^i_t \geq \sum_{k : T \leq t \leq k + 1 \leq s} 1_{\{\bar{z}_k \in S_1\}} 1_{\{\bar{t}_k \leq t - \bar{t}_k + N_0\}} g^i_t
\]

\[
\geq \sum_{k_T \leq k \leq k_T + \frac{T - \bar{t}_{k_T}}{N_0} - 2} 1_{\{\bar{z}_k \in S_1\}} 1_{\{\bar{t}_k \leq t - \bar{t}_k + N_0\}} g^i_t
\]

\[
\geq \sum_{k_T \leq k \leq k_T + \frac{T - \bar{t}_{k_T}}{N_0} - 2} 1_{\{\bar{t}_k \leq t - \bar{t}_k + N_0\}} (g^i_t - 1_{\{\bar{z}_k \notin S_1\}}).
\]

Therefore, using the additivity and monotonicity of the expectation, together with the fact that the expectation equals the expectation of the conditional expectation (that is used in the equality below), we deduce that

\[
E^2_\sigma \int_T^s g^i_t dt \geq E^2_\sigma \int_T^s \sum_{k_T \leq k \leq k_T + \frac{T - \bar{t}_{k_T}}{N_0} - 2} 1_{\bar{t}_k + N_0} (g^i_t - 1_{\{\bar{z}_k \notin S_1\}}) dt
\]

\[
= E^2_\sigma \int_T^s \sum_{k_T \leq k \leq k_T + \frac{T - \bar{t}_{k_T}}{N_0} - 2} N_0(y^i_{N_0}(\bar{z}_k, \bar{t}) - 1_{\{\bar{z}_k \notin S_1\}})
\]

\[
\geq E^2_\sigma \int_T^s \sum_{k_T \leq k \leq k_T + \frac{T - \bar{t}_{k_T}}{N_0} - 2} N_0(\psi^i(\bar{z}_k) - 1_{\{\bar{z}_k \notin S_1\}} - O(\epsilon))
\]
where \( e \) is the maximum over \( z \in S \) of the expectation with respect to \( P_\sigma^z \) of \([k < \infty : \tilde{z}_k \notin S_1]\). In particular, by setting \( T = 0 \), we have
\[
E_\sigma^z \frac{1}{s} \int_0^s g^i_t dt \geq v^i(z) - O(\varepsilon) \text{ for all } s \text{ sufficiently large.} \tag{54}
\]

Let \( \tilde{\sigma}^i \) be a strategy of player \( i \) and let \( T_\ell \) be the minimum of \( s \) and the (random) smallest time \( j/\ell \) such that \( \tilde{\sigma}^i \) deviates from the play under \( \sigma \) at some time in the interval \([j - 1/\ell, j/\ell]\). Let \( \tilde{\sigma} \) denote for short the strategy profile \((\tilde{\sigma}^1, \tilde{\sigma}^2)\). Note that the \( P_\sigma \)-distribution and the \( P_{\tilde{\sigma}} \)-distribution of \( T_\ell \) (and thus also of \( s - T_\ell \)) coincide. By the additivity of the expectation,
\[
\gamma_s^i(z, \sigma^{-i}, \tilde{\sigma}^i) = E_\sigma^z \frac{1}{s} \int_0^s g^i_t dt = E_\sigma^z \frac{1}{s} \int_{T_\ell}^s g^i_t dt + E_\sigma^z \frac{1}{s} \int_{T_\ell}^s g^i_t dt. \tag{55}
\]

The \( P_\sigma \)-distribution and the \( P_{\tilde{\sigma}} \)-distribution of \( 1_{\{t < T_\ell - 1/\ell\}} g^i_t \) (where \( 1_{\{t < T_\ell - 1/\ell\}}(t) = 1 \) if \( t < T_\ell - 1/\ell \) and \( 1_{\{t < T_\ell - 1/\ell\}}(t) = 0 \) otherwise) coincide, and \( 0 \leq g^i_t \leq 1 \). Therefore
\[
E_\sigma^z \frac{1}{s} \int_0^s g^i_t dt \leq E_\sigma^z \frac{1}{s} \int_{T_\ell}^s g^i_t dt + \frac{1}{s\ell}. \tag{56}
\]

By the definition of \( \tilde{\sigma} \),
\[
E_\sigma^z \frac{1}{s} \int_{T_\ell}^s g^i_t dt \leq -\frac{1}{s} E_\sigma^z (s - T_\ell) v^i(z_{T_\ell}) + \varepsilon + \frac{5\varepsilon}{s}. \tag{57}
\]

Recall that \( 0 \leq v^i(z) \leq 1 \) and that for every stopping time \( T \) and strategy profile \( \tau \), \( P_T(z_T = z_T \forall T \leq T + 1/\ell) \geq e^{-1\mu}/\ell \geq 1 - 1/\ell \). Therefore,
\[
-\frac{1}{s} E_\sigma^z (s - T_\ell) v^i(z_{T_\ell}) \leq -\frac{1}{s} E_\sigma^z (s - T_\ell) v^i(z_{T_\ell}) + \frac{1}{\ell}. \tag{58}
\]

By (53)
\[
-\frac{1}{s} E_\sigma^z (s - T_\ell) v^i(z_{T_\ell}) \leq E_\sigma^z \frac{1}{s} \int_{T_\ell}^s g^i_t dt + \varepsilon + \frac{5\varepsilon}{s}. \tag{59}
\]

Let \( \ell \) be sufficiently large so that \( \frac{1}{\ell} < e \). Summing inequalities (55)–(59), we deduce that
\[
\gamma_s^i(z, \sigma^{-i}, \tilde{\sigma}^i) \leq \gamma_s^i(z, \sigma) + 4\varepsilon \quad \text{for all } s \text{ sufficiently large.} \tag{60}
\]

Similarly, the limit of \( \frac{1}{s} \int_0^s g^i_t dt \) as \( s \to \infty \) exists a.e. with respect to \( P_\sigma^z \), and \( E_\sigma^z \lim_{s \to \infty} \frac{1}{s} \int_0^s g^i_t dt = u^i(z) \geq -4\varepsilon + E_\sigma^z \limsup_{s \to \infty} \frac{1}{s} \int_0^s g^i_t dt \).

This completes the proof of Theorem 4. \( \square \)

5. Proof of Theorem 5

Theorem 5 is straightforward if \( \|\mu\| = 0 \). Therefore, we assume that \( \|\mu\| > 0 \).

In order to prove Theorem 5, it suffices to prove that for every continuous-time stochastic game \( \Gamma \), \( \tilde{w} \in (W_1^d)^N \), and \( \varepsilon > 0 \), there is a vector payoff \( v \in \mathbb{R}^{S \times N} \), a strategy profile \( \sigma \), and a neighborhood \( U \) of \( \tilde{w} \) in \( W_N^d \), such that for every \( \tilde{w}^* \in U \), a player \( i \), and a strategy profile \( \tau \) such that \( \tau^j = \sigma^j \) for all \( j \neq i \), we have
\[
\varepsilon + E_\sigma^z \int_{[0, \infty]} g^i_t d\tilde{w}^*(t) \geq v^i(z) \geq E_\sigma^z \int_{[0, \infty]} g^i_t d\tilde{w}^*(t) - \varepsilon, \tag{61}
\]
where \( \tilde{w}^* \) stands for \((w^*)^i\).
The idea of the proof is as follows.

We consider an auxiliary k-stage discrete-time game that depends on a uniform-l-average payoff vector \( u \), a sufficiently large \( t_0 \), and a partition \( t_0 = 0 < t_1 < \ldots < t_k = t_\infty \) of \([0, t_\infty)\).

The set of feasible actions \( a_i^j \) of player \( i \) at stage \( j = 0, \ldots, k-1 \) is \( A_i^j(z_i) \), the conditional probability that \( z_{i+1} = z' \) given \( z_i = z \neq z' \) and \( a_i \) is \( (t_{j+1} - t_j)\mu(z', z, a_i) \), and the payoff to player \( i \) in this auxiliary game is \( w^i(\infty)u^i(z_k) + \sum_{j=0}^{k-1} w^i((t_{j+1}, t_{j+1}))g^i(z_i, a_i) \).

By backward induction, the auxiliary game has a Markov equilibrium strategy \( \tau \) and a corresponding equilibrium payoff vector \( v \).

We associate to an equilibrium strategy profile \( \tau \) of this auxiliary game a strategy profile \( \sigma \) of the continuous-time game. At time \( t_j \leq t < t_{j+1}, \ 0 \leq j < k \), the strategy \( \sigma \) plays the \( j \)-th stage mixed action of \( \tau \) and starting at time \( t_k = t_\infty \), the strategy \( \sigma \) follows a uniform-\( \frac{1}{\gamma} \)-equilibrium that corresponds to the uniform-l-average equilibrium payoff vector \( u \).

We prove that, if \( t_\infty \) is sufficiently large and \( \max_j(t_{j+1} - t_j) \) is sufficiently small, then the equilibrium payoff \( v \) of the auxiliary game, the associated strategy \( \sigma \), and a suitable neighborhood \( U \) of \( \bar{w} \), satisfy (61).

Now we turn to the formal proof.

Let \( 1 > \varepsilon > 0, \bar{w} \in (W^1)^N \), and let \( \Gamma \) be a continuous-time stochastic game. By the covariance properties, we assume w.l.o.g. that \( 0 \leq g^i \leq 1 \) and \( \|\mu\| = 1/2 \).

Let \( u \in \mathbb{R}^{\times N} \) be a uniform l-average payoff of \( \Gamma \), and let \( \sigma \) be a strategy profile and \( t_\infty > 1 \) such that for every \( z \in S, t \geq t_\infty \), and a strategy profile \( \tau \) such that \( \tau^i = \sigma^i \) for every \( j \neq i \),

\[
\varepsilon + \frac{1}{L} E_{\sigma^j} \int_0^t g^j_i dt \geq u^i(z) \geq \frac{1}{L} E_{\tau} \int_0^t g^j_i dt - \varepsilon, \text{ and}
\]

\[
\varepsilon + E_{\sigma^j} g^j_i \geq u^i(z) \geq E_{\tau^j} g^j_i - \varepsilon. \tag{62}
\]

Fix \( t_\infty \) such that \( 2(t_\infty + \varepsilon) \max_{i \in [N]} w^i((t_\infty, \infty)) < \varepsilon/3 \).

Fix an increasing sequence \( 0 = t_0 < t_1 < \ldots < t_k = t_\infty \) and \( (\varepsilon_j)_{j=0}^k \) such that \( \sum_{0 \leq j \leq k} \varepsilon_j < \varepsilon/2 \) and for \( 1 \leq j < k \),

\[
(t_{j+1} - t_j)w^j_i + (t_{j+1} - t_j)\mu(z', z, a)w^j_i \text{ and for } 0 \leq j < k, \text{ and } \varepsilon_j \text{ where } w^j_i := w^i((t_j, t_{j+1})).
\]

The existence of such sequences, \( (t_j)_{j=0}^k \) and \( (\varepsilon_j)_{j=0}^k \), follows from the fact that for a fixed \( t_\infty, \sum_{j=0}^{k-1} (t_{j+1} - t_j)w^j_i \) goes to 0 as \( \max_{j \leq k}(t_{j+1} - t_j) \to 0 \).

We define the strategy profile \( \sigma \) as follows. For \( t \geq t_k, \sigma(h, t) = \sigma_1(L_k h, t - t_k) \), where \( (L_k h)(t) = h(s + t) \), and for \( 0 \leq j < k \) and \( t_j \leq t < t_{j+1}, \sigma(h, t) = \sigma_1(z_{t_j}) \), where \( \sigma_1 \) is defined below.

The definition of the stationary strategies \( \sigma_j (1 \leq j < k) \) and the auxiliary vectors \( u^j_i \in [0, 1]^S \) \((i \in N \text{ and } 0 \leq j \leq k) \) is recursive.

\[
u^j_i = w^i((\infty))u^i. \text{ For } 0 \leq j < k \text{ and } z \in S, \sigma_j(z) \text{ is an equilibrium of the stage game with a payoff function to player } i \text{ given by}
\]

\[
a \mapsto w^j_i g^i(z, a) + u^j_{i+1}(z) + (t_{j+1} - t_j) \sum_{z' \in S} \mu(z', z, a)u^j_{i+1}(z'), \ \text{ and}
\]

\[
u^j_i(z) = w^j_i g^i(z, \sigma_j(z)) + u^j_{i+1}(z) + (t_{j+1} - t_j) \sum_{z' \in S} \mu(z', z, \sigma_j(z))u^j_{i+1}(z').
\]

Now we show that for every \( 0 \leq j \leq k \) we have \( 0 \leq \nu^j_i(z) \leq w^i((t_j, \infty)). \)

As \( 0 \leq g^i \leq 1, 0 \leq u^i(z) \leq 1 \). Therefore \( 0 \leq \nu^j_i(z) \leq w^i((\infty)) \leq w^i((t_k, \infty)). \)

Fix \( 0 \leq j < k \) and assume that \( 0 \leq \nu^j_i(z) \leq w^i((t_{j+1}, \infty)) \). As \( t_{j+1} - t_j \leq 1 \) (which follows from the assumption \( (t_{j+1} - t_j)^2 < \varepsilon_j < 1 \) on the increasing sequence \( t_j \)) we have \( 0 \leq 1 + \sum_{z' \in S} \mu(z', z, a) \leq 1 \). In addition, \( \mu(z', z, a) \geq 0 \) for \( z' \neq z \), and \( \sum_{z' \in S} \mu(z', z, a) = 0 \). Therefore, the sum \( u^j_{i+1}(z') + \sum_{z' \in S} \mu(z', z, a)u^j_{i+1}(z') \) is a convex combination of the nonnegative numbers \( u^j_{i+1}(z') \), \( z' \in S \). Each of these nonnegative numbers is \( \leq w^i((t_{j+1}, \infty)) \) (by assumption) and therefore also their convex combination is \( \leq w^i((t_{j+1}, \infty)) \).

Therefore, by the definition of \( u^j_i(z) \), we have

\[
0 \leq u^j_i(z) = w^j_i g^i(z, \sigma_j(z)) + u^j_{i+1}(z) + (t_{j+1} - t_j) \sum_{z' \in S} \mu(z', z, \sigma_j(z))u^j_{i+1}(z') \leq w^j_i + w^i((t_{j+1}, \infty)) = w^i((t_j, \infty)).
\]

Therefore, \( 0 \leq u^j_i \leq w^i((t_j, \infty)). \)
Fix a player \( i \). We claim that there is a neighborhood \( U \) of \( w^i \) such that for every \( w^* \in U \) and every strategy profile \( \tau \) such that \( \tau^j = \sigma^j \) for every player \( j \neq i \), we have

\[
2\varepsilon + E^*_\alpha \left( \int_{[t_j, t_{j+1})} g^i \, dw^*(t) + u^j_{t+1}(z_{t+1}) | h_{t_j} \right) \geq E^*_\tau \left( \int_{[0, \infty)} g^i \, dw^i(t) \right) - 4\varepsilon. \tag{64}
\]

For \( 0 \leq j < k \), let \( U_j := \{ w^* \in W : |w^*_j - w^j| < \varepsilon_j \} \) where \( w^*_j := w^*(t_j, t_{j+1}) \).

By Lemma 2 and the inequality \((t_{j+1} - t_j) w^j_j + (t_{j+1} - t_j)^2 < \varepsilon_j \), for any \( 0 \leq j < k \) and \( w^* \in U_j \),

\[
E^*_\alpha \left( \int_{[t_j, t_{j+1})} g^i \, dw^*(t) + u^j_{t+1}(z_{t+1}) | h_{t_j} \right) \geq E^*_\tau \left( \int_{[t_j, t_{j+1})} g^i \, dw^i(t) + u^j_{t+1}(z_{t+1}) | h_{t_j} \right) - |w^*_j - w^j| \geq g^i(z_{t_j}, \sigma(z_{t_j})) w^j_j + u^j_{t+1}(z_{t+1}) - \sum_{z' \in S} u^j_{t+1}(z') (t_{j+1} - t_j) \mu(z', z_j(z)) - 2\varepsilon_j \geq u^j_{t+1}(z_{t+1}) - 2\varepsilon_j \tag{65}
\]

and by denoting by \( \tilde{y}_t \) the mixed action profile of \( \tau \) at time \( t_j \leq t < t_{j+1} \), conditional on \( h_{t_j} \) and \( z_s = z_{t_j} \) for every \( t_j \leq s \leq t \),

\[
u^j_{t+1}(z_{t+1}) = w^j_j g^i(z_{t_j}, \sigma(z_{t_j})) + u^j_{t+1}(z_{t+1}) + (t_{j+1} - t_j) \sum_{z' \in S} u^j_{t+1}(z') \mu(z', z_j(z_{t_j})) \geq w^*_j g^i(z_{t_j}, \tilde{y}_t) + u^j_{t+1}(z_{t+1}) + (t_{j+1} - t_j) \sum_{z' \in S} u^j_{t+1}(z') \mu(z', z_j, \tilde{y}_t) - |w^*_j - w^j| \]

\[
E^*_\tau \left( \int_{[t_j, t_{j+1})} g^i \, dw^*(t) + u^j_{t+1}(z_{t+1}) | h_{t_j} \right) - 2\varepsilon_j. \tag{66}
\]

Therefore, if \( w^* \in \cap_{0 \leq j < k} U_j \), then, by summing inequalities (65), respectively (66), over \( 0 \leq j < k \), we have

\[
\varepsilon + E^*_\alpha \left( \int_{[0,\infty)} g^i \, dw^*(t) + u^i_{t}(z_{t}) \right) \geq u^i_{t}(z_{t}) \geq E^*_\tau \left( \int_{[0,\infty)} g^i \, dw^i(t) + u^i_{t}(z_{t}) \right) - \varepsilon. \tag{67}
\]

Set \( \bar{t} := t_k \) and \( U_k := \{ w^* \in W : |w^*(t_k, \bar{t})| = |w^*(t_k, \infty)| - |w^*(t_k, 0)| < \varepsilon_k \} \). Let \( w^* \in U_k \) and define \( f(t) := \frac{dw^*}{dt} \). Note that as \( w^* \) is a discounting measure, \( f(t) \) exists for all, but countably many exceptions, \( 0 < t < \infty \), and \( f \) is monotonic nonincreasing with \( f(t) \to 0 \) as \( t \to \infty \). We refer to the Riemann–Stieltjes integral. Note that \( - \int_{[t_k, \infty)} (t - t_k) f(t) \, dw^*(t) = w^*((t_k, \infty)) \). Then, by integration by part and by using inequality (62), we have

\[
E^*_\alpha \left( \int_{[t_k, \infty)} g^i \, dw^*(t) | h_{t_k} \right) = -E^*_\alpha \left( \int_{[t_k, \infty)} (\int_{t_k}^t g^i \, dt) \, df(t) | h_{t_k} \right) \geq -E^*_\alpha \left( \int_{[t_k, \infty)} (\int_{t_k}^t g^i \, dt) \, df(t) | h_{t_k} \right) \geq -\int_{[t_k, \infty)} (t - t_k) (u^i(z_{t_k}) - \varepsilon) \, df(t) \]

\[
= (u^i(z_{t_k}) - \varepsilon) w^*(t_k, \infty) - 2w^*(([t_k, \bar{t}])) (\bar{t} - t_k) \geq u^i(z_{t_k}) w^*(t_k, \infty) - \varepsilon,
\]

and if \( \tau \) is a strategy profile such that \( \tau^j = \sigma^j \) for every \( j \neq i \) then
These inequalities imply that

\[ 2\varepsilon + E^Z_\sigma \left( \int_{[t_k, \infty)} g^i \, dw^*(t) \mid h_{t_k} \right) \geq w^*(\{t_k, \infty\}) u^i_k(z_{t_k}) \geq E^Z_\tau \left( \int_{[t_k, \infty)} g^i \, dw^*(t) \mid h_{t_k} \right) - 2\varepsilon. \]  \hspace{1cm} (68)

As \( w^* \in U_k \) (and thus \( |w^*(\{t_k, \infty\}) - w^k(\{t_k, \infty\})| < \varepsilon_k \)), and

\[ \varepsilon + E^Z_\sigma \left( g^i_\infty \, w^*(\infty) \mid h_{t_k} \right) \geq w^*(\infty) u^i_k(z_{t_k}) \geq E^Z_\tau \left( g^i_\infty \, w^*(\infty) \mid h_{t_k} \right) - \varepsilon \]  \hspace{1cm} (69)

(by inequality (63), we deduce that)

\[ 4\varepsilon + E^Z_\sigma \left( \int_{[t_k, \infty)} g^i \, dw^*(t) \mid h_{t_k} \right) \geq w^l(\{t_k, \infty\}) u^l_k(z_{t_k}) \geq E^Z_\tau \left( \int_{[t_k, \infty)} g^i \, dw^*(t) \mid h_{t_k} \right) - 4\varepsilon. \]  \hspace{1cm} (70)

Inequalities (67) and (70) imply that for \( w^* \in U := \cap_{j=0}^\infty U_j \),

\[ 6\varepsilon + E^Z_\sigma \int_{[0, \infty)} g^i \, dw^*(t) \geq u^l_k(z) \geq E^Z_\tau \int_{[0, \infty)} g^i \, dw^*(t) - 6\varepsilon. \]  \hspace{1cm} \square

The (first) part of the proof that ends with inequality (67) proves that for every \( \tilde{w} \in (\tilde{W})^N \) that is supported on \([0, \infty)\), \( \Gamma' \) has a \( \tilde{w} \)-time-separable robust equilibrium.

\section{The continuous-time process}

Recall that \( s_k \) stands for the \( k \)-th time of a state change.

In this section we study properties of the stochastic process of states \((Z_t)_{t \geq 0}\) that is defined by a correlated strategy. Recall that the correlated strategy choice at time \( t \) is, for \( s_k \leq t < s_{k+1} \), a function of \( t \) and the state process up to time \( s_k \).

A special case of a correlated strategy is a Markov correlated strategy. If \( \sigma \) is a correlated strategy and \( \tau \) is a Markov strategy such that \( \sigma(h, t) = \tau(h, t) \) whenever \( s_1(h) > t \), then the \( P_\sigma \) and the \( P_\tau \) distributions of the play up to time \( s_1 \) coincide.

Therefore, for a correlated strategy \( \sigma \), the conditional \( P_\sigma \) distribution of \((Z_t)_{s_k \leq t < s_{k+1}}\), \((X_t)_{s_k \leq t < s_{k+1}}\), given \((Z_t)_{0 \leq t \leq s_k}\), \((X_t)_{0 \leq t \leq s_k}\), is the \( P_\tau \) distribution of \((Z_t)_{0 \leq t \leq s_k}\), \((X_t)_{0 \leq t \leq s_k}\) of a Markov correlated strategy \( \tau \).

Therefore, commenting on the \( P_\tau \) distribution of plays defined by a Markov correlated strategy \( \tau \) is useful for studying the \( P_\sigma \) distribution of plays defined by a correlated strategy \( \sigma \).

A Markov correlated strategy \( \sigma \) defines transition rates as a function of time \( t \). These transition rates are represented by \( S \times S \) matrices \( Q(t) \) where \( Q_{z,z'}(t) = \mu(z', z, \sigma(z, t)) \). If \( \sigma \) is stationary then the transition matrices \( Q(t) \) are independent of \( t \). If \( \sigma \) is a continuous Markov correlated strategy then the map \( t \mapsto Q(t) \) is continuous.

The next theorem is a classic\(^{\text{11}}\) result on non-homogeneous continuous-time Markov processes. The result shows that the transition matrices \( F(s, t) \) of the Markov stochastic process with state space \( S \) that is defined by a Markov strategy are well defined. The starting point is the fundamental assumption that \( F^\sigma_{z,z'}(s, s + \delta) = 1_{z'=z} + \int_{s}^{s+\delta} \mu(z', z, \sigma(z, t)) \, dt + o(\delta) \).

\(^{\text{11}}\) For a proof in the case where \( Q(t) \) is continuous in \( t \), see, e.g., Rosenblatt (1962). Variants of this result are stated in, e.g., Miller (1968a) and Martin-Lof (1967). For the present formulation and its proof, see Neyman (2012).
In addition, we derive a quantified continuity property of the state stochastic process as a function of the Markov correlated strategy $\sigma$.

Let $S$ and $T$ be finite sets. An $S \times T$ matrix $Q = (Q_{z,z'})_{(z,z') \in S \times T}$ is a transition matrix if for all $(z,z') \in S \times T$ we have $Q_{z,z'} \geq 0$ and for all $z \in S$ we have $\sum_{z' \in T} Q_{z,z'} = 1$.

Let $M$ be the space of all $S \times S$ matrices $Q$, let $M_0$ be its subspace of all matrices $Q$ with $\sum_{z \in S} Q_{z,z} = 0$ for every $z \in S$ and $Q_{z,z'} \geq 0$ for all $z \neq z'$, and let $M_1$ be the subspace of $M$ of all transition matrices. The identity matrix is denoted by $I$.

The space $M$ is a (noncommutative) Banach algebra with the norm $\|Q\| = \max_{z \in S} \sum_{z' \in S} |Q_{z,z'}|$, and the spaces $M_0$ and $M_1$ are closed subalgebras of $M$. For an ordered list $F_1, \ldots, F_J \in M$ we denote by $\prod_{i=1}^J F_i$ the matrix (ordered) product $F_1 F_2 \ldots F_J$.

**Theorem 6.** Let $Q : [0, \infty) \to M$ (respectively, $\to M_0$) be measurable, with $\|Q(t)\| \leq C$, and let $Q^\eta : [0, \infty) \to M$ converge $w^*$ to $Q$. Then there is a unique family $F(s,t) = F^Q(s,t)$, $t \geq s \geq 0$, of matrices in $M$ (respectively, in $M_1$) such that for all $0 \leq s \leq t_1 \leq t$ and $\delta > 0$ we have

$$F(s,s) = I$$

(71)

$$\|F(t, t + \delta) - I - \int_t^{t+\delta} Q(x) \, dx\| \leq o(\delta),$$

(72)

where $o(\delta)$ is a function of $\delta$ such that $o(\delta)/\delta \to 0$ as $\delta \to 0^+$, and

$$F(s, t) = F(s, t_1) F(t_1, t).$$

(73)

This unique family satisfies

$$\frac{\partial F}{\partial t}(s,t) = F(s, t) Q(t)$$

for almost every $t$

(74)

and for $\delta \leq 1/C$ we have

$$\|F(t, t + \delta) - I - \delta \int_t^{t+\delta} Q(x) \, dx\| \leq C \delta^2.$$ (75)

The above theorem derives the unique transition matrices $F(s,t)$ that are defined by a Markov strategy $\sigma$ via the transition rate matrices $Q$, where $Q_{z,z'} = \mu(z', z, \sigma(z'))$.

It is easy (and classic) to prove that any state process $(z_t)_{t \geq 0}$ with law $P$, such that $P(z_t = z' | z_s = z) = F_{z,z'}(s,t)$, has, with probability 1, left and right limits everywhere, and for each fixed $t$ it is, with probability 1, continuous at $t$.

Therefore, as our game payoffs are of the form $\int g \, dw(t)$, and $w$ has at most countably many atoms, there is no loss of generality in assuming in our game model that the state process is right continuous.

A correlated strategy $\sigma$ defines a probability distribution $P_\sigma$ on the state process $h^5$. The mixed-action choice of $\sigma$ at time $t$ is a function of the state process (up to time $t$) and the time $t$, and therefore is denoted by $\sigma(h^S, t)$.

**Lemma 12.** Let $\sigma$ and $\tau$ be correlated strategies, $z = z_0$ an initial state, and $T$ a bounded $(H^5)$-stopping time. For any positive integer $\ell$, let $T(\ell) := \min\{j/\ell : j/\ell \geq T\}$, and let $T_1(\ell)$, or $T_1$ for short, be defined by $T_1 = T(\ell) \wedge i/\ell := \min(T(\ell), i/\ell)$.

For $i \geq 0$ and $T_i \leq t < T_{i+1}$, we denote by $\hat{X}_i^\ell$, respectively $\hat{Y}_i^\ell$, the mixed-action choice of the correlated strategy $\sigma$, respectively $\tau$, at time $t$ conditional on no state change in the time interval $[T_i, T_{i+1})$.

Then, the norm distance between the $P_\sigma$ and $P_\tau$ distributions of $z_i^\ell := z_0, z_1^\ell, \ldots, z_{T_i}$, respectively $(z_t)_{t \leq T}$, is bounded by

$$4n \|\mu\|/\ell^2 + 2\|\mu\| E_\sigma^2 \sum_{i=1}^n \int_{T_{i-1}}^{T_i} (\hat{X}_i^\ell - \hat{Y}_i^\ell) \, dt \| \leq 4n \|\mu\|/\ell^2 + 2\|\mu\| E_\sigma^2 \int_0^{T_\sigma} \|\hat{X}_i^\ell - \hat{Y}_i^\ell\| \, dt.$$ (76)

respectively $2\|\mu\| E_\sigma^2 \int_0^T \|\sigma(h^5, t) - \tau(h^5, t)\| \, dt$.

The proof uses Lemma 2 and the following lemma.

**Lemma 13.** Let $(\emptyset, \Omega) = (F_0 \subset F_1 \subset \ldots \subset F_n)$ be an increasing sequence of $\sigma$-algebras of subsets of a nonempty set $\Omega$, and let $P$ and $Q$ be two probability measures on the measurable space $(\Omega, F_n)$. Denote by $F_i$ the set of all $[-1, 1]$-valued $F_i$-measurable functions on $\Omega$ and by $P_i$ and $Q_i$ the probability measures $P$ and $Q$ on $(\Omega, F_i)$. Then,
\[ \| P_i - Q_i \| = \sup \{ E_P f - E_Q f : f \in F_i \}, \quad \text{and} \]
\[ \| P - Q \| \leq \sup \{ E_P \sum_{i=1}^{n} (E_P(f_i | F_{i-1}) - E_Q(f_i | F_{i-1})) : f_i \in F_i \}. \]  

(76) \[
\text{Proof.} \quad \| P_i - Q_i \| := \sup \{ P(C) - Q(C) + Q(\Omega \setminus C) - P(\Omega \setminus C) : C \in F_i \} = \sup \{ E_P(1_C - 1_{\Omega^i}(C)) - E_Q(1_C - 1_{\Omega^i}(C)) : C \in F_i \} \leq \sup \{ E_P f - E_Q f : f \in F_i \}.
\]

As \( L_{\infty}(P_i + Q_i) \supset f \mapsto E_P f - E_Q f \) is linear and \( 1_C - 1_{\Omega^i}(C) \), \( C \in F_i \), are the extreme points of \( L_{\infty}(P_i + Q_i) \), the last inequality is an equality. This proves the first part of the lemma, and the second part for \( n = 1 \).

Assume that (76) holds for \( n = k \) and let \( f \in F_{k+1} \). As \( E_Q(f | F_k) \in F_k \), \( E_P f - E_Q f = E_P E_P(f | F_k) - E_Q E_Q(f | F_k) \leq E_P(E_P(f | F_k) - E_Q(f | F_k)) + \| P_k - Q_k \| \leq \sum_{i=1}^{k+1} (E_P(f_i | F_{i-1}) - E_Q(f_i | F_{i-1})) : f_i \in F_i \). Therefore, (76) holds for \( n = k + 1 \), and thus by induction for every \( n \).

\[
\text{Proof of Lemma 12.} \quad \text{Set in Lemma 13 } F_i = \mathcal{H}^X_{T_i}, \text{ where } T_i := T \wedge i/\ell = \min(T, i/\ell).
\]

Note that \( \bar{x}_i^T \) and \( \bar{y}_i^T \), \( T_i < T \), are measurable with respect to \( F_{i-1} \).

By Lemma 2, for every \( f_i : S \to [-1, 1] \), \( E^T_{F_i} (f_i(z_T^i) | F_{T_i-1}) - E^T (f_i(z_T^i) | F_{T_i-1}) \mid \leq 4 \| \mu \|^2/\ell^2 + \| \sum_{z' \in S} f_i(z') \| f_{T_i-1}^{T_i} \mu(z', z, \bar{x}_i - \bar{y}_i) \| dt \leq 4 \| \mu \|^2/\ell^2 + 2 \| \mu \| \| f_{T_i-1}^{T_i} \| \| \bar{x}_i - \bar{y}_i \| \| dt \|
\]

By summing these inequalities over \( i = 1, \ldots, n \), the first bound in the lemma follows from Lemma 13.

As \( \| \bar{x}_i^T - \bar{y}_i^T \| \) is bounded and with \( P_\sigma \)-probability 1 converges, as \( \ell \to \infty \), a.e. to \( \| \sigma(h^5, t) - \tau(h^5, t) \| \), the second part of the lemma follows from the bounded convergence theorem.

7. Related literature

7.1. Continuous-time supergames and bargaining

Simon and Stinchcombe (1989) develop a general theory of continuous-time games. The pathologies of continuous-time strategies are resolved there by imposing assumptions (i.e., restrictions) on pure strategies that identify a class of strategies that yield a well-defined outcome. The assumptions are, essentially, that 1) the number of action changes is uniformly bounded on every finite time interval, 2) the strategies are piecewise continuous with respect to time, and 3) there is strong right continuity with respect to histories.

Bergin and MacLeod (1993) develop a model of continuous-time supergames of complete information and study the strategic equilibrium behavior of these games. The pathologies of continuous-time strategies are resolved there by considering only strategies that have inertia. A strategy with inertia is essentially a limit, in a proper sense, of strategies that as a function of history and time select a constant action over a short time interval. In addition, the pure strategies choose a pure stage action, in contrast to our main model, which allows a strategy to choose a mixed action as a function of history.

Perry and Reny (1993) develop a theory of continuous-time bargaining, where players can make offers whenever they like. The pathologies of continuous-time strategies are resolved there by considering only strategies where upon making an offer players must wait a fixed amount of time before making another offer. The continuous-time conclusions are obtained by studying the results when the waiting times go to zero.

7.2. Continuous-time stochastic games

Zachrisson (1964) introduced the theory of continuous-time two-person zero-sum stochastic games. He proves the existence of a value and optimal strategies in a finite-horizon game where the payoff of a play is the sum of the integral of the payoffs \( g_t = g(t, z, x) \) and a terminal payoff that is a function of the state at the end of the game.

In Zachrisson’s model it is assumed that the space of strategies of each player is the space of Markov strategies. The assumption that all strategies are Markov strategies is restrictive but indeed innocuous in the model of two-person zero-sum games in the sense that Zachrisson’s solution – the value and optimal strategies – is also a solution when one considers all strategies.

On the other hand, Zachrisson’s model allows for time-dependent compact action spaces, transition rates, and payoffs. Explicitly, the action space of player \( i \) is a compact subset \( X_i(t, z) \) of a Euclidean space that depends continuously (in the Hausdorff metric) on \( t \), the transition rates \( \mu(z', z, t, x) \) depend continuously on \( t \) and \( x \), and the payoff function \( g(t, z, x) \) depends also on the time \( t \). The payoff in Zachrisson’s \( T \)-horizon game is \( \int_0^T g(t, z, x) dt + f(z_T) \). A condition on the existence of optimal strategies in a family of auxiliary one-stage games \( x_t \mapsto g(t, z_t, x_t) + \sum_z \mu(z, z_t, t, x_t) v(z), v \in \mathbb{R}^2 \) is shown to suffice for the existence of a value and optimal strategies.
Lai and Tanaka (1982) prove that a finite12 continuous-time stochastic game has a stationary equilibrium in the game where the players are restricted to Markov strategies.

Jasso-Fuentes (2005) studies the equilibrium conditions for Markov games with Polish state and action spaces when the game is restricted to Markov strategies. However, existence is not addressed there.

The relation between the value of the discounted continuous-time and the discounted discrete-time Markov decision process, called the uniformization technique, is discussed also in Howard (1960).

Guo and Hernandez-Lerma (2005) study two-person non-zero-sum continuous-time stochastic games with stationary discounted payoff criteria (and Borel action spaces), and give conditions that ensure the existence of Nash equilibrium in stationary strategies.

Yehuda Levy (2013) proved that a non-zero-sum stochastic game of fixed duration has a Markovian correlated equilibrium.

References

Miller, B.L., 1968a. Finite state continuous time Markov decision processes with a finite planning horizon. SIAM J. Control 6, 266–280.

12 The existence result in Lai and Tanaka (1982) applies also to a class of games with a countable state space and infinite action spaces.